# Food Stamp Certification Periods and Payment Accuracy 

# State Experience During 1997-2001 

By Gregory Mills, Don Laliberty, Christopher Rodger, Abt Associates Inc.

ERS project representative: William Levedahl, 202-694-5431, levedahl@ers.usda.gov


#### Abstract

Quality control (QC) reviews of Food Stamp Program (FSP) cases show that error rates across States range from less than 5 percent to more than 25 percent when both overpayment and underpayment error are combined. This study uses QC data for 1997-2001 and a Markov probability framework to characterize year-to-year national and State error rates into variations due to errors occurring with first-month cases (those approved at initial certification); ongoing cases (those subject to the interim action process); and expiring cases (those subject to the recertification process). This information can be used in planning corrective actions by focusing attention on phases of the administrative process that are more responsible for errors. This study also explores the effect on payment accuracy and FSP participation of more frequent recertification of food stamp cases. A motivating concern is that the use of short certification periods ( 3 months or less) as a strategy to reduce case error may unintentionally reduce program participation.


This report was prepared byAbt Associates Inc. under a research contract with the Economic Research Service. The views expressed are those of the authors and not necessarily those of ERS or USDA.

## Acknowledgments

The authors wish to acknowledge the many individuals whose guidance, support, and cooperation have been essential to completing this research. Most importantly, the following current and former staff members of the Economic Research Service, U.S. Department of Agriculture (USDA), provided useful feedback throughout the project: Bill Levedahl (Project Officer), David Smallwood, Margaret Andrews, Mark Prell, Parke Wilde, and Nader Kabbani. Staff from USDA’s Food and Nutrition Service (FNS), especially Jenny Genser and Karen Peko, were helpful in providing national tabulations of food stamp quality control (QC) data. Karen Cunnyngham of Mathematica Policy Research also assisted in making available, at FNS' request, annual QC data files containing detailed case review information necessary to implement the modeling approach. Chris Hamilton and John Kirlin of Abt Associates offered very help suggestions on many technical aspects of the research as the project evolved. Missy Robinson ably assembled the text, exhibits, and appendices for submission.

## Table of Contents

Executive Summary ..... ii
Analytic Framework ..... ii
Findings: National Error Rate Trends ..... v
Findings: State-by-State Error Patterns ..... v
Findings: Effects of More Frequent Recertification ..... v
Chapter One: Background and Objectives ..... 1
Trends in Food Stamp Participation and Error. ..... 2
Developments in Food Stamp QC Policy ..... 6
Chapter Two: Modeling Program Participation and Error ..... 8
Illustrative Models ..... 9
Modeling Approach Adopted for This Study ..... 12
Chapter Three: Deriving the Transition Matrix ..... 15
Representing Food Stamp Administrative Procedures ..... 15
Using National Food Stamp Quality Control Data ..... 17
Calculation of Cell Counts and Transition Probabilities ..... 22
Derivation of Row Totals ..... 23
Long-Term Outcomes Implied by the Transition Matrix ..... 28
Chapter Four: National Trends in Program Participation and Error Rates ..... 31
Underlying Trends in Participation Rates ..... 31
Underlying Trends in Error Rates ..... 37
Chapter Five: State-by-State Error Patterns ..... 42
Deriving State-Specific Models ..... 42
Differences Among States in Their Underlying Error Patterns ..... 53
Chapter Six: Effects of More Frequent Recertification ..... 54
Using the Model to Test Alternative Scenarios ..... 54
Findings ..... 56
Bibliography ..... 61
Appendix A - Distribution of Food Stamp Households by Length of Certification Period
Appendix B - Procedures For Deriving The Transition MatrixAppendix C - National Estimates of Cell Counts and Transition ProbabilitiesAppendix D - State-by-State Estimates of Cell Counts and Transition ProbabilitiesAppendix E - Discrete Time Markov Chains

## Executive Summary

This study explores the effects on payment accuracy and household participation in the Food Stamp Program of more frequent recertification of food stamp cases. A motivating concern for this research is that the use of short certification periods (three months or less in duration) as a strategy to reduce error among cases with earnings may unintentionally reduce program participation among such households.

This research on food stamp error and participation has three major objectives:

- to examine recent year-to-year national trends in error rates as affected by three food stamp administrative processes-initial certification, interim action, and recertification;
- to examine state-by-state differences in error rates and to diagnose the performance of low-error and high-error states in terms of their effectiveness at conducting initial certification, interim action, and recertification; and
- to examine the extent to which more frequent recertification, as a measure intended to reduce payment error especially among cases with earnings, may inadvertently lower program participation.

This research uses a probability model to explain the underlying month-to-month dynamics of the food stamp case error rate (the average monthly percentage of active cases that are in error) and the aggregate food stamp participation rate (the average monthly percentage of U.S. households that receive food stamps).

There is considerable variation in case error rates among states, ranging from less than 5 percent to more than 25 percent (combining both overpayment and underpayment errors, and including both agency-caused and client-caused errors). An important application of the model is to help explain this variation, by establishing the extent to which a state's error rate is attributable to errors among: first-month cases (those approved at initial certification); ongoing cases (those subject to the interim action process); and expiring cases (those subject to the recertification process). This information is especially important for planning corrective actions, so that state and local program directors can focus attention on the phase of the administrative process that is most responsible for errors.

## Analytic Framework

In analyzing food stamp error, this study focuses on the case error rate, defined as the percentage of active cases whose benefits have been incorrectly computed. Each U.S. household is considered as belonging each month to one of five groups, according to whether
the household is participating in the Food Stamp Program and (if so) whether the household's food stamp payment is correct and whether the household is in the final month of its current food stamp certification period. The five groups are as follows:

- households not participating in the Food Stamp Program ("nonparticipating");
- correctly paid food stamp cases, not in their final certification month ("ongoing correct");
- incorrectly paid food stamp cases, not in their final certification month ("ongoing error");
- correctly paid food stamp cases, in their final certification month ("expiring correct"); and
- incorrectly paid food stamp cases, in their final certification month ("expiring error").

From one month to the next, each household may remain in its group or experience a transition to another. One can express the pattern of month-to-month changes in a five-byfive transition matrix. The entries in this matrix are probabilities, indicating the proportion of households in each group that will remain in their group or shift to a different group in the following month.

At the national level, we have derived the transition matrix for each year 1998 through 2001 using data from food stamp quality control (QC) reviews conducted by all states on an annual sample of their active food stamp cases. The nationwide annual sample consists of more than 46,000 cases. For each state (and the District of Columbia), we have calculated the transition matrix from the state's pooled sample over the period 1998-2001. At both the national and state levels, the model was derived separately for households with and without earnings.

From the estimated parameters of the transition matrix, a series of error indicators has been computed nationally and by state, as follows:

## - Total error rate

$=$ case error rate among all active cases, indicating the percentage of cases that are ineligible, eligible but overpaid (by $\$ 25$ or more monthly), or eligible but underpaid (by $\$ 25$ or more monthly)

## - First-month error rate

$=$ case error rate among first-month (newly certified) active cases

- Next-month error rate: ongoing correct cases
$=$ among ongoing cases not in error this month, the expected percentage in error next month
- Next-month error rate: ongoing error cases
$=$ among ongoing cases in error this month, the expected percentage in error next month
- Next-month error rate: expiring correct cases
$=$ among expiring cases not in error this month, the expected percentage in error next month
- Next-month error rate: expiring error cases
$=$ among expiring cases in error this month, the expected percentage in error next month

The total error rate (also referred to as the combined case error rate, in reported QC statistics) is the most general indicator of the presence of error among active cases. This measure is computed nationally and by state, for each fiscal year; it is included in the annual error rate statistics published by the Food and Nutrition Service (FNS). The other indicators are not included in FNS' conventional error statistics and are components of the total error rate, as follows:

- The first-month error rate is a measure of payment accuracy at initial certification.
- The next-month error rates for ongoing cases indicate the effectiveness of interim actions at preventing errors (among ongoing correct cases) and at detecting and correcting errors (among ongoing error cases).
- The next-month error rates for expiring cases indicate the effectiveness of recertifications at preventing errors (among cases that are correct as they enter recertification) and at detecting and correcting errors (among cases that are in error as they enter recertification).

The total case error rate is very highly correlated with the total dollar error rate (or payment error rate, the summed dollar amount of overpayment and underpayment errors as a percentage of total payment dollars). In FY 2001, the pairwise correlation between the two measures was 0.923 for the fifty states and the District of Columbia. For this reason, there is little loss of generality in focusing here on the total case error rate, even though the total dollar error rate is the basis for determining fiscal liabilities and enhanced funding.

## Findings: National Error Rate Trends

The year-by-year national estimates of the component error rates provide a basis for explaining the downward trend in the national food stamp case error rate during 1998-2001. On a consistently measured basis, applying each year the same $\$ 25$ error threshold for eligible cases, the total case error rate declined nationally from 16.7 percent in 1998 to 12.8 percent in 2001.

The analysis conducted here indicates the factors contributing to the reduction in the national case error rate, separately for cases with earnings and for cases without earnings. For both caseload segments, one factor was a drop in the error rate at initial certification-i.e., the first-month error rate. This trend was accentuated by the fact that first-month cases came to comprise an increasing share of the caseload. Far more significant, however, was the improvement in the rate at which errors were prevented or corrected between formal case actions, through the interim action process. Also noteworthy was the drop in the next-month error rate for expiring error cases, as recertification procedures also appeared to improve in correcting errors.

## Findings: State-by-State Error Patterns

Historically, food stamp error rates have shown substantial variations among states. Based on the state-by-state estimates of the model, some states achieve low total case error rates through strong performance in all three phases of the administrative process: initial certification, interim action, and recertification. Three states are in the lowest quartile for their total case error rate and are below the median in all of the component error parameters: Arizona, Kentucky, and Wyoming. Other states have low overall rates and show strong performance on some but not all phases of certification, indicating the potential for further improvement. Minnesota and South Carolina, for instance, have comparatively low total error rates despite evidence that their recertification procedures are not as effective as most other states in preventing and correcting errors among expiring cases. In contrast, Oklahoma and Oregon do reasonably well in containing errors at interim action and recertification, but each has a very high error rate at initial certification.

## Findings: Effects of More Frequent Recertification

One can use the estimated model to test the long-term effects of alternative scenarios regarding the frequency of recertification. We define the "recertification rate" as the percentage of current-month active cases whose certification is about to expire and who are thus about to undergo recertification. The test conducted here examined an increase of 5 percentage points in the recertification rate for cases with earnings. Using 2000 as the base year, the recertification rate would be increased from 15.5 to 20.5 percent for earnings cases.

This corresponds approximately to a shortening of the average certification period from 6.5 months to 4.9 months.

Such a shift could either decrease or increase the case error rate, depending on whether or not the proportional reduction in error cases (the numerator of the case error rate) exceeds the proportional reduction in active cases (the denominator). Although in principle the effect on the "aggregate participation rate" (the percentage of households receiving food stamps) could also be either upward or downward, one expects a downward change. The reason is that a case termination is more likely when a case is subject to a recertification than in all other months (when the case is, by definition, subject to the interim action process). If recertifications occur more frequently and increase the rate of monthly case closure, this will expectedly lower the aggregate participation rate. Because of the possible sensitivity of estimates to the particular base year used for the calculations, separate estimates were computed using 2000 and 2001 as the base year.

For households with earnings, the alternative scenario involving more frequent recertification was found to result in a small long-term reduction in the number of error cases among households with earnings. The proportional reduction was estimated at 1.7 to 3.3 percent. The associated effect on the case error rate ranged from a small reduction of 0.20 percentage points to an increase of 0.76 percentage points. The latter finding implies that the effects of more frequent recertification on case closure may be proportionally larger for correct cases than for error cases.

The estimates obtained here for the effect of more frequent recertification on program participation among households with earnings is similar to that found in two recent econometric studies that used pooled cross-sectional time-series data at the state level. ${ }^{1}$ In contrast, the effects estimated here on the error rate for earnings cases is less favorable than the error rate reductions found in the earlier research.

These preliminary findings suggest that more frequent recertification for cases with earnings may have effects that are more pronounced in reducing the rate of participation than in reducing the rate of error. As intended, shorter certification periods are shown to lead to higher closure rates for error cases than would otherwise occur through interim action. It also appears, however, that more frequent recertification leads to higher closure rates for correct cases, mitigating the intended reduction in the case error rate.

[^0]
## Chapter One: Background and Objectives

This research on food stamp error rates has three major objectives:

- to examine recent year-to-year national trends in error rates as affected by three food stamp administrative processes-initial certification, interim action, and recertification;
- to examine state-by-state differences in error rates and to diagnose the performance of low-error and high-error states in terms of their effectiveness at conducting initial certification, interim action, and recertification; and
- to examine the extent to which more frequent recertification, as a measure intended to reduce payment error especially among cases with earnings, may also unintentionally reduce program participation.

This research uses a probability model to explain the underlying month-to-month dynamics of the annually measured food stamp case error rate and the level of food stamp participation among U.S. households.

The food stamp administrative processes examined here-initial certification, recertification, and interim action-influence the monthly opening and closing of food stamp cases, and the extent to which active food stamp cases are correctly or incorrectly paid. With respect to recertification, specific attention is given to both the accuracy of recertification decisions and the frequency of recertification. Of particular interest are the patterns of participation and error among cases with earnings. A motivating concern for this research is that the use of short certification periods (three months or less in duration) as a strategy to reduce error among cases with earnings may reduce program participation among such households; thus, some that are circumstantially eligible for assistance may be discouraged from participating by the added procedural requirements.

Two recent studies have provided evidence supporting the proposition that short certification periods tend to reduce food stamp participation for cases with earnings.

- Using national quality control (QC) data for 1990 through 2000, Kabbani and Wilde (2003) found that a 10-percentage-point increase in the proportion of cases subject to short certification periods was associated with a caseload decline of 2.6 percent for households with earnings. Under this same scenario, the corresponding reduction in the payment error rate was estimated at 0.8 percentage points (compared to the state average error rate of 13.1 percent).
- Kornfeld (2002) also obtained estimates of the effect on food stamp participation of a 10-percentage-point increase in the proportion of cases subject to short certification periods, using national QC data for 1987 through 1999. He found the associated caseload decline to be 2.3 percent for cases with adults living separately and 2.4 percent for cases comprised of multiple adults with children.


## Trends in Food Stamp Participation and Error

To set the context for this research, we review here the major national trends in the rates of food stamp participation and error that occurred during federal fiscal years 1997 through 2001. This is the five-year period that followed the enactment of major federal policy reforms to cash assistance and food stamps under the Personal Responsibility and Work Opportunity Reconciliation Act (PRWORA) of 1996 (P.L. 104-193, enacted August 22, 1996 and first effective on October 1, 1996, the start of fiscal year 1997). Exhibit 1 shows key national statistics relating to food stamp participation. During this period the following national trends are noteworthy. (All of the changes noted here at the national level since 1997 are statistically significant, based on tests that make use of the standard errors of the national point estimates.)

- The monthly food stamp caseload declined from 9.452 million households in 1997 to 7.335 million in 2000, before rising to 7.450 million in 2001. The percentage of cases with earnings rose from 24.2 percent to 27.2 percent between 1997 and 2000 and then declined to 26.7 percent in 2001. The percentage of total benefits paid to cases with earnings also rose, from 26.8 percent to 33.7 percent between 1997 and 2000, and then dropped slightly to 33.3 percent in $2001 .^{2}$
- Among individuals eligible for food stamps, the estimated participation rate declined from 64.0 percent in 1997 to 61.6 percent in $2001 .^{3}$ For program-eligible individuals in households with earnings, the "conditional" participation rate fell slightly from 52.9 percent in 1997 to 52.1 percent in 2001. In both instances, the rate moved downward during 1998 and 1999 and then upward in the following two years.

[^1]Exhibit 1: National Food Stamp Program Indicators, 1997-2001

|  | 1997 | 1998 | 1999 | 2000 | 2001 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Annual food stamp benefits (fiscal year, billions) ${ }^{\text {a }}$ | \$19.5 | \$16.9 | \$15.8 | \$15.0 | \$15.5 |
| Monthly food stamp caseload (households in thousands) ${ }^{\text {a }}$ | 9,452 | 8,246 | 7,670 | 7,335 | 7,450 |
| Percentage of food stamp households with earnings (\%) ${ }^{\text {a }}$ | 24.2 | 26.3 | 26.8 | 27.2 | 26.7 |
| Percentage of benefits paid to households with earnings (\%) ${ }^{\text {b }}$ | 26.8 | 30.2 | 33.1 | 33.7 | 33.3 |
| "Conditional" participation rate (among eligible individuals, \%) ${ }^{\text {c }}$ |  |  |  |  |  |
| Individuals in all households | 64.0 | 59.8 | 58.8 | 59.7 | 61.6 |
| Individuals in households with earnings | 52.9 | 49.9 | 42.7 | 50.9 | 52.1 |
| Average length of certification period (months) ${ }^{\text {a }}$ |  |  |  |  |  |
| Total food stamp households | 9.9 | 9.9 | 9.7 | 9.6 | 9.7 |
| Food stamp households with earnings | 8.1 | 7.8 | 7.5 | 7.2 | 7.3 |
| Cases with certification period of 3 months or less (\%) ${ }^{\text {a }}$ |  |  |  |  |  |
| Total food stamp households | 12.1 | 15.0 | 16.9 | 18.5 | 17.2 |
| Food stamp households with earnings | 25.6 | 30.5 | 32.9 | 36.5 | 33.1 |

## Sources:

a. Food and Nutrition Service, U.S. Department of Agriculture, "Characteristics of Food Stamp Households," fiscal years 1997-2001.
b. Food and Nutrition Service, U.S. Department of Agriculture, "Food Stamp Quality Control Annual Report," fiscal years 1997-2001. Overpayment error includes payments to ineligible cases and overpayments to eligible cases. Underpayment error includes underpayments to eligible cases.
c. Karen Cunnyngham, "Trends in Food Stamp Program Participation Rates: 1994 to 2000," Mathematica Policy Research, June 2002; and Karen Cunnyngham, "Trends in Food Stamp Program Participation Rates: 1999 to 2001," Mathematica Policy Research, July 2003. The indicated participation rates are for September of each year. We use the term "conditional" participation rate to distinguish this measure from the later-defined "aggregate" participation rate.
d. Beginning in 2000, the error tolerance was raised from $\$ 5$ to $\$ 25$ in classifying eligible cases as overpaid or underpaid.

- The average length of assigned certification periods declined between 1997 and 2000, but then rose slightly in 2001. For total cases, the average length dropped from 9.9 to 9.7 months over the entire period. The reduction was more pronounced for cases with earnings, from 8.1 to 7.3 months. Exhibits A-1, A-2, and A-3 in Appendix A show the detailed distribution of food stamp households by length of certification period, for total households, households with earnings, and households without earnings, respectively.
- Contributing to the reduced average length of certification periods was the more prevalent use of certifications of three months or less. The percentage of all cases assigned to these "short" certification periods rose from 12.1 to 18.5 percent between 1997 and 2000, but then dropped to 17.2 percent in 2001. For cases with earnings, the corresponding increase was from 25.6 to 36.5 percent during 19972000, declining to 33.1 percent in 2001.

During this same time period, USDA allowed states increasing flexibility to adopt reporting systems that eased the requirements upon clients for reporting income changes or other circumstantial changes within a certification period. Some of the new options (such as quarterly or semiannual reporting) called for less frequent client reporting of changes affecting one's eligibility or benefit. Other options (such as status reporting) limited the reporting requirement only to major shifts in one's employment situation. This increasing federal flexibility was intended to encourage program participation, reduce administrative burdens, and help states control their error rates.

The trends in food stamp error rates during this period are somewhat difficult to track because of a change in error measurement instituted in 2000. Beginning in that year, the error tolerance was raised from $\$ 5$ to $\$ 25$ for eligible cases. Thus, eligible cases with variances of $\$ 5$ to $\$ 24$ are now considered correctly paid, although previously they were considered in error. This change did not affect the error findings for ineligible cases, which are considered in error (overpaid) regardless of their monthly benefit. Nor did this affect the error findings for negative actions (denials and terminations of assistance).

Exhibit 2 shows the national error rates during 1997-2001 for active cases and for negative actions (denials and terminations). For active cases, the exhibit shows both the dollar error rates (dollars in error as a percentage of total dollars paid) and case error rates (active cases in error as a percentage of total active cases); these error rates are shown for overpayment errors (payments to ineligible cases and overpayments to eligible cases), underpayment errors (underpayments to eligible cases), and combined errors.

Exhibit 2: National Food Stamp Error Rates, 1997-2001

|  | 1997 | 1998 | 1999 | 2000 | 2001 |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Active cases ${ }^{\text {a }}$ |  |  |  |  |  |
| Overpayment dollar error rate (\%) | 7.28 | 7.63 | 7.01 | $6.51{ }^{\text {b }}$ | $6.47{ }^{\text {b }}$ |
| Underpayment dollar error rate (\%) | 2.47 | 3.07 | 2.85 | $2.40{ }^{\text {b }}$ | $2.19{ }^{\text {b }}$ |
| Total dollar error rate (\%) | 9.75 | 10.70 | 9.86 | $8.91{ }^{\text {b }}$ | $8.66{ }^{\text {b }}$ |
| Standard error (\%) | 0.17 | 0.17 | 0.18 | 0.18 | 0.18 |
| Overpayment case error rate (\%) | 15.18 | 15.98 | 14.27 | $9.92{ }^{\text {b }}$ | $9.54{ }^{\text {b }}$ |
| Underpayment case error rate (\%) | 8.51 | 9.42 | 9.05 | $4.64{ }^{\text {b }}$ | $4.56{ }^{\text {b }}$ |
| Total case error rate (\%) | 23.68 | 25.40 | 23.32 | $14.56{ }^{\text {b }}$ | $14.11^{\text {b }}$ |
| Distribution of overpayment dollars (\%) |  |  |  |  |  |
| Households with earnings | 38.17 | 43.57 | 49.14 | 49.83 | 53.21 |
| Households without earnings | 61.83 | 56.43 | 50.86 | 50.17 | 46.79 |
| Total | 100.00 | 100.00 | 100.00 | 100.00 | 100.00 |
| Negative actions (denials and terminations) ${ }^{\text {a }}$ |  |  |  |  |  |
| Negative case error rate (\%) | 3.25 | 2.44 | 2.61 | 3.57 | 5.49 |

Sources:
a. Food and Nutrition Service, U.S. Department of Agriculture, "Food Stamp Quality Control Annual Report," fiscal years 1997-2001. Overpayment error includes payments to ineligible cases and overpayments to eligible cases. Underpayment error includes underpayments to eligible cases. All estimates include Guam and Virgin Islands.
b. Beginning in 2000, the error tolerance was raised from $\$ 5$ to $\$ 25$ in classifying eligible cases as overpaid or underpaid.

The published error rates for active cases generally rose slightly from 1997 to 1998, fell marginally from 1998 to 1999, fell dramatically from 1999 to 2000 (reflecting to some degree the change in error measurement), and then fell slightly from 2000 to 2001. In contrast, the negative case error rate rose between 1998 and 2001. It is important to note that all of these measures are sample-based and that some of the national year-to-year changes are not statistically significant. For the total dollar error rate (whose standard error is shown in Exhibit 2), all year-to-year changes are statistically significant except for that between 2000 and 2001.

During this period, errors among cases with earnings comprised an ever-growing share of national overpayment dollars. By 2001, as shown in Exhibit 2, such cases were responsible for 53.2 percent of overpayment dollars, much larger than their 26.7 percent share of active cases and their 33.3 percent share of total benefit dollars in that year (shown in Exhibit 1).

## Developments in Food Stamp QC Policy

Since the early 1980s, food stamp error rates have been a focus of attention among states as a result of federal policies making states vulnerable to some loss of federal funds ("fiscal liabilities") for failing to meet national error rate standards. ${ }^{4}$ During the period examined, the national error rate standard was the national average of the total dollar error rate.

Starting in 1998, the imposition of a liability on a state with an above-average error rate also depended on the share of the state's caseload comprised of households with earnings and households with immigrants. ${ }^{5}$ For each state subject to a "potential liability," an adjusted error rate was calculated by assuming that the caseload shares comprised by households with earners and households with immigrants equaled the national shares in a base year. (For liabilities in 1998 and 1999, the base year was 1996. Starting with 2000, the base year was moved back to 1992, to remove the effect on the adjustment formula of state welfare-to-work initiatives that raised the caseload share of working poor households between 1992 and 1996.)

Only those states whose unadjusted and adjusted error rates both exceeded the national average for a given year then became subject to an "adjusted liability." The liability amount was calculated on a sliding scale, depending on the difference between the state's adjusted error rate and the national average. The number of states subject to adjusted liabilities was 16 in 1998, 16 in 1999, 18 in 2000, and 15 in 2001. The adjusted liability amounts totaled $\$ 27$ million in 1998, $\$ 31$ million in 1999, $\$ 46$ million in 2000, and $\$ 134$ million in 2001. (Of

[^2]the total liability amount in 2001, $\$ 114$ million was associated with California, whose total dollar error rate was 17.37 percent.) ${ }^{6}$

States subject to adjusted liabilities typically entered into negotiation with USDA regarding federal collection of these sanctions. Under these settlements, USDA agreed to waive outright some liability amounts, with states committing to reinvest some amounts in management improvements aimed at reducing error. USDA also then regarded other liability amounts as "at risk," or subject to future collection, if the state's error rate performance did not improve in the following years.

Another component of food stamp QC policy has been the financial incentive of "enhanced funding" of administrative costs (normally matched federally at 50 percent) for states whose total dollar error rate is below 6 percent and whose negative case error rate is below the national average. As with fiscal liabilities, the amount of enhanced funding is computed on a sliding scale, depending on the difference between the state's performance and 5.9 percent. The number of states subject to enhanced funding was 6 in 1998, 6 in 1999, 11 in 2000, and 10 in 2001. (The increase in the latter two years reflected in part the raising of the error threshold for eligible cases from $\$ 5$ to $\$ 25$.) The amount of enhanced funding in total was $\$ 27$ million in 1998, $\$ 39$ million in 1999, $\$ 55$ million in 2000, and $\$ 52$ million in 2001. (For each of these years, Texas accounted for more than half of the nationwide amount of enhanced funding.)

[^3]
## Chapter Two: <br> Modeling Program Participation and Error

This analysis’ modeling approach focused on the underlying research questions, the available program data, and the intrinsic nature of food stamp administrative procedures. The research builds upon previous work by Abt Associates on the effects of recertification and monthly reporting on food stamp error rates. ${ }^{7}$ Importantly, however, the current research goes beyond the earlier work, in the following respects:

- By incorporating food stamp participation, as well as food stamp error, as an outcome of administrative procedures;
- By relaxing the previous modeling assumption that the observed error rates reflect a system already at equilibrium in any given year; and
- By undertaking state-by-state estimates, as well as national estimates, of the model.

The foundation for this work, however, remains the concept that the food stamp administrative process can be represented as a discrete-time, multi-status probability model. ${ }^{8}$ The basic unit of observation is the household, and the basic unit of time is the month. At the start of any given month, each household is considered as occupying one of several possible groups regarding their participation in the program and the accuracy of their payment.

The simplest form of such a model is the discrete-time Markov chain, a stochastic process in which the conditional distribution of each unit's future status is dependent on its present status, but is independent of any prior history. The system is thus considered "memoryless." There is a very substantial body of literature on Markov chains, and the attached bibliography identifies a very small subset of the published work. Appendix E shows a formal specification of discrete-time Markov chains.

Under a Markov model, the period-by-period operations of a real-world system are described in terms of a matrix of transition probabilities. The transition probabilities indicate the pattern of changes in the status of items-in this instance, the status of households regarding their food stamp participation and the correctness of their benefit-as time advances from one period to the next-in this context, from one month to the next. The estimated transition

7 See Mills (1988).
8 To avoid confusion in the current research, we use the word "status" rather than "state" in referring to the condition occupied by a household with respect to the Food Stamp Program. We reserve the use of "state" in referring to the 50 states and the District of Columbia.
matrix can then be used to simulate the effect on the system's performance of operational changes. The particular application of Markov modeling here is to examine the effect of the more frequent food stamp recertification on the size of the active food stamp caseload and the error rate among active cases.

## Illustrative Models

In describing below the methodology of a discrete-time Markov chain, we first introduce a basic two-group model and then move progressively to the five-group model used in the analysis.

## Two-Group Model

The basic two-group model can be used to examine changes over time in the percentage of households that are food stamp participants. We refer to this percentage as the "aggregate participation rate," expressed as $p_{t}$ in any given month $t$. In this simple model, the household population can be viewed as divided into two subgroups: food stamp nonparticipants and food stamp participants. ${ }^{9}$ The condition of the system in any given month can be described by $\mathrm{p}_{\mathrm{t}}$.

From one month to the next, any given household can either shift from one group to the other or can remain within its group. We can describe the monthly dynamics of this system using a two-by-two transition matrix, as shown in Exhibit 3. The entries of the transition matrix indicate the probability that a household of given status in the current month (as indicated by the associated row) will have the same status or a different status in the next month (as indicated by the associated column). The entries in each row are conditional probabilities that sum to 1 , fully describing the transitions that can occur for a household, given its current-month status.

## Exhibit 3: Transition Matrix for Two-Group Model

|  |  | Next-month status: |  |  |
| :--- | :---: | :---: | :---: | :---: |
| Current-month status: |  | Nonparticipating |  | Participating |
|  |  |  |  |  |
| Nonparticipating |  | $1-\mathrm{a}$ |  | a |
| Participating | b |  | $1-\mathrm{b}$ |  |

[^4]The parameter "a" indicates the "case opening rate"-the probability that a household not receiving food stamps in one month will participate in the next month. Correspondingly, the parameter "b" represents the "case closure rate"-the probability that a household receiving food stamps in one month will not participate in the following month. This modeling approach assumes that each parameter does not depend on the length of time that a household has been in its current group. It also assumes that each parameter remains constant over time.

The system's condition will change predictably from one month to the next. In particular, the percentage of households participating in the next month, $\mathrm{p}_{\mathrm{t}+1}$, can be expressed as a function of $p_{t}$, $a$, and $b$ as follows:
(Eq. 1) $\quad \mathrm{p}_{\mathrm{t}+1}=\left[\left(1-\mathrm{p}_{\mathrm{t}}\right) \mathrm{a}\right]+\left[\mathrm{p}_{\mathrm{t}}(1-\mathrm{b})\right]$.
Depending on the values of $a$ and $b$, the system will approach a long-term equilibrium. ${ }^{10}$ From Eq. 1, one can derive the steady-state value of the aggregate participation rate, p*, by solving for $\mathrm{p}_{\mathrm{t}+1}=\mathrm{p}_{\mathrm{t}}$. Doing so, one finds that:
(Eq. 2) $\quad p^{*}=a /(a+b)$

The value of $\mathrm{p}^{*}$ is thus a function of the transition probabilities only and does not depend on the starting value of $\mathrm{p}_{\mathrm{t}}$.

## Three-Group Model

One can elaborate on this basic two-group model by allowing a participating household to be either a correct case or an error case. In any given month, households can thus occupy one of three groups: nonparticipating households (group 1), correctly paid cases (group 2), and incorrectly paid cases (group 3). In this model, the condition of the system in month $t$ is described by the three percentages (summing to one) that indicate the proportion of all households in each group: nonparticipating ( $n_{t}$ ), correct ( $c_{t}$ ), and error ( $e_{t}$ ). The system can be regarded as having two key outcome statistics. One indicator, as before, is the aggregate participation rate, $p_{t}$. In the notation of the three-group model, $p_{t}=c_{t}+e_{t}=1-n_{t}$. The other indicator is the case error rate, $r_{t}$, the percentage of active cases that are in error in month $t$. In the terms of the model, $r_{t}=e_{t} /\left(c_{t}+e_{t}\right)=e_{t} / p_{t}$.

For this three-group model, the month-to-month transitions are fully described in the three-by-three matrix shown in Exhibit 4. As with the matrix for the two-group model, the entries

[^5]in each row sum to one. We have introduced subscripts for each cell entry $\left(\mathrm{p}_{\mathrm{ij}}\right)$ to indicate the specific transition from group $i$ in the current month to group $j$ in the next month.

Exhibit 4: Transition Matrix for Three-Group Model

|  |  | Next-month status: |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Current-month status: |  | Nonparticipating |  | Correct | Error |
|  |  |  |  |  |  |
| Nonparticipating |  | $\mathrm{p}_{11}$ |  | $\mathrm{p}_{12}$ | $\mathrm{p}_{13}$ |
| Correct |  | $\mathrm{p}_{21}$ |  | $\mathrm{p}_{22}$ | $\mathrm{p}_{23}$ |
| Error |  | $\mathrm{p}_{31}$ |  | $\mathrm{p}_{32}$ |  |

This three-by-three matrix provides a simplified representation of the monthly dynamics that underlie the case error rate. In any given month, the opening rate (indicated by "a" in the two-group model) is now the sum of $p_{12}$ and $p_{13}$. Note that the error rate among initially certified cases is $\mathrm{p}_{13} /\left(\mathrm{p}_{12}+\mathrm{p}_{13}\right)$. The probability that a correct case becomes in error the following month is $\mathrm{p}_{23}$. The probability that an error case becomes correct the following month is $\mathrm{p}_{32}$. The probability that an error case leaves the active caseload is $\mathrm{p}_{31} .{ }^{11}$ This model does not enable one to examine the separate roles of interim action and recertification in controlling errors among active cases.

## Five-Group Model

To address the issue of more frequent recertification and its effects on both participation and error in food stamps, the model must be further elaborated. We have done this by subdividing food stamp cases according to whether or not their current certification is about to expire (i.e., "ongoing" versus "expiring"), in addition to whether or not their benefit is in error. The five-by-five transition matrix is shown in Exhibit 5, with row totals again summing to one.

This formulation allows the dynamics of participation and error to be expressed separately for initial certifications (in the first row), interim actions (in the second and third rows), and recertifications (in the fourth and fifth rows). More frequent recertifications affect participation and error by subjecting cases to the conditional probabilities for expiring cases (in the fourth and fifth rows) rather than the conditional probabilities for ongoing cases (in the second and third rows).

[^6]Exhibit 5: Transition Matrix for Five-Group Model

| Current-month <br> status: | Next-month status: |  |  |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Non- <br> participating | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error |  |  |
| Nonparticipating |  |  |  | $\mathrm{p}_{11}$ |  | $\mathrm{p}_{12}$ |  | $\mathrm{p}_{13}$ |
| $\mathrm{p}_{22}$ |  | $\mathrm{p}_{23}$ | $\mathrm{p}_{24}$ |  | $\mathrm{p}_{15}$ |  |  |  |
| Ongoing correct | $\mathrm{p}_{21}$ |  | $\mathrm{p}_{22}$ | $\mathrm{p}_{25}$ |  |  |  |  |
| Ongoing error | $\mathrm{p}_{31}$ |  | $\mathrm{p}_{32}$ |  | $\mathrm{p}_{33}$ | $\mathrm{p}_{34}$ | $\mathrm{p}_{35}$ |  |
| Expiring correct | $\mathrm{p}_{41}$ |  | $\mathrm{p}_{42}$ |  | $\mathrm{p}_{43}$ | $\mathrm{p}_{44}$ | $\mathrm{p}_{45}$ |  |
| Expiring error | $\mathrm{p}_{51}$ |  | $\mathrm{p}_{52}$ | $\mathrm{p}_{53}$ | $\mathrm{p}_{54}$ | $\mathrm{p}_{55}$ |  |  |

The condition of the system in any given month is now described by the five percentages indicating the shares of the total household population that are nonparticipating households (group 1), ongoing correct cases (group 2), ongoing error cases (group 3), expiring correct cases (group 4), or expiring error cases (group 5). As these categories are mutually exclusive and collectively exhaustive, the corresponding shares of the population will always sum to one.

## Modeling Approach Adopted for This Study

With the preceding discussion as background, we now develop more formally the five-group model used in this study.

At the start of any given month, we regard each U.S. household as belonging to one of the following five groups, according to whether the household is participating in the Food Stamp Program and (if so) whether the household's food stamp payment is correct and whether the household is in the final month of its current food stamp certification period:

1. households not participating in the Food Stamp Program ("nonparticipating");
2. correctly paid food stamp cases, not in their final certification month ("ongoing correct");
3. incorrectly paid food stamp cases, not in their final certification month ("ongoing error");
4. correctly paid food stamp cases, in their final certification month ("expiring correct"); and
5. incorrectly paid food stamp cases, in their final certification month ("expiring error").

From one month to the next, each household may experience a transition from one group to another or may remain in its group. One can express the pattern of month-to-month changes in status in a five-by-five transition matrix. The entries in this matrix are transition probabilities of the form $p_{i j}$, indicating the probability that a household in group $i$ in one month will enter group $j$ in the following month (where i may equal $\mathbf{j}$ ).

In developing the initial estimates reported here, we have made several simplifying assumptions that rule out some month-to-month transitions. The corresponding entries in the transition matrix are thus set to zero by definition.

- We assume that a nonparticipating household (Group 1) cannot transition immediately to being an expiring case (i.e., in Group 4 or 5 as an active case in its final certification month). That is, we regard the shortest certification length as two months. (For the 0.5 percent of cases found in the data to have one-month certification periods, we treat these as two-month periods.)
- We also assume that active cases in their final certification month (Group 4 or 5) cannot remain in that group for more than one month at a time. That is, a case that becomes due for certification would not then proceed to being overdue for recertification. ${ }^{12}$

These assumptions are made to avoid having cells in the basic transition matrix that are occupied by a trivially small number of cases. In such instances, the simplifying assumptions serve to combine these negligible categories of cases into adjacent cells of the matrix, with negligible effect on the resulting estimates.

The purpose of adopting this general framework is that it enables the case error rate to be modeled in terms of three distinct administrative processes: initial certification, interim action, and recertification. In the terms of the model, the case error rate can be expressed as the number of households in Groups 3 and 5 divided by the number of households in Groups 2 through 5.

This model focuses on the combined or total case error rate, the percentage of active cases that are either ineligible, eligible but overpaid, or eligible but underpaid, including both agency-related and client-related error. The methodological approach could easily be used to construct separate models for the overpayment case error rate (including both the ineligible and those eligible but overpaid) and the underpayment case error rate (those eligible but underpaid). Alternatively, one could focus solely on agency-related error or client-related error. We examine the total measure here for several reasons. First, as noted above, federal standards for food stamp error are based on the combination of overpayment and

[^7]underpayment errors among active cases. The combined measure is thus the focus of executive and legislative attention. Second, to the extent that errors of both overpayment and underpayment are generated jointly from the same administrative process, it would be inappropriate for the analysis to separate them artificially, implying that one is independent of the other or that one carries more significance than the other. ${ }^{13}$

The decision to focus on the total case error rate, rather than the total dollar error rate, is a pragmatic one. The total case error rate, defined as the ratio of active cases in error to total active cases, is a proportional outcome bounded between zero and one for any set of cases. All active cases contribute equally to the denominator; each case contributes either zero or one to the numerator. In contrast, the total dollar error rate, the ratio of error dollars among active cases to total dollars paid to these cases, can potentially range higher than one for any set of cases. Moreover, cases contribute unequally to the denominator, depending on the magnitude of their monthly benefit, which is itself a function of the error amount. Their contribution to the numerator is also a continuous variable, as small as zero (for correct cases) and as large as the full benefit payment (for ineligible cases).

Arithmetically, the total dollar error rate equals the product of the total case error rate and the ratio of the average monthly dollar error amount to the average monthly benefit payment to an active case. A state's total case error rate is highly correlated with its total dollar rate. For Fiscal Year 2001, the correlation between these two measures was 0.923 , implying that more than 85 percent ( 0.853 ) of the variation in the dollar error rate can be explained by the case error rate. ${ }^{14}$

[^8]
## Chapter Three: Deriving the Transition Matrix

The study's basic analytic construct is a five-by-five matrix of transition probabilities that indicate month-to-month patterns of food stamp participation and error. This chapter describes how the transition matrix is derived from available national data, through a five-byfive crosstabulation of households according to their food stamp status in one month and in the following month.

Exhibit 6 shows the generic notation used to identify the cell entries of the basic crosstabulation. Rows of the matrix correspond to current-month status; columns correspond to next-month status. Illustratively, as explained later, we regard the current month as March and the next month as April. The household count in row i and column j (the cell entry $\mathrm{Q}_{\mathrm{ij}}$ ) indicates the number of households in group i at the start of a month who then belong to group $j$ at the start of the following month. In each month, every household is considered to belong in one (and only one) of the five specified groups.

## Representing Food Stamp Administrative Procedures

This formulation of the model is chosen as a way of describing food stamp participation rates and error rates as outcomes of month-to-month household changes, with an explicit focus on initial certification, interim action, and recertification as distinct stages in the food stamp administrative process.

## Initial Certification

The initial certification process is represented by the first row of the matrix. This row corresponds to households that at the start of the current month are not receiving food stamps but who may transition into the program in the subsequent month. For nonparticipating households at the start of the current month, their distribution at the start of the next month depends on whether: (a) they remain nonparticipants $\left(\mathrm{Q}_{11}\right)$, (b) they become correct cases $\left(\mathrm{Q}_{12}\right)$, or (c) they become error cases $\left(\mathrm{Q}_{13}\right)$. As noted above, we make the simplifying assumption that newly-certified cases are not assigned a one-month certification period; this implies that $\mathrm{Q}_{14}$ and $\mathrm{Q}_{15}$ are both zero.

## Interim Action

The interim action process is represented by the second and third rows of the matrix. These rows correspond to ongoing food stamp cases-households that at the start of the current month are active cases not subject to recertification (i.e., not in the final month of their certification period), and are either correct (second row) or in error (third row). Consider the

## Exhibit 6: Basic Transition Matrix

| Current-month ("March") status | Next-month ("April") status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | $\begin{gathered} \text { Expiring } \\ \text { error } \end{gathered}$ | Total (row) |
|  | Number of households |  |  |  |  |  |
| Nonparticipating | $\mathrm{Q}_{11}$ | $\mathrm{Q}_{12}$ | $\mathrm{Q}_{13}$ | $\mathrm{Q}_{14}=0$ | $\mathrm{Q}_{15}=0$ | $\mathrm{R}_{1}$ |
| Ongoing correct | $\mathrm{Q}_{21}$ | $\mathrm{Q}_{22}$ | $\mathrm{Q}_{23}$ | $\mathrm{Q}_{24}$ | $\mathrm{Q}_{25}$ | $\mathrm{R}_{2}$ |
| Ongoing error | $\mathrm{Q}_{31}$ | $\mathrm{Q}_{32}$ | $\mathrm{Q}_{33}$ | $\mathrm{Q}_{34}$ | $\mathrm{Q}_{35}$ | $\mathrm{R}_{3}$ |
| Expiring correct | $\mathrm{Q}_{41}$ | $\mathrm{Q}_{42}$ | $\mathrm{Q}_{43}$ | $\mathrm{Q}_{44}=0$ | $\mathrm{Q}_{45}=0$ | $\mathrm{R}_{4}$ |
| Expiring error | $\mathrm{Q}_{51}$ | Q ${ }^{2}$ | Q 53 | $\mathrm{Q}_{54}=0$ | $\mathrm{Q}_{55}=0$ | $\mathrm{R}_{5}$ |
| Total (column) | $\mathrm{C}_{1}$ | $\mathrm{C}_{2}$ | $\mathrm{C}_{3}$ | $\mathrm{C}_{4}$ | $\mathrm{C}_{5}$ | Q |
| Data source (see Notes below) | (a) | (b) | (b) | (b) | (b) | (c) |

Notes:
a. First-column entries and total: calculated as row residuals.
b. Second-, third-, fourth-, and fifth-column entries and totals: estimated from current-year QC data.
c. Row totals: The values for $R_{1}$ through $R_{5}$ are derived on the basis of linear interpolation between the current-year and prior-year distribution of households, as explained in the text. The value for $Q$ (total households in the population) is obtained from Census data for April of each year.
second row, the ongoing correct cases. In the next month such a household can either: (a) become a nonparticipating household $\left(\mathrm{Q}_{21}\right)$; (b) remain a correct case, not in its final certification month ( $\mathrm{Q}_{22}$ ); (c) become an error case, not in its final certification month $\left(\mathrm{Q}_{23}\right)$; remain a correct case, in its final certification month $\left(\mathrm{Q}_{24}\right)$; or become an error case, in its final certification month $\left(\mathrm{Q}_{25}\right)$. Similar month-to-month transitions can occur for ongoing error cases-i.e., error cases in the midst of a certification period. The corresponding thirdrow cell entries are $\mathrm{Q}_{31}, \mathrm{Q}_{32}, \mathrm{Q}_{33}, \mathrm{Q}_{34}$, and $\mathrm{Q}_{35}$.

## Recertification

The recertification process is represented by the fourth and fifth rows of the matrix. These rows correspond to expiring cases: cases that in the current month are in the final month of their certification period, either correct (fourth row) or in error (fifth row). Consider the fourth row-the expiring correct cases. The recertification about to be conducted on such a case will result in the household's either: (a) becoming a nonparticipating household ( $\mathrm{Q}_{41}$ ); (b) remaining correct $\left(\mathrm{Q}_{42}\right)$; or (c) becoming in error $\left(\mathrm{Q}_{43}\right)$. Similar month-to-month transitions can occur for expiring error cases-those incorrectly paid in the final month of their certification period (corresponding to $\mathrm{Q}_{51}, \mathrm{Q}_{52}$, and $\mathrm{Q}_{53}$ ). We assume that cases subject to recertification do not become overdue for recertification and that newly-recertified cases are not assigned a one-month certification period (implying that $\mathrm{Q}_{44}, \mathrm{Q}_{45}, \mathrm{Q}_{54}$, and $\mathrm{Q}_{55}$ are all zero).

## Using National Food Stamp Quality Control Data

The primary data source for this analysis is the national Food Stamp Program Quality Control (FSPQC) system. Each year a nationally representative sample of between 45,000 and 50,000 active food stamp cases is selected for QC review, to assess the accuracy of eligibility and benefit determinations made by the administering state and local program agencies. Annual samples range from 300 to 2,400 by state, depending on the size of each state's average monthly active caseload. Approximately one-twelfth of a state's annual sample is drawn each month for review by state QC reviewers. The QC review consists of an inspection of the case record, an interview with the household, and additional field investigation that may include contacts with collateral sources such as employers, landlords, and banks. The information recorded by the state QC reviewers is assembled into a national QC database by USDA's Food and Nutrition Service. National and state error rates are calculated using this database. ${ }^{15}$

[^9]The decision to represent household transitions on a month-by-month basis reflects the fact that the active case QC data describe the status of cases in monthly terms. Specifically, the national QC database provides information about the presence of error for sample cases in their review month, along with information enabling one to deduce their participation and error status in the prior month. In the model developed here, we use the term "next month" to refer to the QC review month. The term "current month" refers to the month prior to the review month.

The model makes no assumption about whether the current-month distribution of households across the five groups (as indicated by the row totals, $R_{i}$ ) is the same as the next-month distribution (as indicated by the column totals, $C_{i}$ ). The expected situation, for the nation or for any particular state, is that the distribution of households is in short-term flux from one month to the next. The estimates derived in this analysis reflect such shortterm month-to-month fluctuations. In general, the nature of a Markov model is such that the monthly transitions will (if uninterrupted) lead to a stable long-term distribution of households. The modeling approach here does not assume, however, that the nation or any individual state has already reached such an equilibrium.

As described below, the task of empirically deriving the model is a matter of using the national QC sample data on active food stamp cases (along with Census data on the total household population) to arrive at cell counts for the basic transition matrix. These cell counts, appropriately weighted to account for the QC sampling procedures, are then used to compute the transition probabilities. One can view the estimation as an accounting process whereby households are placed within the five-by-five classification according to their status in the review month and the preceding month. As described below, a number of assumptions are required to overcome the limitations of the QC data.

In principle, one would want to derive the model separately for each calendar month. This would mean, for instance, that one would calculate twelve separate monthly models over any given annual interval. The monthly sample sizes in the national QC data-approximately 4,000 cases reviewed per month-are not sufficient, however, to support the monthly derivation of the model, even at the national level. This is because some of the monthly transitions in question are based on small segments of the active caseload-for example, cases with earnings that are in error at the final month of their certification period. Moreover, some of the transitions in question are relatively rare events, occurring with a probability of less than 5 percent.

For these reasons, we have chosen to estimate the national model using the full QC sample for a federal fiscal year (October through September). We assume that the month-to-month pattern of change is the same in each of the twelve months of the fiscal year. The estimates for the year in question are thus not specific to any particular calendar month; there is no assumed intra-year variation in monthly error patterns.

Under this approach, we have found it helpful (although not necessary) to consider the data for each fiscal year as having been collected in the mid-year review month of April. We thus interpret the month-to-month transitions as March-to-April transitions. In Exhibit 6, we show the current-month status as referring to "March" and the next-month status as referring to "April."

The active case QC data and Census data allow us to estimate directly some, but not all, cell counts in the matrix. Most notably, the number of cases that exit from the caseload each month is not observed in the active case QC data and must be estimated indirectly as a row or column residual. The resulting closure rates are subject to variation, as they reflect the sampling error of all elements entering the calculation. Exhibit 6 indicates which cell entries are estimated directly from QC data or indirectly as row or column residuals. The firstcolumn entries of the matrix correspond to households that are not participating in the review month ("April"). They are not observed in the active case QC data. These cell counts thus cannot be directly computed and must be derived. If one knows the row totals of the matrix-that is, the preceding ("March") monthly count of households in each of the five subgroups-one can compute each first-column entry as a row residual (i.e., by subtracting all other row entries from the row total).

We have calculated the basic matrix of cell counts for each year 1998 through 2001 using national data from the food stamp QC reviews of active cases. We have excluded Guam and the Virgin Islands from the analysis, as the necessary Census data on annual household counts do not include the territories. ${ }^{16}$

Exhibit 7 shows the total number of food stamp QC cases included in the analysis, for 1997 through 2001.

[^10]Exhibit 7: Size of the Analysis Sample, 1997-2001

|  | 1997 | 1998 | 1999 | 2000 | 2001 |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Number of food stamp quality control cases |  |  |  |  |
| Households with earnings | 11,968 | 12,460 | 12,925 | 12,683 | 12,459 |
| Households without earnings | 35,977 | 33,896 | 33,702 | 33,383 | 33,686 |
| Total households | 47,945 | 46,356 | 46,627 | 46,066 | 46,145 |

Note: Analysis sample excludes Guam and Virgin Islands.

Several adjustments to the data were necessary, as follows:

- To establish a consistent error standard for the historical period, we retroactively applied to the 1997, 1998, and 1999 data the $\$ 25$ error tolerance for active cases that were eligible for food stamps. As noted earlier, this QC policy provision was first effective in 2000. Thus, for each of the years analyzed here, an eligible active case was considered correct if the monthly benefit differed from the correct amount by $\$ 24$ or less. (There is no error tolerance for active cases that are ineligible.) Prior to 2000, the error tolerance for active cases was $\$ 5$. Consistency across years required use of the $\$ 25$ tolerance rather than the $\$ 5$ tolerance, as it was not possible to apply the $\$ 5$ tolerance to cases in 2000 and 2001. (We are unable to identify the eligible cases with error amounts of $\$ 5$ to $\$ 24$, as these cases were found correctly paid and no error amount was recorded.) As shown in Exhibit 8, approximately one-third of the cases classified as errors in 1997-1999 were eligible cases with overpayment or underpayment of $\$ 5$ to $\$ 24$. Such cases were reclassified as correct in this analysis, to establish a consistent error definition.
- We considered cases with a one-month certification period as having a two-month certification period. This was to avoid the calculation of some transition probabilities on the basis of very small samples. Note that less than 0.5 percent of all sample cases were assigned a one-month certification period. (Separately, two cases with an indicated certification length of 0 were excluded from the analysis.)

Exhibit 8: National Case Error Rates with \$5 and \$25 Error Tolerance, 1997-2001



Exhibit 8: National Case Error Rates with \$5 and \$25 Error Tolerance, 1997-2001 (Continued)


Notes:
${ }^{\text {a }}$ Eligible cases with overpayment or underpayment of less than $\$ 5$ are not classified as error cases. (Data not available for 2000 and 2001, as explained in text.)
${ }^{\text {b }}$ Eligible cases with overpayment or underpayment of less than $\$ 25$ are not classified as error cases.

The distinction between cases with earnings and cases without earnings is based on whether the case record indicates the presence of earnings. To some degree, errors among cases "without earnings" are associated with unreported earnings amounts. The model does not address explicitly the extent to which households move between the "with earnings" and "without earnings" categories.

## Calculation of Cell Counts and Transition Probabilities

The model's rules for determining cell counts within the five-by-five transition matrix are described in Appendix B. In applying these rules, one uses the information from the QC review to infer the prior-month status of active cases. We tested numerous alternative specifications of these rules, with specific attention to the resulting closure rates. One consideration was to minimize the number of instances of negative group-specific closure rates. Such anomalous values were present to some degree under all specifications, reflecting the variability of the weighted sample data as well as possible oversimplification in the modeling assumptions.

We applied the procedures in Appendix B in calculating the cell entries within the five-byfive matrix nationally for 1998 through 2001. For 2001, these estimates are shown in Exhibit 9. (Appendix Exhibits C-1 through C-4 show the estimates for each of the four years). We then used the estimated cell counts to compute the associated transition probabilities for each period. These are shown in Exhibit 10 for 2001 (and in Appendix Exhibits C-5 through C-8 for all four estimated years). For any given year, the transition probabilities in each row sum to 1 , as they are computed by dividing each cell count by its corresponding row total. To recall, each transition probability $\mathrm{p}_{\mathrm{ij}}$ indicates the probability that a household of status i in the current month will occupy status j in the next month.

## Derivation of Row Totals

The row totals in the basic transition matrix ( $\mathrm{R}_{1}$ through $\mathrm{R}_{5}$ in Exhibit 6) represent the current-month (March) distribution of households. Row totals cannot be directly estimated from the QC data, because cases that have just closed (the first entry in each row) are not observed in the next-month (April) QC data. Some simplifying assumption is thus required to derive these row totals, which can then in turn be used to calculate the number of case closures.

In this section, we discuss at some length the approach used in setting the row totals. The modeling assumptions here deserve attention, for several reasons. First, there are a number of different approaches that one can take, with no compelling logic pointing toward any single correct method. Second, to the extent that different approaches yield different closure rates, the model's estimates will vary according to one's choice of method.

In principle, one would ideally want the marginal totals in the basic transition matrix to reflect both (a) the change from one month to the next in the total population of households and (b) the shift from one month to the next in the distribution of households among the five groups, within the total population count. The approach adopted here is to derive the row totals in such a way as to incorporate the month-to-month distributional shift, as described below, but not the month-to-month population change. The assumption of a constant total population within a given fiscal year is only a very slight abstraction from reality, as the monthly trend rate of national population growth during this period was 0.1 percent.

The model assumes that, over the course of a fiscal year, the change in the distribution of households takes place in twelve equal monthly steps. The month-to-month change in the distribution of households across the five groups is computed as one-twelfth of the observed year-to-year change in the household distribution.

Exhibit 9: Cell Counts, 2001

|  | Next-month ("April") status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Current-month ("March") status | $\begin{array}{r} \text { Non- } \\ \text { participating } \end{array}$ | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total (row) |

Total households (in thousands)

| Nonparticipating | 98,385 | 574 | 57 | 0 | 0 | 99,015 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 485 | 4,302 | 337 | 633 | 56 | 5,814 |
| Ongoing error | 23 | 380 | 360 | 2 | 63 | 829 |
| Expiring correct | 72 | 509 | 58 | 0 | 0 | 640 |
| Expiring error | 51 | 56 | 14 | 0 | 0 | 121 |
| Total (column) | 99,015 | 5,821 | 827 | 636 | 119 | 106,418 |
|  | Households with earnings (in thousands) |  |  |  |  |  |
| Nonparticipating | 83,049 | 190 | 29 | 0 | 0 | 83,268 |
| Ongoing correct | 141 | 846 | 160 | 211 | 29 | 1,387 |
| Ongoing error | 43 | 159 | 102 | , | 27 | 331 |
| Expiring correct | 10 | 169 | 33 | 0 | 0 | 212 |
| Expiring error | 25 | 28 | 4 | 0 | 0 | 58 |
| Total (column) | 83,268 | 1,391 | 329 | 212 | 56 | 85,257 |
|  | Households without earnings (in thousands) |  |  |  |  |  |
| Nonparticipating | 15,318 | 384 | 27 | 0 | 0 | 15,729 |
| Ongoing correct | 344 | 3,456 | 177 | 422 | 27 | 4,427 |
| Ongoing error | 0* | 221 | 259 | 1 | 36 | 517 |
| Expiring correct | 60 | 340 | 25 | 0 | 0 | 425 |
| Expiring error | 25 | 28 | 10 | 0 | 0 | 63 |
| Total (column) | 15,747 | 4,430 | 498 | 423 | 63 | 21,161 |

Note: Row (or column) entries may not sum to the indicated row (or column) total due to rounding. Asterisk (*) indicates an imputed zero value.

Exhibit 10: Transition Probabilities, 2001

| Current-month <br> ("March") status | Next-month ("April") status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Total households |  |  |  |  |  |
| Nonparticipating | 0.994 | 0.006 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.083 | 0.740 | 0.058 | 0.109 | 0.010 | 1.000 |
| Ongoing error | 0.028 | 0.459 | 0.435 | 0.003 | 0.076 | 1.000 |
| Expiring correct | 0.113 | 0.796 | 0.091 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.420 | 0.462 | 0.119 | 0.000 | 0.000 | 1.000 |
|  | Households with earnings |  |  |  |  |  |
| Nonparticipating | 0.997 | 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.102 | 0.610 | 0.116 | 0.152 | 0.021 | 1.000 |
| Ongoing error | 0.128 | 0.480 | 0.307 | 0.004 | 0.081 | 1.000 |
| Expiring correct | 0.049 | 0.795 | 0.157 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.440 | 0.484 | 0.076 | 0.000 | 0.000 | 1.000 |
|  | Households without earnings |  |  |  |  |  |
| Nonparticipating | 0.974 | 0.024 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.078 | 0.781 | 0.040 | 0.095 | 0.006 | 1.000 |
| Ongoing error | 0.000* | 0.445 | 0.520 | 0.002 | 0.072 | 1.039 |
| Expiring correct | 0.141 | 0.800 | 0.059 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.395 | 0.447 | 0.158 | 0.000 | 0.000 | 1.000 |

[^11]To illustrate the linear interpolation method used to calculate the row totals for the annual national estimates, we describe below how we calculated the national row totals for the 2001 matrix; these are the marginal row totals shown in the top panel of Exhibit 9. As previously noted, we refer to these row totals as the "March 2001" estimates, as they represent the distribution of households one month before the "April 2001" column totals.

Columns (a) through (f) of Exhibit 11 show the calculations for this example. In this exhibit, the weighted household counts are not rounded. The corresponding numbers in Exhibit 8, as elsewhere in the report, are rounded to the nearest thousand.

The steps in the calculation of the March 2001 row totals are as follows:

- Start with the household counts determined in the prior year, 2000. These "April 2000" counts (by group) are the column totals from the top panel of Exhibit C-3, as repeated in column (a) of Exhibit 11. For the four subgroups of active cases, these counts are obtained directly from the 2000 QC data. The total household count for April 2000 ( 104.705 million, from Census data) is used to derive the April 2000 count of nonparticipating households, as a residual.
- Compute the 2000 ("April 2000") percentage distribution of households across the five groups. This is shown in column (b) of Exhibit 11.
- Apply the percentage distribution of households in column (b) to the national population base for April 2001 (106.418 million, from Census data), thus deriving the counts for each group that would result if the percentage distribution of households had remained unchanged from 2000 to 2001 (i.e., from April 2000 to April 2001). These counts are shown in column (c) of Exhibit 11. This step is in keeping with the model's focus on the aggregate participation rate (versus the number of active cases) and on the case error rate (versus the number of error cases).
- Column (d) shows the household counts for 2001 ("April 2001"), as shown in the column totals in the top panel of Exhibit 9. Compute the average monthly change for each group as $1 / 12$ of the difference between the "April 2001" household count in column (d) and the population-adjusted "April 2000" household count in column (c). For each group, this average monthly change is shown in column (e). It can be either a positive or negative value. By definition, the average monthly changes for the five groups sum to zero.

Exhibit 11: Calculating "March 2001" Row Totals for FY 2001 Transition Matrix

| Household group | $\begin{array}{r} \text { FY2000 } \\ \text { ("April 2000") } \\ \text { count } \end{array}$ | $\begin{array}{r} \text { FY2000 } \\ \text { ("April 2000") } \\ \text { distribution } \end{array}$ | FY2000 distribution, applied to FY2001 population | FY2001 ("April 2001") count | Average monthly change, FY2000 to FY2001 | FY2001 ("March 2001") row total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | (a) | (b) | (c) | (d) | (e) | (f) |
| Nonparticipating | 97,415,884 | 93.04\% | 99,009,632 | 99,015,357 | 477 | 99,014,880 |
| Ongoing correct | 5,642,513 | 5.39\% | 5,734,826 | 5,821,020 | 7,183 | 5,813,837 |
| Ongoing error | 832,461 | 0.80\% | 846,080 | 827,098 | -1,582 | 828,680 |
| Expiring correct | 676,094 | 0.65\% | 687,155 | 635,508 | -4,304 | 639,812 |
| Expiring error | 138,048 | 0.13\% | 140,306 | 119,017 | -1,774 | 120,791 |
| Total | 104,705,000 | 100.00\% | 106,418,000 | 106,418,000 | 0 | 106,418,000 |
|  | See column totals in top panel of | $=(\mathrm{a} / 104,705,000$ | $=(\mathrm{b})^{*} 106,418,000$ | See column totals in top panel of | $=[(\mathrm{d})$-(c) $] / 12$ | $=(\mathrm{d})-(\mathrm{e})$ <br> See row totals in |
|  |  |  |  |  |  | top panel o Exhibit C-4 |

- For each group, compute the marginal row total for 2001 ("March 2001") by subtracting the average monthly change from the 2001 ("April 2001") count. The result of this calculation is shown in column (f). These counts, rounded to the nearest thousand, then appear as the marginal row totals in the top panel of Exhibit 9. ${ }^{17}$

The marginal row totals provide the basis for calculating the number of monthly case closures for each of the defined subgroups of active cases (ongoing correct, ongoing error, expiring correct, and expiring error). Case closures are not observed in the active case QC data and also cannot be derived by group from negative action QC data. The marginal row totals allow one to compute the number of closures as the row residual. These closure counts comprise the first column of the basic transition matrix. Although the cell counts elsewhere in the five-by-five matrix are not contingent on the marginal row totals, the row totals form the denominators for computing the transition probabilities in each row.

## Long-Term Outcomes Implied by the Transition Matrix

One advantage of the Markov modeling approach is that the estimated transition matrix can be used to compute the projected long-term distribution of households across the specified groups. This projected steady-state distribution is the system outcome that would ultimately occur if the process of month-to-month transitions were to continue indefinitely. One can expect that shifts in the underlying pattern of participation and error, such as those associated with more frequent recertification, will play themselves out progressively over time. For this reason, it is useful to assess such shifts in terms of the long-run outcomes.

In the previous chapter, the equilibrium value of the aggregate participation rate was derived for the two-group model as $a /(a+b)$, where $a$ is the case opening rate and $b$ is the case closure rate. We show below the formulas for deriving the equilibrium values for the aggregate participation rate ( $\mathrm{p}^{*}$ ) and the case error rate ( $\mathrm{r}^{*}$ ) from the five-group model, using the notation introduced in Exhibit 6. These formulas express each of these long-term outcomes as a function of the associated opening rates and closure rates. In turn, these opening rates and closure rates are computed from cell counts contained in the basic transition matrix.

The formula below for the long-term participation rate is a direct extension of the calculation shown earlier for the two-group model. The long-term case error rate is derived below as a ratio of two proportions, both expressed as shares of the total household population. One ( $\mathrm{e}^{*}$ ) is the long-term share of the population consisting of error cases; the other ( $\mathrm{p}^{*}$ ) is the long-term participation rate.

[^12]$\begin{aligned} \text { Long-term aggregate participation rate } & =\quad \mathrm{p}^{*} \\ & =\quad \mathrm{a} /(\mathrm{a}+\mathrm{b})\end{aligned}$

$$
\begin{aligned}
\text { Where } \mathrm{a} & =\text { case opening rate } \\
& =\left(\mathrm{Q}_{12}+\mathrm{Q}_{13}\right) / \mathrm{R}_{1} \\
\mathrm{~b} & =\text { case closure rate } \\
& =\left(\mathrm{Q}_{21}+\mathrm{Q}_{31}+\mathrm{Q}_{41}+\mathrm{Q}_{51}\right) /\left(\mathrm{R}_{2}+\mathrm{R}_{3}+\mathrm{R}_{4}+\mathrm{R}_{5}\right)
\end{aligned}
$$

Long-term case error rate $=\quad r^{*}$

$$
=\quad \mathrm{e}^{*} / \mathrm{p}^{*}
$$

Where $\mathrm{e}^{*}=$ aggregate error rate

$$
=\quad \mathrm{c} /(\mathrm{c}+\mathrm{d})
$$

$\mathrm{c}=$ error opening rate

$$
=\left(\mathrm{Q}_{13}+\mathrm{Q}_{23}+\mathrm{Q}_{24}+\mathrm{Q}_{43}\right) /\left(\mathrm{R}_{1}+\mathrm{R}_{2}+\mathrm{R}_{4}\right)
$$

$\mathrm{d}=$ error closure rate

$$
=\left(\mathrm{Q}_{31}+\mathrm{Q}_{32}+\mathrm{Q}_{34}+\mathrm{Q}_{51}+\mathrm{Q}_{52}\right) /\left(\mathrm{R}_{3}+\mathrm{R}_{5}\right)
$$

Using these formulas, Exhibit 12 shows the long-term participation rate and case error rate implied by the 2001 transition matrices for total households, households with earnings, and households without earnings. For each segment of the population, the matrices imply a slight downward or stable path for both the participation rate and the error rate. ${ }^{18}$

[^13]Exhibit 12: Projected Long-Term Outcomes, Based on 2001 Matrix

|  | Observed <br> FY 2001 <br> value (\%) | Projected <br> long-term <br> value (\%) |
| :--- | ---: | ---: |
| Total households |  |  |
| $\quad$ Aggregate participation rate | 6.96 | 6.95 |
| $\quad$ Case error rate | 12.78 | 12.72 |
| Households with earnings |  |  |
| $\quad$ Aggregate participation rate | 2.33 | 2.33 |
| $\quad$ Case error rate | 19.37 | 19.19 |
| Households without earnings |  |  |
| $\quad$ Aggregate participation rate | 25.58 | 24.86 |
| $\quad$ Case error rate | 10.36 | 10.28 |

Source: Based on formulas shown in text, using the transition matrices derived for 2001 (see Exhibit 9).

## Chapter Four: National Trends in Program Participation and Error Rates

Recent national trends in food stamp participation and error are examined in this chapter. The annually estimated model is used here to examine the patterns that underlie year-to-year changes in the rate at which households participate in the program (the aggregate participation rate) and the rate at which active cases are incorrectly paid (the case error rate).

## Underlying Trends in Participation Rates

The model produces a series of statistics pertaining to the monthly size, composition, and dynamics of the active food stamp caseload. The first of these is a general indicator of food stamp receipt within the household population, in a given month. (The definitions below use the notation from Exhibit 6.)

- Aggregate participation rate
$=$ percentage of all households that are active cases
$=\left(\mathrm{C}_{2}+\mathrm{C}_{3}+\mathrm{C}_{4}+\mathrm{C}_{5}\right) / \mathrm{Q}$
For 1997 through 2001, Exhibit 13 shows the average monthly number of active food stamp cases, the total number of U.S. households, and the corresponding aggregate participation rate. (Recall that this is distinct from the "conditional" participation rate that shows the percentage of eligible households that participate in the program.) Estimates are shown separately for total households, households with earnings, and households without earnings.

There are three caseload share parameters, summing to one, that indicate the proportion of the active caseload comprised by first-month cases, ongoing cases, or expiring cases:

- Caseload share: first-month cases
$=$ percentage of active cases that are newly certified
$=\left(\mathrm{Q}_{12}+\mathrm{Q}_{13}\right) /\left(\mathrm{C}_{2}+\mathrm{C}_{3}+\mathrm{C}_{4}+\mathrm{C}_{5}\right)$
- Caseload share: ongoing cases
$=$ percentage of active cases that are subject to interim action
$=\left(\mathrm{C}_{2}+\mathrm{C}_{3}-\mathrm{Q}_{12}-\mathrm{Q}_{13}\right) /\left(\mathrm{C}_{2}+\mathrm{C}_{3}+\mathrm{C}_{4}+\mathrm{C}_{5}\right)$
- Caseload share: expiring cases
$=$ percentage of active cases that are subject to recertification
$=\left(\mathrm{C}_{4}+\mathrm{C}_{5}\right) /\left(\mathrm{C}_{2}+\mathrm{C}_{3}+\mathrm{C}_{4}+\mathrm{C}_{5}\right)$

Exhibit 13: Aggregate Participation Rates, 1997-2001

|  | 1997 | 1998 | 1999 | 2000 | 2001 |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Total households |  |  |  |  |
| Number of food stamp cases (thousands) | 9,393 | 8,196 | 7,612 | 7,289 | 7,403 |
| Number of U.S. households (thousands) | 101,018 | 102,528 | 103,874 | 104,705 | 106,418 |
| Aggregate participation rate (\%) | 9.3\% | 8.0\% | 7.3\% | 7.0\% | 7.0\% |
|  | Households with earnings |  |  |  |  |
| Number of food stamp cases (thousands) | 2,273 | 2,158 | 2,048 | 1,985 | 1,989 |
| Number of U.S. households (thousands) | 79,790 | 81,248 | 82,611 | 84,184 | 85,257 |
| Aggregate participation rate (\%) | 2.8\% | 2.7\% | 2.5\% | 2.4\% | 2.3\% |
|  | Households without earnings |  |  |  |  |
| Number of food stamp cases (thousands) | 7,120 | 6,038 | 5,564 | 5,304 | 5,414 |
| Number of U.S. households (thousands) | 21,228 | 21,280 | 21,263 | 20,521 | 21,161 |
| Aggregate participation rate (\%) | 33.5\% | 28.4\% | 26.2\% | 25.8\% | 25.6\% |

Sources and notes: Number of food stamp cases-weighted count of participating cases, from the analysis sample shown in Exhibit 7.
Number of U.S. households-U.S. Department of Commerce, Bureau of the Census, Current Population Survey in March of each year.
(See www.census.gov/hhes/income/dinctabs.html.)
Aggregate participation rate-number of food stamp cases divided by the corresponding total number of U.S. households.

Exhibit 14 shows the annual estimates of these parameters, for 1998 through 2001. Over this period there was a slight increase in the caseload share associated with first-month cases and a slight decline in the share associated with expiring cases. These trends occurred both for households with earnings and without earnings. For both types of households, the caseload share associated with ongoing cases remained stable and large (more than three-fourths of active cases). In later explaining national error trends, it will be important to recognize that ongoing cases comprise so large a segment of the active caseload.

## Exhibit 14: Caseload Shares, 1998-2001

|  | 1998 | 1999 | 2000 | 2001 |
| :---: | :---: | :---: | :---: | :---: |
|  | Total households |  |  |  |
| Caseload share (\%): |  |  |  |  |
| First-month cases | 7.0 | 7.6 | 7.8 | 8.5 |
| Ongoing cases | 81.3 | 81.5 | 81.0 | 81.3 |
| Expiring cases | 11.8 | 10.8 | 11.2 | 10.2 |
|  | Households with earnings |  |  |  |
| Caseload share (\%): |  |  |  |  |
| First-month cases | 8.6 | 9.0 | 10.1 | 11.0 |
| Ongoing cases | 76.2 | 76.4 | 74.3 | 75.5 |
| Expiring cases | 15.2 | 14.6 | 15.6 | 13.5 |
|  | Households without earnings |  |  |  |
| Caseload share (\%): |  |  |  |  |
| First-month cases | 6.4 | 7.1 | 7.0 | 7.6 |
| Ongoing cases | 83.1 | 83.5 | 83.5 | 83.4 |
| Expiring cases | 10.5 | 9.4 | 9.5 | 9.0 |

One can also define the closure rates, indicating the percentage of cases that exit the active caseload in each month:

- Closure rate: total cases
$=$ closure rate among all active cases
$=\left(\mathrm{Q}_{21}+\mathrm{Q}_{31}+\mathrm{Q}_{41}+\mathrm{Q}_{51}\right) /\left(\mathrm{R}_{2}+\mathrm{R}_{3}+\mathrm{R}_{4}+\mathrm{R}_{5}\right)$
- Closure rate: ongoing cases
$=$ closure rate among active cases subject to interim action
$=\left(\mathrm{Q}_{21}+\mathrm{Q}_{31}\right) /\left(\mathrm{R}_{2}+\mathrm{R}_{3}\right)$
- Closure rate: expiring cases
$=$ closure rate among active cases subject to recertification
$=\left(\mathrm{Q}_{41}+\mathrm{Q}_{51}\right) /\left(\mathrm{R}_{4}+\mathrm{R}_{5}\right)$

Exhibit 15 shows the estimated closure rates provided by the model. As noted earlier, the monthly closure rates for active cases are subject to considerable sampling error, as they are computed as the residual of sample-estimated values in each row of the matrix. The estimated monthly closure rate for total active cases ranged between 8.2 and 8.5 percent during this period. Estimated closure rates were somewhat higher for households with earnings ( 9.5 to 12.3 percent) than for households without earnings ( 7.4 to 7.9 percent).

## Exhibit 15: Closure Rates, 1998-2001

|  | 1998 | 1999 | 2000 | 2001 |
| :---: | :---: | :---: | :---: | :---: |
|  | Total households |  |  |  |
| Closure rate (\%): |  |  |  |  |
| Total cases | 8.4 | 8.3 | 8.2 | 8.5 |
| Ongoing cases | 4.9 | 6.4 | 6.0 | 7.6 |
| Expiring cases | 33.6 | 24.5 | 26.2 | 16.2 |
|  | Households with earnings |  |  |  |
| Closure rate (\%): |  |  |  |  |
| Total cases | 12.3 | 9.5 | 10.5 | 11.0 |
| Ongoing cases | 9.7 | 7.1 | 8.0 | 10.7 |
| Expiring cases | 27.2 | 23.4 | 23.9 | 13.0 |
|  | Households without earnings |  |  |  |
| Closure rate (\%): |  |  |  |  |
| Total cases | 7.9 | 7.9 | 7.4 | 7.9 |
| Ongoing cases | 4.4 | 6.1 | 5.3 | 7.0 |
| Expiring cases | 36.9 | 24.7 | 27.5 | 17.4 |

For total active cases, the model yields monthly closure rates higher than those for June 2000 shown in Food Stamp Program Access Study: Local Office Policies and Practices. ${ }^{19}$ Compared to the estimates here for fiscal year 2000, Bartlett et al. found the closure rate to be 5.4 percent for total cases (versus the 8.2 percent above), 2.9 percent for ongoing cases (versus the 6.0 percent above), and 21.9 percent for expiring cases (versus the 26.2 percent above).

[^14]The higher rates here for total closures and interim closures appear to reflect some overstatement in the number of first-month (newly certified) cases and a corresponding understatement in the number of second-month (ongoing) cases and expiring cases. As noted earlier, the share of the caseload comprised by first-month cases was estimated here at 7.8 percent in FY 2000 (compared to 4.9 percent for Bartlett et al.). This higher-than-expected proportion of new openings may result from the way in which QC reviewers enter information on the timing of the initial certification in relation to the review month. Some of these cases (now classified in cells $\mathrm{Q}_{12}$ and $\mathrm{Q}_{13}$ of Exhibit 6) should perhaps be classified as in their second month and thus as ongoing cases. Any misclassifications of this type have the effect of raising the rate of new openings, the total closure rate, and the interim closure rate. In other instances, cases that are overdue for recertification (and should be classified as expiring) may be misclassified as first-month cases, given the instructions to QC reviewers for entering information on "sample month in certification." 20

Can any other data be used to indicate whether the closure rates estimated here are indeed higher than one should have expected? For any given fiscal year, it is possible to derive an "implied monthly closure rate" based on information regarding the monthly active caseload and the number of monthly case openings. To do this, one starts with the following identity: the net monthly change in active cases equals monthly case openings minus monthly case closings. Rearranging terms, it follows that: monthly case closings equal monthly case openings minus the net monthly change in active cases.

We have applied this relationship to compute implied monthly closure rates for 1998 through 2001, using national administrative data from FNS on monthly caseloads and also using the counts of monthly case openings calculated in this study from the annual QC data. As shown in Exhibit 16, the implied monthly closure rates are in the range of 7.7 to 8.0 percent during this period, compared to 8.2 to 8.5 percent for the model. This suggests that the monthly closure rates provided by the model may be somewhat overstated, but not to the degree suggested by the comparison with Abt's other recently completed study. We acknowledge that the implied closure rates calculated in Exhibit 16 do not provide a truly independent test, as they rely on the counts of case openings computed from the QC data analyzed here. To our knowledge, however, there is no independent source of data on food stamp case openings.

[^15]Exhibit 16: Monthly Closure Rates Implied by Year-to-Year Change in National Caseload

| Fiscal Year | Monthly caseload at start of current year | Monthly caseload at start of next year | Change in caseload over 12 months | Average monthly caseload change | Average monthly openings | Implied average monthly closings | Average monthly caseload | Implied monthly closure rate |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | (a) | (b) | (c) | (d) | (e) | (f) | (g) | (h) |
|  |  |  | $=(\mathrm{b})-(\mathrm{a})$ | $=(\mathrm{c}) / 12$ |  | $=(\mathrm{e})$-(d) |  | $=(\mathrm{f}) /(\mathrm{g})$ |
| 1998 | 8,658,521 | 7,858,938 | -799,583 | -66,632 | 571,574 | 638,206 | 8,248,741 | 0.077 |
| 1999 | 7,858,938 | 7,440,073 | -418,865 | -34,905 | 580,617 | 615,522 | 7,668,372 | 0.080 |
| 2000 | 7,440,073 | 7,315,526 | -124,547 | -10,379 | 571,395 | 581,774 | 7,324,628 | 0.079 |
| 2001 | 7,315,526 | 7,812,305 | 496,779 | 41,398 | 630,323 | 588,925 | 7,446,981 | 0.079 |

Sources:
(a) Food and Nutrition Service, National Data Bank, caseload in first month of current fiscal year.
(b) Food and Nutrition Service, National Data Bank, caseload in first month of next fiscal year.
(e) Abt Associates, Exhibits C-1 through C-4 of this report, sum of cell counts $\mathrm{Q}_{12}$ and $\mathrm{Q}_{13}$ for total households.
(g) Food and Nutrition Service, National Data Bank, average caseload over the twelve months of current fiscal year.

## Underlying Trends in Error Rates

Study findings are derived from the five-by-five transition matrices calculated nationwide and by state, using the annual QC data. The matrices are used to compute a series of parameters that underlie the total case error rate. These parameters are defined below, using the notation from Exhibit 6 relating to cell counts $\left(\mathrm{Q}_{\mathrm{ij}}\right)$, column totals $\left(\mathrm{C}_{\mathrm{j}}\right)$, and row totals $\left(\mathrm{R}_{\mathrm{i}}\right)$ in the transition matrix.

- Total error rate
$=$ case error rate among all active cases
$=\left(\mathrm{C}_{3}+\mathrm{C}_{5}\right) /\left(\mathrm{C}_{2}+\mathrm{C}_{3}+\mathrm{C}_{4}+\mathrm{C}_{5}\right)$
- First-month error rate
$=$ case error rate among first-month (newly certified) cases
$=\mathrm{Q}_{13} /\left(\mathrm{R}_{1}-\mathrm{Q}_{11}\right)$
- Next-month error rate: ongoing correct cases
$=$ next-month case error rate among current-month ongoing correct cases
$=\left(\mathrm{Q}_{23}+\mathrm{Q}_{25}\right) /\left(\mathrm{R}_{2}-\mathrm{Q}_{21}\right)$
- Next-month error rate: ongoing error cases
$=$ next-month case error rate among current-month ongoing error cases
$=\left(\mathrm{Q}_{33}+\mathrm{Q}_{35}\right) /\left(\mathrm{R}_{3}-\mathrm{Q}_{31}\right)$
- Next-month error rate: expiring correct cases
$=$ next-month case error rate among current-month expiring correct cases
$=\mathrm{Q}_{43} /\left(\mathrm{R}_{4}-\mathrm{Q}_{41}\right)$
- Next-month error rate: expiring error cases
$=$ next-month case error rate among current-month expiring error cases
$=\mathrm{Q}_{53} /\left(\mathrm{R}_{5}-\mathrm{Q}_{51}\right)$

The denominator for each error rate is the number of cases that are active in the next month, within the corresponding caseload group. ${ }^{21}$

[^16]The first-month error rate is a measure of payment accuracy at initial certification. The nextmonth error rates for ongoing cases indicate the effectiveness of interim actions at preventing errors (among ongoing correct cases) and at detecting and correcting errors (among ongoing error cases). Similarly, the next-month error rates for expiring cases indicate the effectiveness of recertifications at preventing errors (among cases that are correct as they enter recertification) and at detecting and correcting errors (among cases that are in error as they enter recertification).

In general, the modeling approach yields parameter estimates at the national level that either remain stable or trend progressively upward or downward over the period 1998 to 2001. Exhibit 17 shows the above-defined key error indicators by year, for total households and for the two subgroups (households with and without earnings). The major descriptive findings are as follows:

- Total error rate-On a consistently measured basis, the total error rate steadily declined during these years, from 16.7 percent in 1998 to 12.8 percent in 2001 for the total caseload. ${ }^{22}$ As typically observed in QC data, the error rate for cases with earnings (19.4 percent in 2001) was consistently about twice as high as that for cases without earnings (10.4 percent in 2001).
- First-month error rate-The error rate for newly certified cases (reflecting the errors that occur at initial certification) declined markedly during these years (from 12.2 percent in 1998 to 9.0 percent in 2001 for total cases). This was partly responsible for the reduction in the case error rate, especially in light of the growing caseload share comprised by first-month cases (as explained below, from 7.0 percent in 1998 to 8.5 percent in 2001). As with the total error rate, the firstmonth error rate was about twice as high for cases with earnings (13.2 percent in 2001) as for cases without earnings ( 6.6 percent in 2001).

[^17]Exhibit 17: Error Rates, 1998-2001

|  | 1998 | 1999 | 2000 | 2001 |
| :---: | :---: | :---: | :---: | :---: |
|  | Total households |  |  |  |
| Total error rate (\%) | 16.7 | 14.7 | 13.3 | 12.8 |
| First-month error rate (\%): | 12.2 | 10.8 | 8.9 | 9.0 |
| Next-month error rate (\%): |  |  |  |  |
| Ongoing correct cases | 9.6 | 8.2 | 7.7 | 7.4 |
| Ongoing error cases | 75.5 | 61.8 | 57.0 | 52.5 |
| Expiring correct cases | 11.8 | 11.5 | 9.4 | 10.2 |
| Expiring error cases | 38.6 | 29.4 | 25.0 | 20.0 |
|  | Households with earnings |  |  |  |
| Total error rate (\%) | 25.9 | 23.3 | 21.5 | 19.4 |
| First-month error rate (\%) | 18.9 | 16.8 | 13.5 | 13.2 |
| Next-month error rate (\%): |  |  |  |  |
| Ongoing correct cases | 20.0 | 17.5 | 17.3 | 15.2 |
| Ongoing error cases | 66.1 | 54.6 | 50.5 | 44.8 |
| Expiring correct cases | 19.3 | 19.5 | 15.6 | 16.3 |
| Expiring error cases | 35.5 | 18.9 | 19.4 | 12.1 |
|  | Households without earnings |  |  |  |
| Total error rate (\%) | 13.4 | 11.5 | 10.3 | 10.4 |
| First-month error rate (\%) | 9.0 | 8.1 | 6.5 | 6.6 |
| Next-month error rate (\%): |  |  |  |  |
| Ongoing correct cases | 6.4 | 5.3 | 4.7 | 5.0 |
| Ongoing error cases | 80.4 | 66.0 | 61.3 | 57.1 |
| Expiring correct cases | 7.7 | 7.3 | 5.8 | 6.8 |
| Expiring error cases | 40.7 | 36.4 | 30.6 | 26.3 |

- Next-month error rate: ongoing correct cases-During these years there was a decline in the rate at which, in the midst of a certification period, correct cases became in error the following month. This might reflect greater month-to-month stability in household circumstances among food stamp cases. It also may reflect some improvement in the extent to which interim actions prevented errors from occurring. This rate dropped from 9.6 percent in 1998 to 7.4 percent in 2001. The decline was especially pronounced among cases with earnings, from 20.0 percent in 1998 to 15.2 percent in 2001. Even with this lowering, the rate for earnings cases remained approximately three times the rate for cases without earnings ( 5.0 percent in 2001).
- Next-month error rate: ongoing error cases-There was also a marked decline in the rate at which ongoing error cases remained in error the next month, a measure of the extent to which interim actions serve to detect and correct errors). This rate was 52.5 percent for total cases in 2001, and it was much lower for cases with earnings ( 44.8 percent) than for cases without earnings ( 57.1 percent). For both caseload segments, these rates dropped by more than 20 percentage points between 1998 and 2001. It is interesting to note that errors among cases with earnings appear less likely to persist from one month to the next (in comparison to the errors among cases without earnings). It may be that reporting systems are better at capturing income changes than other changes in household circumstances.
- Next-month error rate: expiring correct cases-The percentage of correct cases that became in error upon recertification remained stable during this period, equaling 10.2 percent for total cases in 2001 ( 16.3 percent for cases with earnings and 6.8 percent for cases without earnings). This measure reflects the extent to which errors are newly created in the course of a recertification. Errors can arise at recertification if, for instance, a caseworker misapplies policies in acting on new information about the household's circumstances. It is seemingly for this reason (i.e., agency-related errors at recertification) that the onset of error occurs at a higher rate among expiring correct cases than among ongoing correct cases (10.2 percent versus 7.4 percent for 2001), as recertifications presumably serve to reduce the onset of client-related errors. Note also that the rate of error onset is higher among newly-recertified cases than among those initially certified (10.2 percent versus 9.0 percent in 2001). This may reflect a greater degree of attention (and caseworker labor) devoted by program offices to initial certifications than to recertifications, on a per-case basis.
- Next-month error rate: expiring error cases-The percentage of expiring error cases that remained in error in the course of recertification was 20.0 percent in 2001, showing a dramatic reduction from 38.6 percent in 1998. This "error survival rate at recertification" (the rate at which errors escape detection and correction at recertification) was less than half as large for cases with earnings (12.1 percent) than for cases without earnings ( 26.3 percent). As was noted above for ongoing cases, errors among cases without earnings tend to be more persistent than errors among cases with earnings.

To summarize these results, the reduction in the national case error rate from 16.7 percent in 1998 to 12.8 percent in 2001 reflected improvements both for cases with earnings and for cases without earnings. For both caseload segments, one minor factor was the drop in error rates at initial certification, accentuated by the fact that first-month cases came to comprise an increasing share of the caseload. Far more important, however, was the reduction in nextmonth error rates for ongoing cases. As noted earlier, such cases comprise the bulk of the
active caseload. The major contributing factor was the dramatic decline in the next-month error rate for ongoing error cases. It appears that interim action procedures became much better at detecting and correcting errors between formal case actions. Also noteworthy, but of lesser importance in explaining the national error trend, was the drop in the next-month error rate for expiring error cases. Both for cases with and without earnings, recertification procedures appeared to capture and correct more errors.

To the extent that there was a reduction in the next-month error rates for ongoing cases, it is difficult to know whether this occurred more as a result of (a) the adoption by states of new client reporting systems that were more error-tolerant or (b) the improved administrative performance of states under client reporting systems that remained largely unchanged. In 1998-2001, the Food and Nutrition Service increasingly granted waivers allowing states to adopt reporting systems that were more forgiving. ${ }^{23}$ For instance, quarterly or semiannual reporting systems tended to ease the burdens upon both clients and agencies to respond to household changes affecting the monthly benefit, by extending the time interval allowed for reporting circumstantial changes. States could also adopt "status reporting", which limited the household's obligation to report changes in earnings to those situations involving major shifts in employment status. However, not until 2001 did the Food and Nutrition Service collect systematic information on the client reporting systems used by states, making it impossible to assess the effect of changes in reporting systems on error rate trends.

The comparison of error patterns between cases with and without earnings is informative. The higher case error rate among cases with earnings results from the higher probability that such cases will be in error at initial certification or (if correct at intake) will later fall into error. Errors among cases with earnings tend not to persist as long, however, as such error cases are more likely (than error cases without earnings) to leave the caseload during interim months or at recertification. Stated otherwise, errors tend to both start and end at a higher rate among cases with earnings than among cases without earnings.

[^18]
## Chapter Five: State-by-State Error Patterns

State-by-state patterns are discussed in this chapter. The model can be used to help explain the variation in states' error rates, by diagnosing whether a state's errors are attributable to errors among first-month cases (at initial certification), among ongoing cases (at interim action), or among expiring cases (at recertification). This information is especially important for planning corrective actions, so that one can focus attention on the phase of the administrative process that is most responsible for errors.

Historically, food stamp error rates have shown substantial interstate variation. For FY 2001, as shown in Exhibit 18, the total case error rate ranged from 4.46 percent in South Dakota to 31.12 percent in California.

## Deriving State-Specific Models

We calculated the model for each state and the District of Columbia using the state's pooled QC sample for the period 1998-2001. Initially, state-specific single-year models were computed. For states with annual QC samples of fewer than 800 cases, the majority of all states, sampling variability resulted in considerable year-to-year changes in the estimated transition probabilities. For this reason, and because of the desire to compare the model's results across all states, it was decided to pool the data across years for each state.

As with the single-year estimates at the national level, row totals were first derived for each state's multi-year model. In the state-specific models, as in the national estimates, the row totals ( $\mathrm{R}_{1}$ through $\mathrm{R}_{5}$ ) represent the current-month distribution of households. As with the national estimates, the row totals assume that the change in the distribution of households takes place in twelve equal monthly steps for each year. For each state, the average month-to-month changes for each of the four years 1998-2001 are themselves averaged and are then used to derive the row totals for the state's multi-year estimates. Stated otherwise, we assume that the change in the distribution of households over the four-year period takes place in 48 equal monthly steps.

Exhibit 19 shows for each state the distribution of active cases between households with earnings and households without earnings. For any given state, the respective sizes of these caseload segments will affect the role played by the respective error rate parameters for cases with and without earnings in influencing the state's total error rate.

Exhibits 20, 21, and 22 show the key error findings from the state-by-state models. These exhibits display the previously defined case error indicators, for total households (Exhibit 20), households with earnings (Exhibit 21), and households without earnings (Exhibit 22). In each exhibit, the state values of each error indicator that lie in the bottom quartile of the
distribution are shown in bold. These states are the exemplary performers with respect to that error indicator.

The underlying cell counts and transition probabilities are shown by state in Appendix D.

Exhibit 18: Case Error Rates by State, Fiscal Year 2001

| State | Overpayment case error rate (\%) | Underpayment case error rate (\%) | Total case error rate (\%) |
| :---: | :---: | :---: | :---: |
| Alabama | 12.82 | 2.99 | 15.81 |
| Alaska | 13.61 | 5.38 | 18.99 |
| Arizona | 6.76 | 3.86 | 10.62 |
| Arkansas | 4.46 | 2.04 | 6.50 |
| California | 20.00 | 11.12 | 31.12 |
| Colorado | 10.78 | 4.70 | 15.48 |
| Connecticut | 9.79 | 5.29 | 15.08 |
| Delaware | 11.18 | 5.80 | 16.98 |
| District of Columbia | 12.42 | 5.70 | 18.12 |
| Florida | 10.16 | 4.34 | 14.50 |
| Georgia | 7.74 | 2.51 | 10.25 |
| Hawaii | 8.83 | 5.68 | 14.51 |
| Idaho | 8.37 | 4.28 | 12.65 |
| Illinois | 9.50 | 3.13 | 12.63 |
| Indiana | 7.83 | 3.84 | 11.67 |
| lowa | 8.87 | 2.38 | 11.25 |
| Kansas | 10.71 | 3.39 | 14.10 |
| Kentucky | 8.08 | 3.57 | 11.65 |
| Louisiana | 7.88 | 4.03 | 11.91 |
| Maine | 9.46 | 5.27 | 14.73 |
| Maryland | 9.01 | 4.00 | 13.01 |
| Massachusetts | 9.05 | 4.47 | 13.52 |
| Michigan | 12.21 | 6.30 | 18.51 |
| Minnesota | 5.69 | 3.05 | 8.74 |
| Mississippi | 4.93 | 2.63 | 7.56 |
| Missouri | 10.74 | 4.18 | 14.92 |
| Montana | 11.32 | 3.70 | 15.02 |
| Nebraska | 11.83 | 3.23 | 15.06 |
| Nevada | 8.07 | 3.16 | 11.23 |
| New Hampshire | 9.75 | 4.18 | 13.93 |

## Exhibit 18: Case Error Rates by State, Fiscal Year 2001 (Continued)

|  | Overpayment <br> case error <br> rate (\%) | Underpayment <br> case error <br> rate (\%) | Total <br> case error <br> rate (\%) |
| :--- | ---: | ---: | ---: |
| State |  |  |  |
| New Jersey | 8.03 | 3.42 | 11.45 |
| New Mexico | 8.79 | 3.08 | 11.87 |
| New York | 7.06 | 6.26 | 13.32 |
| North Carolina | 7.75 | 2.10 | 9.85 |
| North Dakota | 4.69 | 3.19 | 7.88 |
| Ohio | 8.06 | 3.70 | 11.76 |
|  |  |  |  |
| Oklahoma | 10.29 | 3.33 | 13.62 |
| Oregon | 12.37 | 3.11 | 15.48 |
| Pennsylvania | 9.29 | 4.94 | 14.23 |
| Rhode Island | 6.44 | 4.19 | 10.63 |
|  |  |  |  |
| South Carolina | 5.37 | 2.81 | 8.18 |
| South Dakota | 3.62 | 0.84 | 4.46 |
| Tennessee | 8.35 | 2.51 | 10.86 |
| Texas | 5.30 | 2.58 | 7.88 |
| Utah | 12.12 | 4.27 | 16.39 |
| Vermont | 12.00 | 2.93 | 14.93 |
| Virginia | 7.61 | 4.64 | 12.25 |
| Washington | 8.97 | 3.67 | 12.64 |
| West Virginia | 8.40 | 2.40 | 10.80 |
| Wisconsin | 12.09 | 6.23 | 18.32 |
| Wyoming | 3.30 | 1.80 | 5.10 |
| U.S. Average |  |  |  |
|  | 9.55 |  | 14.11 |
|  |  |  |  |

Source: U.S. Department of Agriculture, Food and Nutrition Service, "Food Stamp Program Quality Control Annual Report, Fiscal Year 2001."

Notes:
U.S. average is weighted by state issuance and includes Guam and Virgin Islands.

Row entries may not sum to the indicated row total due to rounding.

Exhibit 19: Households With and Without Earnings, Caseload Shares by State, 1998-2001 Combined

| State | Share of caseload (\%) |  |  |
| :---: | :---: | :---: | :---: |
|  | Households with earnings | Households without earnings | Total |
| Alabama | 29.8 | 70.2 | 100.0 |
| Alaska | 32.6 | 67.4 | 100.0 |
| Arizona | 32.9 | 67.1 | 100.0 |
| Arkansas | 28.5 | 71.5 | 100.0 |
| California | 30.5 | 69.5 | 100.0 |
| Colorado | 29.8 | 70.2 | 100.0 |
| Connecticut | 12.3 | 87.7 | 100.0 |
| Delaware | 27.4 | 72.6 | 100.0 |
| District of Columbia | 10.4 | 89.6 | 100.0 |
| Florida | 25.4 | 74.6 | 100.0 |
| Georgia | 28.7 | 71.3 | 100.0 |
| Hawaii | 27.0 | 73.0 | 100.0 |
| Idaho | 41.2 | 58.8 | 100.0 |
| Illinois | 27.2 | 72.8 | 100.0 |
| Indiana | 27.9 | 72.1 | 100.0 |
| lowa | 32.1 | 67.9 | 100.0 |
| Kansas | 28.1 | 71.9 | 100.0 |
| Kentucky | 25.9 | 74.1 | 100.0 |
| Louisiana | 33.0 | 67.0 | 100.0 |
| Maine | 20.2 | 79.8 | 100.0 |
| Maryland | 21.6 | 78.4 | 100.0 |
| Massachusetts | 15.5 | 84.5 | 100.0 |
| Michigan | 31.2 | 68.8 | 100.0 |
| Minnesota | 22.0 | 78.0 | 100.0 |
| Mississippi | 28.8 | 71.2 | 100.0 |
| Missouri | 26.5 | 73.5 | 100.0 |
| Montana | 34.0 | 66.0 | 100.0 |
| Nebraska | 32.7 | 67.3 | 100.0 |
| Nevada | 21.1 | 78.9 | 100.0 |
| New Hampshire | 20.6 | 79.4 | 100.0 |

Exhibit 19: Households With and Without Earnings, Caseload Shares by State, 1998-2001 Combined (Continued)

| State | Share of caseload (\%) |  |  |
| :---: | :---: | :---: | :---: |
|  | Households with earnings | Households without earnings | Total |
| New Jersey | 15.8 | 84.2 | 100.0 |
| New Mexico | 32.2 | 67.8 | 100.0 |
| New York | 16.7 | 83.3 | 100.0 |
| North Carolina | 27.2 | 72.8 | 100.0 |
| North Dakota | 40.8 | 59.2 | 100.0 |
| Ohio | 22.8 | 77.2 | 100.0 |
| Oklahoma | 29.3 | 70.7 | 100.0 |
| Oregon | 31.0 | 69.0 | 100.0 |
| Pennsylvania | 26.1 | 73.9 | 100.0 |
| Rhode Island | 19.6 | 80.4 | 100.0 |
| South Carolina | 29.1 | 70.9 | 100.0 |
| South Dakota | 38.6 | 61.4 | 100.0 |
| Tennessee | 25.5 | 74.5 | 100.0 |
| Texas | 38.3 | 61.7 | 100.0 |
| Utah | 34.9 | 65.1 | 100.0 |
| Vermont | 23.7 | 76.3 | 100.0 |
| Virginia | 28.7 | 71.3 | 100.0 |
| Washington | 22.2 | 77.8 | 100.0 |
| West Virginia | 24.1 | 75.9 | 100.0 |
| Wisconsin | 32.3 | 67.7 | 100.0 |
| Wyoming | 42.6 | 57.4 | 100.0 |
| U.S. Average | 26.8 | 73.2 | 100.0 |

Source: Food stamp QC data by state, pooled over the period 1998-2001.

Exhibit 20: Case Error Indicators by State, 1998-2001 Combined: Total Households

| State | Total error rate | First-month error rate | Next-month error rate for: |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Ongoing correct cases | Ongoing error cases | Expiring correct cases |  |
| Alabama | 14.4 | 16.5 | 8.5 | 45.0 | 9.0 | 2.0 |
| Alaska | 22.4 | 18.1 | 20.9 | 13.9 | 30.5 | 0.0 |
| Arizona | 8.7 | 6.2 | 6.3 | 29.6 | 5.3 | 7.8 |
| Arkansas | 7.2 | 4.8 | 3.0 | 27.2 | 6.2 | 29.3 |
| California | 24.1 | 12.5 | 17.1 | 42.1 | 5.2 | 7.2 |
| Colorado | 13.2 | 8.5 | 5.8 | 46.7 | 9.0 | 29.0 |
| Connecticut | 16.7 | 13.7 | 6.8 | 62.3 | 8.6 | 31.4 |
| Delaware | 19.8 | 14.8 | 10.5 | 54.9 | 6.6 | 9.9 |
| District of Columbia | 16.6 | 9.0 | 9.3 | 50.1 | 12.8 | 7.4 |
| Florida | 13.7 | 11.0 | 8.7 | 47.4 | 5.7 | 3.3 |
| Georgia | 14.4 | 9.8 | 7.3 | 47.7 | 11.2 | 13.4 |
| Hawaii | 13.9 | 14.3 | 8.1 | 46.3 | 5.1 | 29.0 |
| Idaho | 12.2 | 10.3 | 5.9 | 48.3 | 9.8 | 19.0 |
| Illinois | 16.4 | 7.9 | 7.6 | 39.3 | 18.8 | 32.3 |
| Indiana | 10.2 | 9.2 | 4.2 | 43.3 | 8.3 | 14.5 |
| lowa | 12.4 | 10.8 | 7.1 | 25.4 | 14.6 | 22.7 |
| Kansas | 14.2 | 12.3 | 8.5 | 41.6 | 8.9 | 31.2 |
| Kentucky | 10.2 | 8.1 | 5.5 | 45.3 | 6.7 | 5.2 |
| Louisiana | 12.2 | 11.3 | 5.7 | 60.5 | 5.4 | 5.6 |
| Maine | 13.4 | 16.4 | 6.3 | 47.6 | 13.7 | 25.8 |
| Maryland | 17.4 | 12.4 | 6.9 | 68.3 | 6.3 | 20.6 |
| Massachusetts | 12.5 | 11.5 | 6.5 | 47.5 | 9.6 | 14.1 |
| Michigan | 21.9 | 14.2 | 8.5 | 68.4 | 8.6 | 22.4 |
| Minnesota | 7.3 | 3.8 | 3.6 | 14.3 | 2.5 | 34.5 |
| Mississippi | 8.3 | 6.4 | 5.2 | 15.6 | 7.5 | 20.3 |
| Missouri | 11.9 | 7.8 | 5.8 | 50.5 | 10.6 | 14.8 |
| Montana | 13.5 | 8.7 | 7.6 | 23.0 | 9.8 | 9.2 |
| Nebraska | 18.3 | 15.4 | 10.6 | 40.5 | 13.2 | 29.4 |
| Nevada | 10.6 | 8.7 | 7.9 | 13.0 | 11.0 | 8.6 |
| New Hampshire | 16.1 | 13.6 | 8.6 | 57.8 | 6.5 | 9.6 |

Exhibit 20: Case Error Indicators by State, 1998-2001 Combined: Total Households (Continued)

| State | Total error rate | First-month error rate | Next-month error rate for: |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  |  | Ongoing correct cases | Ongoing error cases | Expiring correct cases | Expiring error cases |
| New Jersey | 15.9 | 8.3 | 5.5 | 76.0 | 6.6 | 13.0 |
| New Mexico | 15.9 | 14.5 | 9.7 | 51.5 | 5.2 | 5.8 |
| New York | 17.0 | 14.0 | 5.9 | 71.5 | 8.2 | 17.2 |
| North Carolina | 11.6 | 8.8 | 6.8 | 40.5 | 11.5 | 6.8 |
| North Dakota | 9.8 | 10.9 | 6.5 | 15.9 | 14.4 | 8.8 |
| Ohio | 11.3 | 9.9 | 4.9 | 54.3 | 7.9 | 17.5 |
| Oklahoma | 14.9 | 14.9 | 8.9 | 30.6 | 15.7 | 10.7 |
| Oregon | 14.4 | 16.6 | 8.8 | 18.2 | 13.2 | 5.7 |
| Pennsylvania | 14.3 | 7.9 | 6.8 | 56.1 | 7.0 | 37.0 |
| Rhode Island | 10.8 | 6.9 | 5.6 | 46.8 | 3.7 | 28.4 |
| South Carolina | 9.9 | 7.0 | 4.9 | 39.8 | 10.8 | 21.7 |
| South Dakota | 4.7 | 3.0 | 2.6 | 42.8 | 1.0 | 42.6 |
| Tennessee | 10.3 | 9.9 | 5.7 | 52.3 | 4.5 | 11.1 |
| Texas | 8.4 | 5.1 | 4.8 | 55.1 | 2.2 | 11.1 |
| Utah | 15.4 | 10.7 | 10.1 | 40.5 | 13.3 | 13.3 |
| Vermont | 14.5 | 13.1 | 6.6 | 60.8 | 5.2 | 0.0 |
| Virginia | 14.0 | 11.0 | 9.8 | 40.7 | 11.3 | 0.0 |
| Washington | 12.4 | 12.7 | 7.8 | 34.0 | 8.9 | 16.4 |
| West Virginia | 12.5 | 9.6 | 6.6 | 58.2 | 3.7 | 5.1 |
| Wisconsin | 15.5 | 15.0 | 9.2 | 45.8 | 10.0 | 10.0 |
| Wyoming | 6.3 | 6.7 | 3.9 | 34.5 | 2.9 | 8.1 |
| U.S. Average | 14.4 | 10.2 | 7.7 | 48.1 | 7.8 | 11.9 |
| 25th percentile | 10.7 | 8.2 | 5.7 | 36.9 | 5.6 | 7.6 |
| 50th percentile | 13.5 | 10.7 | 6.8 | 45.8 | 8.6 | 13.3 |
| 75th percentile | 15.7 | 13.7 | 8.5 | 53.3 | 11.1 | 22.6 |

Note: Values less than or equal to the 25th percentile are shown in bold.

Exhibit 21: Case Error Indicators by State, 1998-2001 Combined: Households with Earnings

|  |  |  | Next-month error rate for: |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | :---: |

Exhibit 21: Case Error Indicators by State, 1998-2001 Combined: Households with Earnings (Continued)

|  |  |  | Next-month error rate for: |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |

Note: Values less than or equal to the 25th percentile are shown in bold.

Exhibit 22: Case Error Indicators by State, 1998-2001 Combined: Households without Earnings

|  |  |  |  |  |  | Next-month error rate for: |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |

Exhibit 22: Case Error Indicators by State, 1998-2001 Combined: Households without Earnings (Continued)

|  |  |  | Next-month error rate for: |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |

Note: Values less than or equal to the 25th percentile are shown in bold.

## Differences Among States in Their Underlying Error Patterns

Total household findings, shown in Exhibit 20, are:

- Some states achieve low total case error rates through strong performance in initial certification, interim action, and recertification. Arizona, Kentucky, and Wyoming, as examples, are in the lowest-quartile for their total error rate and are below the median in all five of the component error parameters.
- Other states have low overall rates and show strong performance on some but not all phases of certification, indicating the potential for further improvement. Minnesota and South Carolina, for instance, have total error rates in the lowest quartile despite having recertification procedures that appear not as effective as most other states in preventing and correcting errors. In contrast, Oklahoma and Oregon do reasonably well in containing errors at interim action and recertification, but each has a high error rate at initial certification.

A contrasting pattern of error is evident among households with earnings, as shown in Exhibit 21.

- The following states are in the lowest quartile for the total error rate and are below the median in all five component parameters, for households with earnings: Arizona, Minnesota, South Dakota, Tennessee, and Wyoming.
- A number of other states have low total error rates for households with earnings, but with potential for improvement on some phases of case action. For example, Idaho, Indiana, and Texas are in the lowest quartile for the total error rate, despite having a high next-month error rate for ongoing error cases. This suggests that their interim action procedures are not effective in detecting and correcting errors. Mississippi and South Carolina, in contrast, do very well at avoiding errors at initial certification and interim action, but could improve the accuracy of their recertification decisions.

In possible future research, it would be instructive to consider the error patterns by stateespecially for households with earnings-in the context of the client reporting provisions that states have adopted for such households. These reporting provisions are presumed to influence especially the "client error" component of the error rate-i.e., the extent to which error is attributable to a household not having correctly or completely reported information that it was required to report, such as a change in income or other financial circumstances. Throughout this report, the error measures include both client error and agency error.

## Chapter Six: Effects of More Frequent Recertification

Short certification periods are frequently used to control error, especially among cases with earnings. Such periods may unintentionally reduce program participation.

## Using the Model to Test Alternative Scenarios

One can use the derived transition matrix for each year to examine the effects of alternative scenarios regarding more frequent recertification. In doing so, we make use of the formulas provided at the end of Chapter Three for the projected long-term aggregate participation rate and case error rate. To recall, the steady-state distribution expresses the eventual outcome that would be obtained if the pattern of monthly transitions, as estimated for a given "base year," were continued uninterrupted. As explained in this chapter, changes in the frequency of recertification can be evaluated in terms of the shift in the long-term distribution of households according to their participation and error status.

The alternative scenario examined here is one that increases by 5 percentage points the "recertification rate" for cases with earnings. We define the recertification rate as the percentage of current-month active cases that are expiring cases (i.e., cases in the final month of their certification period). For these cases, their next-month status will be a result of the recertification process. The remaining balance of current-month active cases consists of ongoing cases, whose next-month status will be subject to the interim action process.

In the terms of the basic transition matrix (Exhibit 6), the recertification rate is measured as the share of the current-month caseload comprised by expiring correct cases (the fourth-row total) and expiring error cases (the fifth-row total), as follows:

$$
\text { Recertification rate }=\left(\mathrm{R}_{4}+\mathrm{R}_{5}\right) /\left(\mathrm{R}_{2}+\mathrm{R}_{3}+\mathrm{R}_{4}+\mathrm{R}_{5}\right)
$$

One can carry this logic further to specify the recertification rate separately for current-month correct cases as $\mathrm{R}_{4} /\left(\mathrm{R}_{2}+\mathrm{R}_{4}\right)$ and for current-month error cases as $\mathrm{R}_{5} /\left(\mathrm{R}_{3}+\mathrm{R}_{5}\right)$.

Note that the recertification rate will approximately equal the inverse of the average length of certification period. For cases with earnings, whose average certification length is about 7 months (as shown in Exhibit A-2), the recertification rate will thus be about $1 / 7$ or 14.3 percent. An increase of 5 percentage points in the recertification rate, to 19.3 percent, corresponds to a reduction in the average certification length from 7 months to about 5.2 months (1/0.193).

In understanding how the model is used to test such an alternative scenario, it is useful to regard the five-group model as reducible to a three-group equivalent (as shown in Exhibit 4 of Chapter Two), whose groups consist of nonparticipating households, correct cases, and error cases. The transition probabilities for correct cases (the second-row entries of the threegroup transition matrix) can be viewed as weighted averages of the probabilities specific to ongoing correct cases (i.e., those subject to interim action) and expiring correct cases (i.e., those subject to recertification). The weights reflect the recertification rates for correct cases. A similar logic applies to the transition probabilities for error cases (the third-row entries of the three-group matrix).

In using the model to simulate an assumed increase in the recertification rate, one allows the next-month status of current-month cases to reflect more heavily the transition probabilities specific to recertifications. Thus, the probabilities specific to interim actions are reflected less heavily. The model does this separately for current-month correct cases and current-month error cases. Specifically, we assume that the increase of 5 percentage points in the overall recertification rate occurs through a 5-percentage-point increase in the recertification rate for both current-month correct cases and current-month error cases.

To explain further, the "alternative scenario" is simulated by modifying the transition matrix for the "base scenario" in the following manner. We first convert the 5-by-5 matrix of cell counts (previously estimated for the chosen base year) into an equivalent 3-by-3 matrix of cell counts. This reconfigured matrix maintains nonparticipating households as one group and establishes participating correct cases and participating error cases as the other two groups. (Thus, the distinction between correct cases and error cases is maintained in the new matrix, but the explicit distinction between ongoing cases and expiring cases is removed.) The higher recertification rate under the alternative scenario can then be translated into a recomputed set of transition probabilities along the newly defined second row (current-month correct cases) and the newly defined third row (current-month error cases). Within each of these rows, the modified pattern of next-month outcomes under the alternative scenario can be represented through a re-weighting of the component transition probabilities specific to ongoing cases and expiring cases. In each row, the assumed 5-percentage-point increase in the recertification rate implies a reduction in the weights specific to ongoing cases and a corresponding increase in the weights specific to expiring cases.

The alternative scenario is assumed to have no effect on the rate of case openings and the accuracy of initial certifications (i.e., the first-row matrix entries). The alternative scenario is also assumed to cause no behavioral response among cases during interim months. For example, the simulation assumes no change in the likelihood that an ongoing error case voluntarily withdraws from the program.

Having recalculated the transition matrix through such a re-weighting, one can than apply the formulas shown at the end of Chapter Three to re-compute the long-term values for the
aggregate participation rate and the case error rate. (The formulas, developed in the context of a 5-by-5 matrix, are also applicable to the reconfigured 3-by-3 framework described above.) The effect of the alternative scenario is assessed by comparing these long-term values with those corresponding to the base scenario.

Note that such a shift could either increase or decrease the case error rate, depending on whether the reduction in error cases (the numerator of the case error rate) falls short of or exceeds the reduction in active cases (the denominator). Although in principle the effect on the aggregate participation rate (and thus on the size of the caseload) could also be either upward or downward, one expects a downward change. The reason is that a case termination is more likely when a case is subject to a recertification than when it is subject to the interim action process. If recertifications occur more frequently, the higher resulting rate of monthly case closure (along with an unchanged case opening rate) will expectedly reduce the aggregate participation rate.

To see whether the results are sensitive to the particular base year, we tested the alternative scenario using both 2000 and 2001 as base years.

## Findings

Exhibit 23 shows the estimated long-term effect on participation and error for cases with earnings of the assumed 5 percentage point increase in the recertification rate for such cases. Using 2000 as the base year, the alternative scenario implies an increase in the recertification rate from 15.50 to 20.50 percent, roughly equivalent to a reduction in the average certification length from 6.5 to 4.9 months. With 2001 as the base year, the implied increase in the recertification rate is from 13.53 to 18.53 percent.

Under both sets of estimates, the results are consistent with respect to the direction of change in the components of the aggregate participation rate and the aggregate error rate, both of which are expressed in terms of the total population of households with earnings.

- Aggregate participation rate: Under both sets of estimates, the alternative scenario was found to increase the case closure rate (b). In conjunction with an unchanged case opening rate (a), this leads to a lower aggregate participation rate $(\mathrm{a} /(\mathrm{a}+\mathrm{b}))$. In proportional terms, the estimated effect on the eventual size of the caseload for earnings cases is a 6.5 percent reduction for the 2000-based estimates, versus only a 0.7 percent reduction for the 2001-based estimates. The difference appears attributable to the lower rate of case closure for expiring correct cases in the 2001 matrix ( 4.9 percent) that in the 2000 matrix (15.4 percent), as shown in Exhibits C-7 and C-8.

Exhibit 23: Estimated Effects of More Frequent Recertification, Households with Earnings

|  |  | Long-term outcome |  | Absolute change | Proportional change |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Base scenario | Alternative scenario |  |  |
|  |  | Households with earnings: base year 2000 |  |  |  |
| Recertification rate (\%) |  | 15.50 | 20.50 | 5.00 |  |
| Case opening rate (\%) | a | 0.24 | 0.24 |  |  |
| Case closure rate (\%) | b | 10.29 | 11.02 |  |  |
| Aggregate participation rate (\%) | $p^{*}=a /(a+b)$ | 2.31 | 2.16 | -0.15 | -6.5\% |
| Error opening rate (\%) | c | 0.33 | 0.33 |  |  |
| Error closure rate (\%) | d | 64.97 | 66.54 |  |  |
| Aggregate error rate (\%) | $e^{*}=c /(c+d)$ | 0.51 | 0.49 | -0.02 | -3.3\% |
| Case error rate (\%) | $r^{*}=e^{\star} / p^{*}$ | 21.92 | 22.68 | 0.76 | 3.5\% |
|  |  | Households with earnings: base year 2001 |  |  |  |
| Recertification rate (\%) |  | 13.53 | 18.53 | 5.00 |  |
| Case opening rate (\%) | a | 0.26 | 0.26 |  |  |
| Case closure rate (\%) | b | 11.02 | 11.09 |  |  |
| Aggregate participation rate (\%) | $p^{*}=a /(a+b)$ | 2.33 | 2.32 | -0.02 | -0.7\% |
| Error opening rate (\%) | C | 0.30 | 0.30 |  |  |
| Error closure rate (\%) | d | 65.81 | 67.36 |  |  |
| Aggregate error rate (\%) | $e^{*}=c /(c+d)$ | 0.45 | 0.44 | -0.01 | -1.7\% |
| Case error rate (\%) | $r^{*}=e^{*} / p^{*}$ | 19.19 | 18.99 | -0.20 | -1.0\% |

Note: All outcomes pertain to households with earnings. See Chapter Three for the formulas used to compute each long-term outcome measure.

- Aggregate error rate: Under both sets of estimates, the alternative scenario was found to increase the error closure rate (d), with no change in the error opening rate (c). This implies a lower aggregate error rate (c/(c+d)). Stated otherwise, the alternative scenario leads to a very slight long-term reduction in the average monthly number of error cases among households with earnings. The projected
reduction is more pronounced in the 2000-based estimates ( 3.3 percent) than in the 2001-based estimates (1.7 percent).
- Case error rate: The effect on the case error rate will depend on whether the effect on the aggregate error rate is proportionally greater than or less than the effect on the aggregate participation rate. Here, the two sets of estimates differ. The 2000-based estimates imply an increase in the case error rate, as the estimated reduction in participation ( 6.5 percent) is greater proportionally that the reduction in error ( 3.3 percent). In contrast, the 2001-based estimates show a decrease in the case error rate, as the estimated drop in participation ( 0.7 percent) is smaller than the reduction in error (1.7 percent).

These findings suggest that the long-term effects of more frequent recertification for earnings cases will be to reduce both the monthly number of participating households and the monthly number of error cases among households with earnings. The effect will be to reduce the case error rate (the percentage of participating cases in error) only if the proportional reduction in error cases exceeds the reduction in participating households. This relationship was found in the 2000-based estimates but not in the 2001-based estimates.

The findings should be interpreted with some caution, however, as they are based on assumptions about the recertification process. These assumptions, as noted earlier, are necessitated by the limitations of the QC data in providing retrospective detail on the status of correct cases. Specifically (with respect to cell $\mathrm{Q}_{52}$ in Exhibit 6), the model assumes that among cases currently in error, the percentage that emerge the following month as correctly paid active cases is the same for those subject to recertification as for those subject to interim action. Under this assumption, the effect of more frequent recertification on the error rate occurs through the higher closure of error cases at recertification. Such case closures reduce both the numerator and denominator of the case error rate, thus moderating the change in the measured case error rate.

Another key modeling assumption (pertaining to cells $\mathrm{Q}_{32}$ and $\mathrm{Q}_{34}$ in Exhibit 6) is that ongoing or expiring cases that are correct in one month were in error the prior month only if an interim change has just occurred. The interim change is thus assumed to have corrected an error.

Estimates based on the matrix for 2000 show a participation effect for cases with earnings similar to that found in the earlier-mentioned studies by Kabbani-Wilde and Kornfeld. These previous studies estimated the effect of an increase of 10 percentage points in the proportion of earnings cases subject to a short certification period. The alternative scenarios tested here represent upward shifts twice as large, increases of 18 to 22 percentage points in the
proportion of earnings cases with a short certification period. ${ }^{24}$ Scaling back the 2000-based finding here by a factor of 2 , the estimated caseload reduction amounts to 3.2 percent ( 6.5 percent divided by 2). This compares to the 2.6 percent caseload decline estimated by Kabbani-Wilde (for cases with earnings) and the 2.3 to 2.4 percent caseload decline estimated by Kornfeld (for cases consisting of multiple adults with children or cases with adults only). The estimated proportional decrease in participation is much smaller here (0.3 percent) if one instead uses 2001 as the base year.

In contrast, the model-estimated effect on the error rate for cases with earnings is negligible compared to the effect found in the Kabbani-Wilde study. The findings here, again scaled back to reflect an increase of 10 percentage points in the proportion of earnings cases subject to a short certification period, range from an error rate increase of 0.4 percentage points for the 2000-based estimate to an error rate decrease of 0.1 percentage point for the 2001-based estimate. The corresponding effect on the dollar error rate would be even smaller, as dollar error rates are consistently smaller than case error rates. ${ }^{25}$ The Kabbani-Wilde study found a reduction of 0.8 percentage points in the dollar error rate for earnings cases.

As to why the model developed here yields a smaller effect on error than found by KabbaniWilde, one can only speculate. One possible explanation is that, although the Kabbani-Wilde estimating equation included many covariates (plus state-specific and year-specific fixed effects), the estimated coefficient on the short-certification variable may have captured the influence of other excluded variables that are correlated with the use of more frequent recertification. One indication of this is that, in the Kabbani-Wilde error rate equation for households without earnings, the short-certification variable was again strongly significant, although somewhat smaller than for earnings cases.

In comparison to the more conventional multivariate econometric approaches used by Kabbani-Wilde and Kornfeld, one seeming advantage of the model developed here is that it constitutes a more direct operational representation of the administrative processes and
${ }^{24}$ Using 2000 as the base year, the increase of 5 percentage points in the recertification rate for earnings cases (from 15.5 to 20.5 percent) is roughly equivalent to a 1.6 -month reduction in the average certification length, from 6.5 months $(1 / 0.155)$ to 4.9 months ( $1 / 0.205$ ). It is reasonable to assume that any increase in the proportion of earnings cases subject to a short certification period is achieved by shortening to 3 months the certification period of those with annual recertification. For every additional 10 percent of earnings cases assigned to a 3-month certification, the reduction in the average certification length is thus 0.9 months. To achieve a reduction of 1.6 months in the average certification length, the implied increase in the proportion of earnings cases subject to a short certification period would be 18 percentage points (1.6/0.9). Using 2001 as the base year, the corresponding estimate is 22 percentage points. These estimates, ranging from 18 to 22 percent, are approximately twice as large as the 10 percent assumed in the Kabbani-Wilde and Kornfeld studies.

25 The dollar error rate is the product of the case error rate and the following ratio: the average error amount divided by the average payment amount. Because the latter ratio is typically less than one, the dollar error rate is typically less than the case error rate.
events that underlie food stamp error and participation. The simplifying assumptions of the model, as necessary to overcome the limitations of QC data, can be explicitly considered and tested. Indeed, the issues raised in this exploratory research suggest that the modeling assumptions should be further scrutinized before undertaking any future applications of the model.

## Bibliography

Affholter, Dennis P. and Fredrica D. Kramer (eds.), Rethinking Quality Control: A New System for the Food Stamp Program, National Academy Press, Washington, DC, 1987.

Agresti, Alan. Categorical Data Analysis. John Wiley \& Sons, 1990. (See especially Section 11.5: "Markov Chain Models.")

Bartholomew, D.J. Stochastic Models for Social Processes. John Wiley \& Sons, 1982. (See especially Section 2.2: "The Markov Model for Social Mobility.")

Bartlett, Susan, et al. Food Stamp Program Access Study: Draft Final Report, Abt Associates Inc., Cambridge, Mass., March 2004.

Cameron, A. Colin and Pravin K. Trivedi. Regression Analysis of Count Data. Cambridge University Press. 1998.

Cox, D.R. and H.D. Miller. The Theory of Stochastic Processes. Chapman \& Hall, 1965. (See especially Chapter 3: "Markov Chains.")

Cunnyngham, Karen. Trends in Food Stamp Program Participation Rates: 1994 to 2000, Mathematica Policy Research, June 2002.

Cunnyngham, Karen. Trends in Food Stamp Program Participation Rates: 1999 to 2001, Mathematica Policy Research, July 2003.

Dean, Stacy. Making Food Stamps Work: Important New State Option for Semi-Annual Reporting Could Dramatically Improve the Food Stamp Program for Working Poor Families, Center on Budget and Policy Priorities, December 7, 2000.

Dean, Stacy, and Dorothy Rosenbaum. Implementing New Changes to the Food Stamp Program: A Provision-by-Provision Analysis of the Farm Bill, Center on Budget and Policy Priorities. January 21, 2003.

Diggle, Peter J., Kung-Yee Liang, and Scott L. Zeger. Analysis of Longitudinal Data. Oxford Statistical Science Series 13. Clarendon Press, 1994. (See especially Section 7.3: "Transition (Markov) Models.")

Drake, Alvin, Fundamentals of Applied Probability Theory. McGraw-Hill. (See especially Chapter Five: "Discrete-State Markov Processes.")

Fishman, George S. Monte Carlo: Concepts, Algorithms, and Applications. Springer Series in Operations Research. Springer-Verlag New York, 1996. (See especially Section 5.1: "Markov Processes.")

Guttorp, Peter. Stochastic Modeling of Scientific Data. Chapman \& Hall, 1995. (See especially Chapter 2: "Discrete time Markov chains.")

Howard, Ronald A. Dynamic Probabilistic Systems, Volume II: Semi-Markov and Decision Processes. John Wiley \& Sons, 1971. (See especially Chapter 10: "The Discrete-Time Semi-Markov Process.")

Kabbani, Nader S. and Parke E. Wilde, "Short Certification Periods in the U.S. Food Stamp Program," Journal of Human Resources 38, pp. 1112-1138, 2003.

Kemeny, John G. and J. Laurie Snell. Finite Markov Chains. Van Nostrand, 1960
Kornfeld, Robert, Explaining Recent Trends in Food Stamp Program Caseloads: Final Report, Abt Associates Inc., March 2002.

Lindsey, J. K. Modeling Frequency and Count Data. Oxford Statistical Science Series 15. Clarendon Press, 1995. (See especially Chapter 8: "Markov chains.")

Mills, Gregory, et al., Redesign of the Negative Action Quality Control System in the Food Stamp Program, Abt Associates Inc., Cambridge, Mass., June 1990.

Mills, Gregory. Reducing Food Stamp Overpayments: More Frequent Recertification and Monthly Reporting, Abt Associates Inc., Cambridge, Mass., December 1988.

MacDonald, Iain L. and Walter Zucchini. Hidden Markov and Other Models for Discretevalued Time Series. Monographs on Statistics and Applied Probability 70. Chapman \& Hall/CRC Press, 1997.

Parzen, Emanuel. Stochastic Processes. Holden-Day, 1962. (See especially Chapter 6: "Markov Chains: Discrete Parameter.")

Rosenbaum, Dorothy. Improving Access to Food Stamps: New Reporting Options Can Reduce Administrative Burdens and Error Rates. Center on Budget and Policy Priorities. September 1, 2000.

Rosenbaum, Dorothy. States Have Significant Flexibility in the Food Stamp Program. Center on Budget and Policy Priorities, June 17, 2002.

Rosenbaum, Dorothy and David Super. Understanding Food Stamp Quality Control. Center on Budget and Policy Priorities, April 27, 2001.

Ross, Sheldon M. Introduction to Probability Models. Academic Press, 1985. (See especially Chapter 4: "Markov Chains.")

Shanmugan, K. Sam and Arthur M. Breipohl. Random Signals: Detection, Estimation, and Data Analysis. John Wiley \& Sons. (See especially Section 5.3: "Markov Sequences and Processes.")

Shirm, Allen L. and Laura A. Castner. Reaching Those in Need: State Food Stamp Participation Rates in 2000, Mathematica Policy Research, December 2002.

Super, David. Food Stamp Overpayment Error Rate Hits Record Low. Center on Budget and Policy Priorities, May 23, 2001.
U.S. Department of Agriculture, Food and Nutrition Service. "Change Reporting Systems and Waivers Summary, As of August 31, 2001."
U.S. Department of Agriculture, Food and Nutrition Service. "Characteristics of Food Stamp Households," fiscal years 1997-2001.
U.S. Department of Agriculture, Food and Nutrition Service. "Food Stamp Quality Control Annual Report," fiscal years 1997-2001.
U.S. Department of Agriculture, Food and Nutrition Service, "FNS-310: Quality Control Review Handbook." fiscal years 1997-2001.
U.S. Department of Agriculture, Food and Nutrition Service. "Error Rates, Potential and Adjusted Liabilities, and Enhancing Funding," unpublished tabulations for fiscal years 1998-2001, the most recent dated May 24, 2002.

Zelterman, Daniel. Models for Discrete Data. Clarendon Press, 1999.

## Appendix A Distribution of Food Stamp Households by Length of Certification Period



Exhibit A-1: Distribution of food stamp households by length of certification period, 1997-2001: all households

| Length of certification period | 1997 | 1998 | 1999 | 2000 | 2001 |
| :---: | :---: | :---: | :---: | :---: | :---: |
|  | Percentage distribution, all households (\%) |  |  |  |  |
| Number of months |  |  |  |  |  |
| 1 | 0.5 | 0.5 | 0.5 | 0.5 | 0.4 |
| 2 | 0.8 | 0.9 | 1.1 | 1.1 | 0.9 |
| 3 | 10.8 | 13.6 | 15.3 | 16.9 | 15.9 |
| 4 | 3.2 | 3.4 | 3.8 | 4.5 | 4.3 |
| 5 | 1.9 | 2.0 | 2.1 | 1.9 | 1.8 |
| 6 | 17.0 | 14.5 | 13.8 | 12.7 | 14.2 |
| 7 | 2.3 | 2.3 | 2.3 | 2.1 | 2.2 |
| 8 | 0.8 | 0.8 | 0.7 | 0.7 | 0.7 |
| 9 | 0.7 | 0.6 | 0.5 | 0.6 | 0.5 |
| 10 | 0.9 | 0.9 | 0.8 | 0.6 | 0.5 |
| 11 | 2.1 | 1.9 | 1.9 | 1.7 | 1.6 |
| 12 | 49.2 | 49.4 | 48.7 | 47.8 | 48.2 |
| 13+ | 9.3 | 9.0 | 8.2 | 8.4 | 8.3 |
| Unknown | 0.5 | 0.2 | 0.4 | 0.4 | 0.3 |
| Total | 100.0 | 100.0 | 100.0 | 100.0 | 100.0 |
|  | Estimated value, all households (months) |  |  |  |  |
| Mean | 9.9 | 9.9 | 9.7 | 9.6 | 9.7 |
| Standard error | 0.03 | 0.05 | 0.03 | 0.03 | 0.03 |

Source: U.S. Department of Agriculture, Food and Nutrition Service, "Characteristics of Food Stamp Households," fiscal years 1997-2001.

Exhibit A-2: Distribution of food stamp households by length of certification period, 1997-2001: households with earnings

| Length of <br> certification <br> period |  |  |  |  |
| ---: | ---: | ---: | ---: | ---: |
|  |  |  |  |  |
| Number of months |  |  |  |  |
| $\mathbf{1}$ |  |  |  |  |
| $\mathbf{2}$ |  |  |  |  |

Source: U.S. Department of Agriculture, Food and Nutrition Service, "Characteristics of Food Stamp Households," fiscal years 1997-2001.


Source: U.S. Department of Agriculture, Food and Nutrition Service, "Characteristics of Food Stamp Households," fiscal years 1997-2001.

## Appendix B Procedures For Deriving The Transition Matrix

This appendix explains the derivation of the cell counts for each column of the transition matrix. The notation here is as shown in Exhibit 6 of the report.

## Variables Used from Food Stamp Quality Control Data

The model makes use of the following variables for each sample case in the food stamp QC database:

- STATUS: case error finding for review month

1 amount correct
2 overissuance (\$25 or more)
3 underissuance (\$25 or more)
4 ineligible

- ACTNTYPE: action type, for most recent action

1 initial certification
3 recertification
5 interim change

- CERTMTH: length of assigned certification period, in months
- LASTCERT: number of months since last formal certification for food stamps (initial certification or recertification).
- RCNTACTN: date (year, month, and day) of the most recent action (either initial certification, recertification, or interim change)
- OCCDATE: date (year and month) of error occurrence (for error cases) ${ }^{1}$
- YRMONTH: date (year and month) for which the case has been sampled and reviewed (review month)
- FSEARN: monthly household earnings, as indicated in the case record (sum of wages and salaries, self-employment income, and other earned income).

It is important to note several limitations of the QC data, which require the creation of additional variables or the use of assumptions in estimating the model:

[^19]- For cases in which the most recent action is an interim change, the variable LASTCERT indicates the number of months since the most recent formal certification (initial certification or recertification). Because of the need to establish the time elapsed since the most recent action (initial certification, interim change, or recertification) for all cases, an additional variable, LASTACT, is computed. This is the number of months since the most recent action, computed as the difference between RCNTACTN and YRMONTH. LASTACT thus differs from LASTCERT for those cases whose most recent action was an interim change.
- For active cases found in error at the QC review, the indicated timing of error occurrence is used to establish whether the case was correct or in error during the preceding month. An error case is considered to have been correct in the preceding month if the occurrence dates for all indicated errors are within two months of the review month. Conversely, an error case is considered to have been in error in the preceding month if the occurrence date for any of its indicated errors is more than two months before the review month. ${ }^{2}$
- For cases that are correct at the QC review, there is no indication in the QC record of the presence of error in any previous month. As indicated below, we assume that an ongoing correct case was also correct in the preceding month, unless the case has just undergone an interim change. The interim change is assumed to have occurred because the case was in error and required a benefit adjustment. ${ }^{3}$

The following sections of this appendix explain the derivation of the cell counts for each column of the matrix, beginning with the right-most of the five columns and then working to the left, using the notation in Exhibit 6 of this report. Using the definitions described below, and as shown in Exhibits B-1 through B-5, all cases in the QC sample are assigned to one (and only one) cell in columns two through five. Then, as described in the final section below, the cells in the first column entries are computed as row residuals.

[^20]In this appendix, we refer to month-to-month transitions that occur between the QC review month and the prior month. In the body of the report, for ease of exposition, we adopt a different terminology, referring to transitions between the current month and the next month. We find it helpful to consider these transitions as occurring from March to April. Under this nomenclature, the review month is the next month, and the month preceding the review is the current month.

## Expiring Error Cases

The fifth column of the transition matrix (shown in Exhibit 6) is comprised of error cases that are at the end of their certification period in the review month. By definition, such cases are already in error as they enter recertification. The column total $\left(\mathrm{C}_{5}\right)$ is the count of all cases that are in error (STATUS=2, 3, or 4) and whose certification is about to expire or has expired (LASTCERT \$CERTMTH-1). ${ }^{4}$ This includes, for example, error cases assigned a 12-month certification period and for whom 11 or more months have elapsed since their most recent certification. ${ }^{5}$ Throughout this analysis, we use the word "expiring" to indicate that the case's certification period is about to end, not that the case is itself about to terminate. As noted above, these cases are ones about to undergo recertification; their error has preceded the recertification action.

We determine the prior-month status of each case in the fifth column, and thus determine its placement in one of the five cells in the fifth column, as follows (and as shown in Exhibit B1):

[^21]Exhibit B-1: Derivation of Fifth-Column Cell Counts: Expiring Error Cases

| Cell Count | Prior-month <br> ("March") status | Review-month <br> ("April") status | Criteria (based on variables from QC review, as defined in text) |
| :---: | :--- | :--- | :--- |
| $Q_{15}$ | Nonparticipating | Expiring error | equals 0, by assumption; see text |
| $Q_{25}$ | Ongoing correct | Expiring error | certification expiring (LASTCERT\$CERTMTH-1) and <br> error in review month (STATUS=2, 3, or 4) and <br> correct in prior month (OCCDATE\$YRMONTH-2) |
| $Q_{45}$ | Ongoing error | Expiring error | certification expiring (LASTCERT\$CERTMTH-1) and <br> error in review month (STATUS=2, 3, or 4) and <br> correct in prior month (OCCDATE<YRMONTH-2) |
| $Q_{55}$ | Expiring correct | Expiring error | equals 0, by assumption; see text |
|  | Expiring error | Expiring error | equals 0, by assumption; see text |

- nonparticipating in prior month-The $\mathrm{Q}_{15}$ cell count is 0 by definition, as the model assumes that one-month certification periods are negligible, implying that a nonparticipant would never in the next month be an active case at the end of its certification period.
- ongoing correct in prior month-The $\mathrm{Q}_{25}$ cell consists of expiring cases for whom an error has just occurred, as indicated by an interval of two months or less between the error occurrence month and the review month (YRMONTHOCCDATE \#2) for the earliest indicated error in the case. ${ }^{6}$
- ongoing error in prior month-The $\mathrm{Q}_{35}$ cell consists of expiring error cases that were previously in error, as indicated by an interval of more than two months between the error occurrence month and the review month (YRMONTHOCCDATE $>2$ ) for the earliest indicated error in the case.
- expiring correct in prior month-The $\mathrm{Q}_{45}$ cell count is 0 by definition, as we assume that cases at the end of their certification period in the review month could not also have been expiring in the prior month.
- expiring error in prior month-For the same reason cited above, the $\mathrm{Q}_{55}$ cell count is also 0 by definition.


## Expiring Correct Cases

The fourth column of the matrix consists of correct cases that in the review month are at the end of their certification period. The column total $\left(\mathrm{C}_{4}\right)$ is the count of all expiring correct cases: cases that are both correctly paid (STATUS=1) and whose certification is about to expire or has expired (LASTCERT \$ CERTMTH-1).

We determine the prior-month status of each case, and thus its contribution to the cell counts in the fourth column, as follows (and as shown in Exhibit B-2):

- nonparticipating in prior month-The $\mathrm{Q}_{14}$ cell count is 0 by definition (as with $\mathrm{Q}_{15}$ ), for we assume no one-month certification periods. A case thus cannot be nonparticipating in one month and an expiring active case in the next month.

[^22]Exhibit B-2: Derivation of Fourth-Column Cell Counts: Expiring Correct Cases

| Cell Count | Prior-month <br> ("March") status | Review-month <br> ("April") status | Criteria (based on variables from QC review, as defined in text) |
| :---: | :--- | :--- | :--- |
| $Q_{14}$ | Nonparticipating | Expiring correct | equals 0, by assumption; see text |
| $Q_{24}$ | Ongoing correct | Expiring correct | certification expiring (LASTCERT\$CERTMTH-1) and <br> correct in review month (STATUS=1) |
| $Q_{44}$ | Ongoing error | Expiring correct | certification expiring (LASTCERT\$CERTMTH-1) and <br> correct in review month (STATUS=1) and <br> interim action in review month (ACTNTYPE=5 and LASTACT=0) |
| $Q_{54}$ | Expiring correct | Expiring correct | equals 0, by assumption; see text |

- ongoing correct in prior month-The $\mathrm{Q}_{24}$ cell consists of expiring correct cases which were also correct in the prior month, as indicated by the absence of any recent interim change $($ ACTNTYPE $=1$ or 3 , or ACTNTYPE $=5$ and LASTACT $>0$ ). Under the model's assumptions, the only circumstance in which an expiring correct case is regarded as having been in error in the prior month is if the case has just undergone an interim change, as accounted for below in cell $\mathrm{Q}_{34}$.
- ongoing error in prior month—The $\mathrm{Q}_{34}$ cell consists of expiring correct cases for whom an error has just been corrected, as indicated by a recent interim change (ACTNTYPE $=5$ and LASTACT $=0$ ).
- expiring correct in prior month-The $\mathrm{Q}_{44}$ cell count is 0 by definition (as with $\mathrm{Q}_{45}$ earlier), assuming that cases reaching the end of their certification period in the review month could not also have been an expiring case in the preceding month.
- expiring error in prior month-For the same reason cited above, the $\mathrm{Q}_{54}$ cell count is also 0 by definition.

The assumption regarding cell $\mathrm{Q}_{34}$ deserves further explanation. In deriving this cell count, we assume that the only scenario in which an ongoing or expiring correct case would have been in error the previous month is the situation in which the case has just undergone an interim action. We assume no self-correction of errors-i.e., no situations in which a case in error one month becomes correct the following month, in the absence of agency action. There are some "transient error" situations, however, in which self-correction might occur in the absence of an interim change. Consider, for instance, a case that is overpaid because of unreported income from a particular source. If the income from that source later returns to zero, the case becomes correct. Conversely, however, there will be some instances in which an interim change has occurred, but where the case has remained correctly paid throughout. For instance, if the interim change were to adjust the benefit by less than $\$ 25$ for an eligible case, there would have been no previous error, because of the $\$ 25$ error tolerance. Also, if the interim change was in response to a client's timely report of a change in household circumstances, an error would have been prevented, but the case would not have actually been in error in the prior month.

In the text of this report, we have tested alternative specifications and examined the sensitivity of the model's findings to this assumption.

## Ongoing Error Cases

The third column consists of error cases that in the review month are not at the end of their certification period. The column total $\left(\mathrm{C}_{3}\right)$ is the count of all cases that are in error in the
review month (STATUS $=2$, 3, or 4) and whose certification has not expired (LASTCERT $<$ CERTMTH-1). This includes, for example, error cases assigned a 12-month certification period and for whom less than 11 months have passed since their last certification.

We determine the prior-month status of each case, and thus its placement in the cells of the third column, as follows (and as shown in Exhibit B-3):

- nonparticipating in prior month-The $\mathrm{Q}_{13}$ cell consists of households that have just been initially certified, as indicated by an initial certification that coincides with the review month (ACTNTYPE $=1$ and LASTCERT=0).
- ongoing correct in prior month-The $\mathrm{Q}_{23}$ cell consists of ongoing cases whose error has just occurred, as indicated by an interval of two months or less between the error occurrence month and the review month (YRMONTH-OCCDATE \#2) for the earliest indicated error in the case.
- ongoing error in prior month-The $\mathrm{Q}_{33}$ cell consists of ongoing cases for whom the error has not just occurred, as indicated by an interval of more than two months between the error occurrence month and the review month (YRMONTHOCCDATE $>2$ ) for the earliest indicated error in the case.
- expiring correct in prior month-The $\mathrm{Q}_{43}$ cell consists of cases that have just been recertified (ACTNTYPE $=3$ and LASTCERT $=0$ ) and whose error has just occurred, as indicated by an interval of two months or less between the error occurrence month and the review month (YRMONTH-OCCDATE \#2) for the earliest indicated error in the case.
- expiring error in prior month-The $\mathrm{Q}_{53}$ cell consists of cases that have just been recertified $($ ACTNTYPE $=3$ and LASTCERT $=0)$ and whose error has not just occurred, as indicated by an interval of more than two months between the error occurrence month and the review month (YRMONTH-OCCDATE $>2$ ) for the earliest indicated error in the case.

Exhibit B-3: Derivation of Third-Column Cell Counts: Ongoing Error Cases

| Cell Count | Prior-month <br> ("March") status | Review-month <br> ("April") status |
| :---: | :--- | :--- |
| $Q_{13}$ | Nonparticipating | Ongoing error |
| $Q_{23}$ | Ongoing correct | Ongoing error |
| $Q_{33}$ | Ongoing error (based on variables from QC review, as defined in text) |  |

## Ongoing Correct Cases

The second column of the matrix consists of correct cases that are not at the end of their certification period in the review month. The column total $\left(\mathrm{C}_{2}\right)$ is the count of all cases that are correctly paid (STATUS=1) and whose certification has not expired (LASTCERT < CERTMTH-1).

We determine the prior-month status of these cases, and thus their contribution to the cell counts in the second column, as follows (and as shown in Exhibit B-4):

- nonparticipating in prior month-The $\mathrm{Q}_{12}$ cell consists of households that have just been initially certified, as indicated by an initial certification that coincides with the review month (ACTNTYPE $=1$ and LASTCERT=0).
- ongoing correct in prior month-The $\mathrm{Q}_{22}$ cell consists of ongoing cases for which no error existed in the preceding month, as indicated by the absence of any recent interim change (i.e., ACTNTYPE $=1$ or 3 , or ACTNTYPE $=5$ and LASTACT $>$ 0 ). The rationale for this is the same as for the $\mathrm{Q}_{24}$ cell.
- ongoing error in prior month-The $\mathrm{Q}_{32}$ cell consists of ongoing cases for which an error has just been corrected, as indicated by a recent interim change ( ACTNTYPE $=5$ and LASTACT $=0$ ). Here the logic is the same as for the $\mathrm{Q}_{34}$ cell.
- expiring correct in prior month-The $\mathrm{Q}_{42}$ cell count is computed as the row residual $\left(\mathrm{R}_{4}-\mathrm{Q}_{41}-\mathrm{Q}_{43}\right){ }^{7}$
- expiring error in prior month-The $\mathrm{Q}_{52}$ cell count is computed under the assumption that, for a case entering recertification in error, the probability of becoming a correctly paid active case in the following month is the same as for an error case not undergoing recertification, which equals $\left(Q_{32}+Q_{34}\right) / R_{3}$.

The assumption used in computing the $\mathrm{Q}_{52}$ cell count warrants further explanation, as it figures importantly in simulations of the effects of more frequent recertification. Because the QC data do not provide information on the prior-month error status of cases that are correct in the review month, some assumption is necessary for either $\mathrm{Q}_{42}$ or $\mathrm{Q}_{52}$. Once one of these values is estimated, the other (along with $\mathrm{Q}_{41}$ and $\mathrm{Q}_{51}$ ) can be computed as either a row or column residual.

[^23]Exhibit B-4: Derivation of Second-Column Cell Counts: Ongoing Correct Cases

| Cell Count | Prior-month <br> ("March") status | Review-month <br> ("April") status | Criteria (based on variables from QC review, as defined in text) |
| :---: | :--- | :--- | :--- |
| $Q_{12}$ | Nonparticipating | Ongoing correct | initial certification in review month (ACTNTYPE=1 and LASTCERT=0) <br> and correct in review month (STATUS=1) |
| $Q_{22}$ | Ongoing correct | Ongoing correct | certification not expiring (LASTCERT<CERTMTH-1) and correct in <br> review month (STATUS=1) |
| $Q_{32}$ | Ongoing error | Ongoing correct | certification not expiring (LASTCERT<CERTMTH-1) and correct in <br> review month (STATUS=1) <br> interim action in review month (ACTNTYPE=5 and LASTACT=0) |
| $Q_{42}$ | Expiring correct | Ongoing correct | derived as $R_{4}-Q_{41}-Q_{43} ;$ see text |

Several alternate assumptions were tested. One was to assume that cases entering recertification as correctly paid (i.e., the expiring correct cases) should emerge upon recertification with an error rate equal to that of cases emerging from initial certification (later defined as the "first-month error rate"). We found, however, that this assumption, when used to compute the $\mathrm{Q}_{52}$ cell count, resulted in negative first-column entries in the national-level matrix, for several of the fiscal years under analysis. We did not adopt this assumption (nor several others that yielded similar results), as negative cell entries also appeared among many state-level matrices.

The assumption used in deriving $\mathrm{Q}_{52}$ minimized the extent of anomalous (negative) cell counts. This assumption might appear to inadequately capture the effectiveness of recertifications in removing errors from the active caseload. To the contrary, as will be seen later, this assumption yields rates of "error survival" that are one-third or one-fourth as large among expiring error cases (error cases subject to recertification) as among ongoing error cases (error cases not subject to recertification). Stated otherwise, a recertification will be shown to reduce very substantially (i.e., by more than one-half) the likelihood that an active error case remains so the following month.

## Nonparticipating Households

The first column is comprised of households not participating in the Food Stamp Program in the review month. This includes the overwhelming majority of total U.S. households. The column total $\left(\mathrm{C}_{1}\right)$ is computed as the residual after subtracting the other four column totals from the total population $\left(\mathrm{Q}-\mathrm{C}_{2}-\mathrm{C}_{3}-\mathrm{C}_{4}-\mathrm{C}_{5}\right)$. None of the cell entries in this column can be derived directly from the active case QC data, as nonparticipating households are not observed in active case reviews, by definition. Although there is a negative case action QC system for food stamp denials and terminations, the associated review process is not sufficient to establish whether a terminated case was correctly or incorrectly paid in its last month of participation. ${ }^{8}$

We determine the prior-month status of each case, and thus its contribution to the cell counts in the first column, as follows (and as shown in Exhibit B-5):

- nonparticipating in prior month-The $\mathrm{Q}_{11}$ cell consists of households that were nonparticipants in both the review month and the preceding month, computed as the first-row residual $\left(\mathrm{R}_{1}-\mathrm{Q}_{12}-\mathrm{Q}_{13}\right)$.

[^24]Exhibit B-5: Derivation of First-Column Cell Counts: Nonparticipating Households

| Cell Count | Prior-month <br> ("March") status | Review-month <br> ("April") status | Criteria (derived as row or column residuals, as explained in text) |
| :---: | :--- | :--- | :--- |
| $Q_{11}$ | Nonparticipating | Nonparticipating | derived as $R_{1}-Q_{12}-Q_{13}$; see text |
| $Q_{21}$ | Ongoing correct | Nonparticipating | derived as $R_{2}-Q_{22}-Q_{23}-Q_{24}-Q_{25}$; see text |
| $Q_{31}$ | Ongoing error | Nonparticipating | derived as $R_{3}-Q_{32}-Q_{33}-Q_{34}-Q_{35}$; see text |
| $Q_{41}$ | Expiring correct | Nonparticipating | derived as $C_{1}-Q_{11}-Q_{21}-Q_{31}-Q_{51}$; see text |
| $Q_{51}$ | Expiring error | Nonparticipating | derived as $R_{5}-Q_{52}-Q_{53} ;$ see text |

- ongoing correct in prior month-The $\mathrm{Q}_{21}$ cell consists of households that have just left the caseload having previously been correctly paid in the midst of a certification period, computed as the second-row residual $\left(R_{2}-Q_{22}-Q_{23}-Q_{24}-\right.$ $\mathrm{Q}_{25}$ ).
- ongoing error in prior month—The $\mathrm{Q}_{31}$ cell consists of households that have just left the caseload having previously been incorrectly paid in the midst of a certification period, computed as the third-row residual $\left(\mathrm{R}_{3}-\mathrm{Q}_{32}-\mathrm{Q}_{33}-\mathrm{Q}_{34}-\right.$ $\mathrm{Q}_{35}$ ).
- expiring correct in prior month-The $\mathrm{Q}_{41}$ cell consists of households that have just left the caseload having previously been correctly paid at the end of a certification period, computed as the first-column residual $\left(\mathrm{C}_{1}-\mathrm{Q}_{11}-\mathrm{Q}_{21}-\mathrm{Q}_{31}-\right.$ $\left.\mathrm{Q}_{51}\right) .{ }^{9}$
- expiring error in prior month-The $\mathrm{Q}_{51}$ cell consists of households that have just left the caseload having previously been incorrectly paid at the end of a certification period, computed as the fifth-row residual $\left(R_{5}-Q_{52}-Q_{53}\right)$.

Because the first-column entries are computed as row residuals, based on cell counts in the other columns that are subject to sampling error, it is possible for the model to yield residuals that are negative in value. That is, for any given row, it is possible for the summed entries of the second through fifth cells to exceed the corresponding row total. Such first-column negative values are typically small, and we have imputed them as zero. This imputation (an upward change from a negative value) requires an offsetting downward adjustment elsewhere in the first column, to leave the first-column total unchanged. We make this downward adjustment in cell $\mathrm{Q}_{11}$, to minimize the effect on the transition probabilities and error rate indicators that are derived from the cell counts. In the exhibits of this report, all imputed zero values are identified by asterisks to distinguish them from entries that are zero by definition or by estimate.

[^25]
# Appendix C National Estimates of Cell Counts and Transition Probabilities 

[In this appendix, asterisked entries in the first column indicate imputed zero values for cells in which initial estimates yielded a negative entry. See Chapter 5 for explanation. Note: Row entries may not sum to the indicated row total due to rounding.]

Exhibit C-1: Cell Counts, 1998

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households (in thousands) |  |  |  |  |  |
| Nonparticipating | 93,636 | 502 | 70 | 0 | 0 | 94,207 |
| Ongoing correct | 0* | 4,798 | 490 | 764 | 102 | 6,154 |
| Ongoing error | 362 | 198 | 516 | 1 | 97 | 1,174 |
| Expiring correct | 187 | 531 | 71 | 0 | 0 | 789 |
| Expiring error | 147 | 34 | 22 | 0 | 0 | 204 |
| Total | 94,332 | 6,064 | 1,168 | 765 | 199 | 102,528 |
|  |  | Cell cou | ouseholds | rnings (in |  |  |


| Nonparticipating | 78,811 | 150 | 35 | 0 | 0 | 78,995 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 0 * | 930 | 233 | 234 | 58 | 1,455 |
| Ongoing error | 187 | 95 | 146 | 0 | 37 | 464 |
| Expiring correct | 26 | 171 | 41 | 0 | 0 | 238 |
| Expiring error | 65 | 20 | 11 | 0 | 0 | 96 |
| Total | 79,090 | 1,365 | 466 | 234 | 94 | 81,248 |


| Nonparticipating | 14,757 | 352 | 35 | 0 | 0 | 15,144 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 68 | 3,868 | 256 | 530 | 44 | 4,768 |
| Ongoing error | 174 | 104 | 371 | 0 | 60 | 710 |
| Expiring correct | 162 | 360 | 30 | 0 | 0 | 551 |
| Expiring error | 81 | 16 | 11 | 0 | 0 | 108 |
| Total | 15,242 | 4,700 | 703 | 531 | 105 | 21,280 |

Exhibit C-2: Cell Counts, 1999

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households (in thousands) |  |  |  |  |  |
| Nonparticipating | 95,623 | 518 | 63 | 0 | 0 | 96,204 |
| Ongoing correct | 276 | 4,431 | 390 | 678 | 67 | 5,841 |
| Ongoing error | 158 | 318 | 438 | 1 | 78 | 993 |
| Expiring correct | 123 | 499 | 65 | 0 | 0 | 687 |
| Expiring error | 82 | 48 | 20 | 0 | 0 | 150 |
| Total | 96,262 | 5,813 | 975 | 679 | 145 | 103,874 |
| Cell counts for households with earnings (in thousands) |  |  |  |  |  |  |
| Nonparticipating | 80,367 | 153 | 31 | 0 | 0 | 80,551 |
| Ongoing correct | 36 | 869 | 188 | 219 | 42 | 1,354 |
| Ongoing error | 89 | 142 | 133 | 1 | 39 | 404 |
| Expiring correct | 26 | 157 | 38 | 0 | 0 | 221 |
| Expiring error | 45 | 29 | 7 | 0 | 0 | 82 |
| Total | 80,563 | 1,351 | 397 | 220 | 80 | 82,611 |
| Cell counts for households without earnings (in thousands) |  |  |  |  |  |  |
| Nonparticipating | 15,257 | 365 | 32 | 0 | 0 | 15,653 |
| Ongoing correct | 240 | 3,561 | 202 | 459 | 25 | 4,487 |
| Ongoing error | 69 | 176 | 304 | 0 | 39 | 589 |
| Expiring correct | 97 | 341 | 27 | 0 | 0 | 466 |
| Expiring error | 35 | 20 | 12 | 0 | 0 | 68 |
| Total | 15,699 | 4,463 | 578 | 459 | 64 | 21,263 |

Exhibit C-3: Cell Counts, 2000

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households (in thousands) |  |  |  |  |  |
| Nonparticipating | 96,812 | 521 | 51 | 0 | 0 | 97,384 |
| Ongoing correct | 320 | 4,259 | 346 | 673 | 64 | 5,661 |
| Ongoing error | 70 | 328 | 368 | 3 | 74 | 845 |
| Expiring correct | 147 | 480 | 50 | 0 | 0 | 677 |
| Expiring error | 67 | 54 | 18 | 0 | 0 | 139 |
| Total | 97,416 | 5,643 | 832 | 676 | 138 | 104,705 |
| Cell counts for households with earnings (in thousands) |  |  |  |  |  |  |
| Nonparticipating | 81,990 | 173 | 27 | 0 | 0 | 82,190 |
| Ongoing correct | 63 | 811 | 177 | 235 | 42 | 1,327 |
| Ongoing error | 72 | 140 | 112 | 2 | 32 | 357 |
| Expiring correct | 36 | 169 | 31 | 0 | 0 | 235 |
| Expiring error | 38 | 30 | 7 | 0 | 0 | 74 |
| Total | 82,199 | 1,323 | 353 | 236 | 74 | 84,184 |
| Cell counts for households without earnings (in thousands) |  |  |  |  |  |  |
| Nonparticipating | 14,821 | 347 | 24 | 0 | 0 | 15,192 |
| Ongoing correct | 257 | 3,448 | 169 | 438 | 22 | 4,333 |
| Ongoing error | 0* | 188 | 257 | 2 | 43 | 489 |
| Expiring correct | 111 | 311 | 19 | 0 | 0 | 441 |
| Expiring error | 28 | 25 | 11 | 0 | 0 | 64 |
| Total | 15,217 | 4,320 | 480 | 440 | 64 | 20,521 |

Exhibit C-4: Cell Counts, 2001

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households (in thousands) |  |  |  |  |  |
| Nonparticipating | 98,385 | 574 | 57 | 0 | 0 | 99,015 |
| Ongoing correct | 485 | 4,302 | 337 | 633 | 56 | 5,814 |
| Ongoing error | 23 | 380 | 360 | 2 | 63 | 829 |
| Expiring correct | 72 | 509 | 58 | 0 | 0 | 640 |
| Expiring error | 51 | 56 | 14 | 0 | 0 | 121 |
| Total | 99,015 | 5,821 | 827 | 636 | 119 | 106,418 |
|  |  | Cell cou | ouseholds | rnings (in |  |  |


| Nonparticipating | 83,049 | 190 | 29 | 0 | 0 | 83,268 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 141 | 846 | 160 | 211 | 29 | 1,387 |
| Ongoing error | 43 | 159 | 102 | 1 | 27 | 331 |
| Expiring correct | 10 | 169 | 33 | 0 | 0 | 212 |
| Expiring error | 25 | 28 | 4 | 0 | 0 | 58 |
| Total | 83,268 | 1,391 | 329 | 212 | 56 | 85,257 |


| Nonparticipating | 15,318 | 384 | 27 | 0 | 0 | 15,729 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 344 | 3,456 | 177 | 422 | 27 | 4,427 |
| Ongoing error | 0 * | 221 | 259 | 1 | 36 | 517 |
| Expiring correct | 60 | 340 | 25 | 0 | 0 | 425 |
| Expiring error | 25 | 28 | 10 | 0 | 0 | 63 |
| Total | 15,747 | 4,430 | 498 | 423 | 63 | 21,161 |

Exhibit C-5: Transition Probabilities, 1998

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.994 | 0.005 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000 * | 0.780 | 0.080 | 0.124 | 0.017 | 1.000 |
| Ongoing error | 0.308 | 0.169 | 0.440 | 0.000 | 0.083 | 1.000 |
| Expiring correct | 0.237 | 0.673 | 0.090 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.724 | 0.169 | 0.107 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.998 | 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000 * | 0.639 | 0.161 | 0.161 | 0.040 | 1.000 |
| Ongoing error | 0.404 | 0.204 | 0.314 | 0.000 | 0.079 | 1.000 |
| Expiring correct | 0.110 | 0.718 | 0.171 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.681 | 0.204 | 0.115 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.974 | 0.023 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.014 | 0.811 | 0.054 | 0.111 | 0.009 | 1.000 |
| Ongoing error | 0.246 | 0.146 | 0.522 | 0.001 | 0.085 | 1.000 |
| Expiring correct | 0.293 | 0.652 | 0.054 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.754 | 0.147 | 0.099 | 0.000 | 0.000 | 1.000 |

Exhibit C-6: Transition Probabilities, 1999

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.994 | 0.005 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.047 | 0.759 | 0.067 | 0.116 | 0.011 | 1.000 |
| Ongoing error | 0.159 | 0.320 | 0.441 | 0.001 | 0.079 | 1.000 |
| Expiring correct | 0.178 | 0.727 | 0.095 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.546 | 0.321 | 0.133 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.998 | 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.027 | 0.642 | 0.139 | 0.162 | 0.031 | 1.000 |
| Ongoing error | 0.220 | 0.351 | 0.330 | 0.002 | 0.096 | 1.000 |
| Expiring correct | 0.119 | 0.711 | 0.170 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.555 | 0.353 | 0.092 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.975 | 0.023 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.054 | 0.794 | 0.045 | 0.102 | 0.006 | 1.000 |
| Ongoing error | 0.117 | 0.299 | 0.517 | 0.001 | 0.066 | 1.000 |
| Expiring correct | 0.209 | 0.732 | 0.059 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.518 | 0.300 | 0.182 | 0.000 | 0.000 | 1.000 |

Exhibit C-7: Transition Probabilities, 2000

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.994 | 0.005 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.056 | 0.752 | 0.061 | 0.119 | 0.011 | 1.000 |
| Ongoing error | 0.083 | 0.389 | 0.436 | 0.004 | 0.088 | 1.000 |
| Expiring correct | 0.217 | 0.709 | 0.074 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.480 | 0.393 | 0.127 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.998 | 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.047 | 0.611 | 0.133 | 0.177 | 0.032 | 1.000 |
| Ongoing error | 0.201 | 0.392 | 0.313 | 0.005 | 0.089 | 1.000 |
| Expiring correct | 0.154 | 0.716 | 0.130 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.511 | 0.397 | 0.092 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.976 | 0.023 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.059 | 0.796 | 0.039 | 0.101 | 0.005 | 1.000 |
| Ongoing error | 0.000 * | 0.385 | 0.525 | 0.003 | 0.087 | 1.000 |
| Expiring correct | 0.251 | 0.705 | 0.044 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.442 | 0.389 | 0.169 | 0.000 | 0.000 | 1.000 |

Exhibit C-8: Transition Probabilities, 2001

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.994 | 0.006 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.083 | 0.740 | 0.058 | 0.109 | 0.010 | 1.000 |
| Ongoing error | 0.028 | 0.459 | 0.435 | 0.003 | 0.076 | 1.000 |
| Expiring correct | 0.113 | 0.796 | 0.091 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.420 | 0.462 | 0.119 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.997 | 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.102 | 0.610 | 0.116 | 0.152 | 0.021 | 1.000 |
| Ongoing error | 0.128 | 0.480 | 0.307 | 0.004 | 0.081 | 1.000 |
| Expiring correct | 0.049 | 0.795 | 0.157 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.440 | 0.484 | 0.076 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.974 | 0.024 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.078 | 0.781 | 0.040 | 0.095 | 0.006 | 1.000 |
| Ongoing error | 0.000 * | 0.428 | 0.501 | 0.002 | 0.070 | 1.000 |
| Expiring correct | 0.141 | 0.800 | 0.059 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.395 | 0.447 | 0.158 | 0.000 | 0.000 | 1.000 |

## Appendix D State-by-State Estimates of Cell Counts and Transition Probabilities

[In this appendix, asterisked entries in the first column indicate imputed zero values for cells in which initial estimates yielded a negative entry. See Chapter 5 for explanation. Note: Row entries may not sum to the indicated row total due to rounding.]

State: Alabama

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 1,548,059 | 8,953 | 1,770 | 0 | 0 | 1,558,783 |
| Ongoing correct | 1,705 | 102,970 | 9,900 | 11,329 | 841 | 126,746 |
| Ongoing error | 7,547 | 4,174 | 8,613 | 0 | 993 | 21,327 |
| Expiring correct | 631 | 9,761 | 1,023 | 0 | 0 | 11,416 |
| Expiring error | 1,461 | 365 | 37 | 0 | 0 | 1,863 |
| Total | 1,559,404 | 126,223 | 21,344 | 11,329 | 1,834 | 1,720,134 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 1,274,413 | 3,379 | 1,299 | 0 | 0 | 1,279,091 |
| Ongoing correct | 2,549 | 27,118 | 4,425 | 1,917 | 75 | 36,084 |
| Ongoing error | 3,641 | 2,340 | 3,526 | 0 | 225 | 9,732 |
| Expiring correct | 0* | 2,953 | 469 | 0 | 0 | 3,422 |
| Expiring error | 194 | 73 | 37 | 0 | 0 | 305 |
| Total | 1,280,797 | 35,863 | 9,756 | 1,917 | 300 | 1,328,634 |
|  | Cell counts for households without earnings |  |  |  |  |  |
| Nonparticipating | 271,280 | 5,574 | 471 | 0 | 0 | 277,325 |
| Ongoing correct | 0* | 75,853 | 5,475 | 9,411 | 766 | 91,505 |
| Ongoing error | 3,913 | 1,834 | 5,087 | 0 | 768 | 11,602 |
| Expiring correct | 2,101 | 6,853 | 554 | 0 | 0 | 9,509 |
| Expiring error | 1,313 | 246 | 0 | 0 | 0 | 1,559 |
| Total | 278,607 | 90,360 | 11,588 | 9,411 | 1,534 | 391,501 |


|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Non- | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error | Total |

## Transition probabilities for total households

| Nonparticipating | 0.993 | 0.006 | 0.001 | 0.000 | 0.000 | 0.007 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Ongoing correct | 0.013 | 0.812 | 0.078 | 0.089 | 1.000 |  |
| Ongoing error | 0.354 | 0.196 | 0.404 | 0.000 | 0.047 |  |
| Expiring correct | 0.055 | 0.855 | 0.090 | 0.000 | 0.000 |  |
| Expiring error | 0.784 | 0.196 | 0.020 | 0.000 | 0.000 |  |

Expiring error 0.784
0.196
0.020
0.000
0.000
1.000

Transition probabilities for households with earnings

| 0.003 | 0.001 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- |
| 0.752 | 0.123 | 0.053 | 0.002 | 1.000 |
| 0.240 | 0.362 | 0.000 | 0.023 | 1.000 |
| 0.863 | 0.137 | 0.000 | 0.000 | 1.000 |
| 0.240 | 0.123 | 0.000 | 0.000 | 1.000 |
|  |  |  |  |  |
| sition probabilities for households without earnings |  |  |  |  |


| Nonparticipating | 0.978 |
| :--- | :--- |
| Ongoing correct | $0.000^{*}$ |
| Ongoing error | 0.337 |
| Expiring correct | 0.221 |
| Expiring error | 0.842 |


| 0.020 | 0.002 |
| :--- | :--- |
| 0.829 | 0.060 |
| 0.158 | 0.438 |
| 0.721 | 0.058 |
| 0.158 | 0.000 |

0.000
0.103
0.000
0.000
0.000
0.000

Nonparticipating
0.996
0.071
0.374
0.000*
0.637
$\begin{array}{cc}\text { Transition probabilities for ho } \\ & \\ 0.020 & 0.002 \\ 0.829 & 0.060 \\ 0.158 & 0.438 \\ 0.721 & 0.058 \\ 0.158 & 0.000\end{array}$

## Onoing correc

Ongoing error
Expiring correct
Expiring error
0.221
0.842

State: Alaska

|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Current-month status | Non- <br> participating | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error | Total |

Nonparticipating
Ongoing correct
Ongoing error
Expiring correct
Expiring error

Cell counts for total households

| Nonparticipating | 208,226 | 1,301 | 286 | 0 | 0 | 209,814 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 1,114 | 6,805 | 2,068 | 258 | 88 | 10,333 |
| Ongoing error | 838 | 1,717 | 413 | 0 | 0 | 2,969 |
| Expiring correct | 0* | 450 | 197 | 0 | 0 | 647 |
| Expiring error | 39 | 54 | 0 | 0 | 0 | 93 |
| Total | 210,217 | 10,327 | 2,965 | 258 | 88 | 223,856 |

## Cell counts for households with earnings

| Nonparticipating | 195,498 | 373 | 85 | 0 | 195,956 |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 36 | 1,599 | 1,169 | 45 | 49 |
| Ongoing error | 556 | 781 | 120 | 0 | 0 |
| Expiring correct | $0^{*}$ | 113 | 83 | 0 | 0 |
| Expiring error | 23 | 27 | 0 | 0 | 197 |
| Total | 196,112 | 2,893 | 1,457 | 457 |  |
|  |  |  |  | 40 | 40 |

Cell counts for households without earnings

| Nonparticipating | 12,735 | 928 | 201 | 0 | 0 | 13,864 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 1,072 | 5,207 | 900 | 213 | 39 | 7,431 |
| Ongoing error | 282 | 936 | 293 | 0 | 0 | 1,510 |
| Expiring correct | 0* | 337 | 114 | 0 | 0 | 451 |
| Expiring error | 16 | 27 | 0 | 0 | 0 | 43 |
| Total | 14,105 | 7,434 | 1,508 | 213 | 39 | 23,299 |


|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Non- | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error | Total |

## Transition probabilities for total households

| Nonparticipating | 0.992 | 0.006 | 0.001 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Ongoing correct | 0.108 | 0.659 | 0.200 | 0.025 | 0.008 |  |
| Ongoing error | 0.282 | 0.578 | 0.139 | 0.000 | 1.000 |  |
| Expiring correct | $0.000^{*}$ | 0.695 | 0.305 | 0.000 |  |  |
| Expiring error | 0.422 | 0.578 | 0.000 | 0.000 | 0.000 | 1.000 |
|  |  |  |  |  | 0.000 | 0.000 |

Expiring error 0.422
0.578
0.000
$0.000 \quad 0.000$
1.000

Transition probabilities for households with earnings

| 0.998 | 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 0.012 | 0.552 | 0.403 | 0.016 | 0.017 | 1.000 |
| 0.381 | 0.536 | 0.083 | 0.000 | 0.000 | 1.000 |
| $0.000^{*}$ | 0.577 | 0.423 | 0.000 | 0.000 | 1.000 |
| 0.464 | 0.536 | 0.000 | 0.000 | 0.000 | 1.000 |

Transition probabilities for households without earnings

| Nonparticipating | 0.919 | 0.067 | 0.015 | 0.000 | 0.000 | 0.005 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Ongoing correct | 0.144 | 0.701 | 0.121 | 0.029 | 1.000 |  |
| Ongoing error | 0.187 | 0.620 | 0.194 | 0.000 | 1.000 |  |
| Expiring correct | $0.000^{*}$ | 0.746 | 0.254 | 0.000 | 0.000 |  |
| Expiring error | 0.380 | 0.620 | 0.000 | 0.000 | 1.000 |  |
|  |  |  |  |  |  |  |

State: Arizona

|  | Next-month status |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | ---: | :--- |
| Current-month status | Non- <br> participating | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error | Total |


|  |  |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | $1,784,093$ | 14,133 | 939 | 0 | 0 | $1,799,164$ |
| Ongoing correct | 9,195 | 47,988 | 3,129 | 15,390 | 1,764 | 77,466 |
| Ongoing error | 2,284 | 2,306 | 1,315 | 0 | 617 | 6,521 |
| Expiring correct | 2,884 | 11,824 | 818 | 0 | 0 | 15,526 |
| Expiring error | 1,371 | 853 | 189 | 0 | 0 | 2,413 |
| Total | $1,799,826$ | 77,103 | 6,390 | 15,390 | 2,381 | $1,901,090$ |

## Cell counts for total households

## Cell counts for households with earnings

| Nonparticipating | $1,441,677$ | 5,081 | 477 | 0 | $1,447,235$ |  |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 3,676 | 12,813 | 1,759 | 4,733 | 1,077 | 24,058 |
| Ongoing error | 1,524 | 996 | 477 | 0 | 271 | 0 |
| Expiring correct | $0^{*}$ | 4,702 | 407 | 0 | 0,109 |  |
| Expiring error | 850 | 419 | 104 | 0 | 0 |  |
| Total | $1,447,727$ | 24,011 | 3,225 | 4,733 | 1,373 |  |
|  |  |  |  |  |  | $0,481,043$ |


| Nonparticipating | 342,099 | 9,052 | 462 | 0 | 351,612 |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 5,520 | 35,175 | 1,370 | 53,409 |  |
| Ongoing error | 760 | 1,309 | 838 | 088 | 3,253 |
| Expiring correct | 3,184 | 7,138 | 411 | 0 | 10,733 |
| Expiring error | 536 | 418 | 85 | 0 | 0 |
| Total | 352,099 | 53,092 | 3,165 | 0 | 0 |
|  |  |  | 10,656 | 1,033 | 420,047 |


|  | Next-month status |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Current-month status | Non- <br> participating | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error | Total |


| Nonparticipating | 0.992 | 0.008 | 0.001 | 0.000 | 0.000 | 1.000 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 0.119 | 0.619 | 0.040 | 0.199 | 0.023 | 1.000 |
| Ongoing error | 0.350 | 0.354 | 0.202 | 0.000 | 0.095 | 1.000 |
| Expiring correct | 0.186 | 0.762 | 0.053 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.568 | 0.354 | 0.078 | 0.000 | 0.000 | 1.000 |
|  | Transition probabilities for households with earnings |  |  |  |  |  |
| Nonparticipating | 0.996 | 0.004 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.153 | 0.533 | 0.073 | 0.197 | 0.045 | 1.000 |
| Ongoing error | 0.466 | 0.305 | 0.146 | 0.000 | 0.083 | 1.000 |
| Expiring correct | 0.000* | 0.920 | 0.080 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.619 | 0.305 | 0.076 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.973 | 0.026 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.103 | 0.659 | 0.026 | 0.200 | 0.013 | 1.000 |
| Ongoing error | 0.234 | 0.403 | 0.258 | 0.000 | 0.106 | 1.000 |
| Expiring correct | 0.297 | 0.665 | 0.038 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.516 | 0.403 | 0.082 | 0.000 | 0.000 | 1.000 |

Expiring error

| Nonparticipating | 0.992 | 0.008 | 0.001 | 0.000 | 0.000 | 1.000 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 0.119 | 0.619 | 0.040 | 0.199 | 0.023 | 1.000 |
| Ongoing error | 0.350 | 0.354 | 0.202 | 0.000 | 0.095 | 1.000 |
| Expiring correct | 0.186 | 0.762 | 0.053 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.568 | 0.354 | 0.078 | 0.000 | 0.000 | 1.000 |
|  | Transition probabilities for households with earnings |  |  |  |  |  |
| Nonparticipating | 0.996 | 0.004 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.153 | 0.533 | 0.073 | 0.197 | 0.045 | 1.000 |
| Ongoing error | 0.466 | 0.305 | 0.146 | 0.000 | 0.083 | 1.000 |
| Expiring correct | 0.000* | 0.920 | 0.080 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.619 | 0.305 | 0.076 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.973 | 0.026 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.103 | 0.659 | 0.026 | 0.200 | 0.013 | 1.000 |
| Ongoing error | 0.234 | 0.403 | 0.258 | 0.000 | 0.106 | 1.000 |
| Expiring correct | 0.297 | 0.665 | 0.038 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.516 | 0.403 | 0.082 | 0.000 | 0.000 | 1.000 |

0.568

Transition probabilities for total households

State: Arkansas

|  | Next-month status |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Current-month status | Non- <br> participating | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error | Total |


|  | Cell counts for total households |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Nonparticipating | 921,020 | 8,443 | 427 | 0 | 0 | 929,890 |
| Ongoing correct | 19,314 | 68,384 | 2,779 | 1,457 | 0 | 91,934 |
| Ongoing error | 0* | 9,691 | 3,522 | 0 | 104 | 13,317 |
| Expiring correct | 0 * | 5,255 | 347 | 0 | 0 | 5,602 |
| Expiring error | 0* | 148 | 61 | 0 | 0 | 209 |
| Total | 940,334 | 91,921 | 7,136 | 1,457 | 104 | 1,040,952 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 780,026 | 2,787 | 268 | 0 | 0 | 783,081 |
| Ongoing correct | 6,505 | 17,114 | 1,395 | 264 | 0 | 25,279 |
| Ongoing error | 0* | 3,705 | 1,075 | 0 | 42 | 4,822 |
| Expiring correct | 0 * | 1,694 | 204 | 0 | 0 | 1,898 |
| Expiring error | 0* | 57 | 61 | 0 | 0 | 118 |
| Total | 786,531 | 25,357 | 3,003 | 264 | 42 | 815,198 |

Cell counts for total households

Cell counts for households without earnings

| Nonparticipating | 140,692 | 5,656 | 159 | 0 | 146,507 |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 13,111 | 51,270 | 1,384 | 0 | 06,957 |
| Ongoing error | $0^{*}$ | 5,986 | 2,448 | 0,495 |  |
| Expiring correct | $0^{*}$ | 3,561 | 143 | 0 | 0 |
| Expiring error | $0^{*}$ | 92 | 0 | 0 | 0 |
| Total | 153,803 | 66,564 | 4,133 | 0 | 0 |
|  |  |  | 1,193 | 61 | 225,754 |


|  | Next-month status |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Current-month status | Non- <br> participating | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error | Total |


| Nonparticipating | 0.990 |
| :--- | :--- |
| Ongoing correct | 0.210 |
| Ongoing error | $0.000^{*}$ |
| Expiring correct | $0.000^{*}$ |
| Expiring error | $0.000^{*}$ |


| 0.009 | 0.000 |
| :--- | :--- |
| 0.744 | 0.030 |
| 0.728 | 0.264 |
| 0.938 | 0.062 |
| 0.707 | 0.293 |

0.000
0.016
0.000
0.000
0.000

| 0.000 | 1.000 |
| :--- | :--- |
| 0.000 | 1.000 |
| 0.008 | 1.000 |
| 0.000 | 1.000 |
| 0.000 | 1.000 |

Transition probabilities for households with earnings

| 0.996 | 0.004 | 0.000 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 0.257 | 0.677 | 0.055 | 0.010 | 0.000 | 1.000 |
| $0.000^{*}$ | 0.768 | 0.223 | 0.000 | 0.009 | 1.000 |
| $0.000^{*}$ | 0.893 | 0.107 | 0.000 | 0.000 | 1.000 |
| $0.000^{*}$ | 0.479 | 0.521 | 0.000 | 0.000 | 1.000 |
|  |  |  |  |  |  |
|  |  |  |  |  |  |
|  | Transition probabilities for households without earnings |  |  |  |  |


| Nonparticipating | 0.960 | 0.039 | 0.001 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Ongoing correct | 0.196 | 0.766 | 0.021 | 0.018 | 0.000 | 1.000 |
| Ongoing error | $0.000^{*}$ | 0.705 | 0.288 | 0.000 | 0.007 | 1.000 |
| Expiring correct | $0.000^{*}$ | 0.961 | 0.039 | 0.000 | 0.000 | 1.000 |
| Expiring error | $0.000^{*}$ | 1.000 | 0.000 | 0.000 | 0.000 | 1.000 |

State: California

|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Non- | Ongoing | Ongoing | Expiring | Expiring |  |
| Current-month status | participating | correct | error | correct | error | Total |


|  | Cell counts for total households |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Nonparticipating | 10,427,591 | 29,742 | 4,251 | 0 | 0 | 10,461,583 |
| Ongoing correct | 0* | 415,381 | 85,524 | 77,217 | 15,840 | 593,962 |
| Ongoing error | 83,471 | 1,595 | 50,617 | 0 | 11,265 | 146,948 |
| Expiring correct | 50,754 | 23,592 | 4,061 | 0 | 0 | 78,406 |
| Expiring error | 24,945 | 295 | 1,948 | 0 | 0 | 27,189 |
| Total | 10,586,761 | 470,605 | 146,401 | 77,217 | 27,104 | 11,308,088 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 9,072,522 | 5,072 | 1,679 | 0 | 0 | 9,079,273 |
| Ongoing correct | 0* | 118,621 | 38,703 | 22,865 | 6,187 | 186,376 |
| Ongoing error | 40,867 | 348 | 14,478 | 0 | 2,602 | 58,294 |
| Expiring correct | 13,634 | 6,827 | 2,522 | 0 | 0 | 22,983 |
| Expiring error | 8,562 | 53 | 297 | 0 | 0 | 8,912 |
| Total | 9,135,585 | 130,920 | 57,679 | 22,865 | 8,788 | 9,355,837 |
|  | Cell counts for households without earnings |  |  |  |  |  |
| Nonparticipating | 1,354,783 | 24,670 | 2,572 | 0 | 0 | 1,382,025 |
| Ongoing correct | 0* | 296,760 | 46,821 | 54,352 | 9,653 | 407,586 |
| Ongoing error | 42,744 | 1,248 | 36,139 | 0 | 8,663 | 88,793 |
| Expiring correct | 37,252 | 16,750 | 1,539 | 0 | 0 | 55,541 |
| Expiring error | 16,397 | 257 | 1,652 | 0 | 0 | 18,306 |
| Total | 1,451,176 | 339,685 | 88,722 | 54,352 | 18,316 | 1,952,251 |


|  | Next-month status |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | ---: | :--- |
|  | Non- | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error | Total |

Transition probabilities for total households

|  | 0.997 | 0.003 | 0.000 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Nonparticipating | $0.000^{*}$ | 0.699 | 0.144 | 0.130 | 0.027 | 1.000 |
| Ongoing correct | 0.568 | 0.011 | 0.344 | 0.000 | 1.000 |  |
| Ongoing error | 0.647 | 0.301 | 0.052 | 0.077 |  |  |
| Expiring correct | 0.917 | 0.011 | 0.072 | 0.000 | 0.000 | 1.000 |
| Expiring error |  |  |  | 0.000 | 1.000 |  |

Expiring error
$0.917 \quad 0.011$
Transition probabilities for households with earnings

| 0.999 | 0.001 | 0.000 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| $0.000^{*}$ | 0.636 | 0.208 | 0.123 | 0.033 | 1.000 |
| 0.701 | 0.006 | 0.248 | 0.000 | 0.045 | 1.000 |
| 0.593 | 0.297 | 0.110 | 0.000 | 0.000 | 1.000 |
| 0.961 | 0.006 | 0.033 | 0.000 | 0.000 | 1.000 |
|  |  |  |  |  |  |


| Nonparticipating | 0.980 | 0.018 | 0.002 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Ongoing correct | $0.000^{*}$ | 0.728 | 0.115 | 0.133 | 0.024 |  |
| Ongoing error | 0.481 | 0.014 | 0.407 | 0.000 | 0.098 |  |
| Expiring correct | 0.671 | 0.302 | 0.028 | 0.000 | 0.000 |  |
| Expiring error | 0.896 | 0.014 | 0.090 | 0.000 | 0.000 |  |
|  |  |  |  |  | 1.000 |  |

State: Colorado

|  | Next-month status |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Current-month status | Non- | Ongoing | Ongoing | Expiring | Expiring |
| enrror | correct | error | correct | Total |  |

## Cell counts for total households

| Nonparticipating | 1,588,329 | 7,013 | 654 | 0 | 0 | 1,595,996 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 14,098 | 41,000 | 2,883 | 3,157 | 676 | 61,814 |
| Ongoing error | 0* | 5,209 | 4,327 | 41 | 264 | 9,840 |
| Expiring correct | 0* | 7,522 | 741 | 0 | 0 | 8,263 |
| Expiring error | 117 | 557 | 275 | 0 | 0 | 949 |
| Total | 1,602,543 | 61,302 | 8,880 | 3,197 | 940 | 1,676,862 |

Cell counts for households with earnings

| Nonparticipating | $1,415,249$ | 2,568 | 319 | 0 | $1,418,137$ |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 4,016 | 8,568 | 1,819 | 1,179 | 505 | 16,087 |
| Ongoing error | 853 | 1,795 | 1,532 | 0 | 135 | 0 |
| Expiring correct | $0^{*}$ | 2,850 | 524 | 0 | 3,373 |  |
| Expiring error | 287 | 268 | 89 | 0 | 643 |  |
| Total | $1,420,405$ | 16,048 | 4,283 | 1,179 | 640 | $1,442,556$ |


| Nonparticipating | 171,909 | 4,445 | 334 | 0 | 0 | 176,688 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 10,229 | 32,432 | 1,064 | 1,978 | 171 | 45,874 |
| Ongoing error | 0* | 3,414 | 2,795 | 41 | 129 | 6,378 |
| Expiring correct | 0 * | 4,734 | 217 | 0 | 0 | 4,952 |
| Expiring error | 0* | 228 | 186 | 0 | 0 | 414 |
| Total | 182,138 | 45,253 | 4,597 | 2,018 | 300 | 234,307 |


|  | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Current-month status | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |

## Transition probabilities for total households

| Nonparticipating | 0.995 | 0.004 | 0.000 | 0.000 | 0.000 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Ongoing correct | 0.228 | 0.663 | 0.047 | 0.051 | 0.011 |
| Ongoing error | $0.000^{*}$ | 0.529 | 0.440 | 0.004 | 0.027 |
| Expiring correct | $0.000^{*}$ | 0.910 | 0.090 | 0.000 | 1.000 |
| Expiring error | 0.123 | 0.587 | 0.290 | 0.000 | 1.000 |
|  |  |  | 0.000 | 1.000 |  |

Expiring error
0.123
0.998
0.250
0.198
$0.000^{*}$
0.446

Transition probabilities for households with earnings

| 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- |
| 0.533 | 0.113 | 0.073 | 0.031 | 1.000 |
| 0.416 | 0.355 | 0.000 | 0.031 | 1.000 |
| 0.845 | 0.155 | 0.000 | 0.000 | 1.000 |
| 0.416 | 0.138 | 0.000 | 0.000 | 1.000 |
|  |  |  |  |  |


| Nonparticipating | 0.973 | 0.025 | 0.002 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Ongoing correct | 0.223 | 0.707 | 0.023 | 0.043 | 1.000 |  |
| Ongoing error | $0.000^{*}$ | 0.535 | 0.438 | 0.006 | 0.020 |  |
| Expiring correct | $0.000^{*}$ | 0.956 | 0.044 | 0.000 | 0.000 |  |
| Expiring error | $0.000^{*}$ | 0.551 | 0.449 | 0.000 | 1.000 |  |
|  |  |  |  | 0.000 | 1.000 |  |

State: Connecticut

|  | Next-month status |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | ---: | :--- |
|  | Non- | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error | Total |

## Cell counts for total households

| Nonparticipating | $1,196,412$ | 4,552 | 725 | 0 | $1,201,689$ |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | $0^{*}$ | 58,839 | 4,370 | 5,049 | 08 |
| Ongoing error | 4,642 | 419 | 7,504 | 0 | 840 |
| Expiring correct | 1,504 | 3,232 | 444 | 0 | 0 |
| Expiring error | 728 | 35 | 349 | 0 | 0,404 |
| Total | $1,203,286$ | 67,076 | 13,392 | 5,049 | 1,112 |
|  |  |  |  | 1,093 | $1,289,897$ |


| Nonparticipating | $1,013,458$ | 658 | 339 | 0 | $1,014,456$ |  |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | $0^{*}$ | 5,320 | 1,474 | 297 | 165 | 7,256 |
| Ongoing error | 1,731 | 280 | 1,063 | 0 | 184 |  |
| Expiring correct | $0^{*}$ | 448 | 312 | 0 | 0 | 760 |
| Expiring error | 251 | 30 | 71 | 0 | 0 | 0 |
| Total | $1,015,441$ | 6,737 | 3,259 | 297 | 348 | $1,026,083$ |


| Nonparticipating | 182,566 | 3,894 | 386 | 0 | 186,846 |
| :--- | :---: | ---: | ---: | ---: | ---: |
| Ongoing correct | $0^{*}$ | 53,518 | 2,896 | 4,751 | 81,254 |
| Ongoing error | 2,870 | 139 | 6,441 | 0 | 656 |
| Expiring correct | 1,942 | 2,778 | 133 | 0 | 4,852 |
| Expiring error | 467 | 10 | 278 | 0 | 0 |
| Total | 187,845 | 60,340 | 10,133 | 4,751 | 0 |
|  |  |  |  |  |  |

State: Connecticut

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.996 | 0.004 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.859 | 0.064 | 0.074 | 0.004 | 1.000 |
| Ongoing error | 0.346 | 0.031 | 0.560 | 0.000 | 0.063 | 1.000 |
| Expiring correct | 0.290 | 0.624 | 0.086 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.655 | 0.031 | 0.314 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.999 | 0.001 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.733 | 0.203 | 0.041 | 0.023 | 1.000 |
| Ongoing error | 0.531 | 0.086 | 0.326 | 0.000 | 0.056 | 1.000 |
| Expiring correct | 0.000* | 0.590 | 0.410 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.712 | 0.086 | 0.202 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.977 | 0.021 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.874 | 0.047 | 0.078 | 0.001 | 1.000 |
| Ongoing error | 0.284 | 0.014 | 0.637 | 0.000 | 0.065 | 1.000 |
| Expiring correct | 0.400 | 0.572 | 0.027 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.619 | 0.014 | 0.368 | 0.000 | 0.000 | 1.000 |

State: Delaware

|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Non- | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error | Total |

## Cell counts for total households

| Nonparticipating | 277,625 | 1,450 | 252 | 0 | 0 | 279,327 |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | $0^{*}$ | 7,347 | 938 | 1,867 | 147 | 10,299 |
| Ongoing error | 837 | 271 | 1,071 | 0 | 275 | 2,454 |
| Expiring correct | 1,147 | 592 | 124 | 0 | 0 | 1,863 |
| Expiring error | 336 | 47 | 42 | 0 | 0 | 425 |
| Total | 279,944 | 9,708 | 2,426 | 1,867 | 422 | 294,368 |

## Cell counts for households with earnings

| Nonparticipating | 233,115 | 415 | 95 | 0 | 0 | 233,624 |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | $0^{*}$ | 1,503 | 558 | 488 | 57 | 2,607 |
| Ongoing error | 644 | 59 | 413 | 0 | 81 | 1,197 |
| Expiring correct | 198 | 169 | 113 | 0 | 0 | 481 |
| Expiring error | 134 | 7 | 0 | 0 | 0 | 141 |
| Total | 234,091 | 2,153 | 1,179 | 488 | 138 | 238,051 |

Cell counts for households without earnings

| Nonparticipating | 44,508 | 1,035 | 157 | 0 | 45,700 |
| :--- | :---: | ---: | ---: | ---: | ---: |
| Ongoing correct | $0^{*}$ | 5,844 | 379 | 1,379 | 90 |
| Ongoing error | 193 | 212 | 658 | 0 | 194 |
| Expiring correct | 957 | 415 | 10 | 0 | 0 |
| Expiring error | 194 | 48 | 42 | 0 | 1,257 |
| Total | 45,853 | 7,554 | 1,247 | 1,383 |  |
|  |  |  |  | 284 | 0 |

State: Delaware

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.994 | 0.005 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.713 | 0.091 | 0.181 | 0.014 | 1.000 |
| Ongoing error | 0.341 | 0.111 | 0.436 | 0.000 | 0.112 | 1.000 |
| Expiring correct | 0.616 | 0.318 | 0.066 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.791 | 0.111 | 0.099 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.998 | 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.577 | 0.214 | 0.187 | 0.022 | 1.000 |
| Ongoing error | 0.538 | 0.049 | 0.345 | 0.000 | 0.068 | 1.000 |
| Expiring correct | 0.413 | 0.352 | 0.236 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.951 | 0.049 | 0.000 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.974 | 0.023 | 0.003 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.760 | 0.049 | 0.179 | 0.012 | 1.000 |
| Ongoing error | 0.154 | 0.169 | 0.523 | 0.000 | 0.155 | 1.000 |
| Expiring correct | 0.692 | 0.300 | 0.007 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.683 | 0.169 | 0.148 | 0.000 | 0.000 | 1.000 |

State: District of Columbia

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 204,928 | 1,732 | 171 | 0 | 0 | 206,830 |
| Ongoing correct | 0* | 23,076 | 2,447 | 2,970 | 234 | 28,727 |
| Ongoing error | 2,522 | 165 | 2,354 | 0 | 341 | 5,382 |
| Expiring correct | 629 | 1,943 | 378 | 0 | 0 | 2,950 |
| Expiring error | 507 | 17 | 42 | 0 | 0 | 567 |
| Total | 208,586 | 26,932 | 5,393 | 2,970 | 575 | 244,456 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 194,890 | 154 | 41 | 0 | 0 | 195,085 |
| Ongoing correct | 0* | 1,561 | 713 | 267 | 96 | 2,637 |
| Ongoing error | 774 | 13 | 473 | 0 | 111 | 1,371 |
| Expiring correct | 0* | 167 | 121 | 0 | 0 | 288 |
| Expiring error | 175 | 2 | 28 | 0 | 0 | 204 |
| Total | 195,839 | 1,897 | 1,375 | 267 | 207 | 199,586 |
|  | Cell counts for households without earnings |  |  |  |  |  |
| Nonparticipating | 9,994 | 1,578 | 130 | 0 | 0 | 11,702 |
| Ongoing correct | 0* | 21,514 | 1,735 | 2,702 | 138 | 26,089 |
| Ongoing error | 1,763 | 152 | 1,881 | 0 | 230 | 4,026 |
| Expiring correct | 655 | 1,778 | 257 | 0 | 0 | 2,690 |
| Expiring error | 335 | 14 | 15 | 0 | 0 | 363 |
| Total | 12,747 | 25,035 | 4,017 | 2,702 | 368 | 44,871 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.991 | 0.008 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.803 | 0.085 | 0.103 | 0.008 | 1.000 |
| Ongoing error | 0.469 | 0.031 | 0.437 | 0.000 | 0.063 | 1.000 |
| Expiring correct | 0.213 | 0.659 | 0.128 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.895 | 0.031 | 0.074 | 0.000 | 0.000 | 1.000 |

Transition probabilities for households with earnings

| Nonparticipating | 0.999 | 0.001 | 0.000 | 0.000 | 0.000 | 1.000 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 0.000* | 0.592 | 0.270 | 0.101 | 0.037 | 1.000 |
| Ongoing error | 0.565 | 0.009 | 0.345 | 0.000 | 0.081 | 1.000 |
| Expiring correct | 0.000* | 0.579 | 0.421 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.855 | 0.009 | 0.135 | 0.000 | 0.000 | 1.000 |
|  | Transition probabilities for households without earnings |  |  |  |  |  |
| Nonparticipating | 0.854 | 0.135 | 0.011 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.825 | 0.066 | 0.104 | 0.005 | 1.000 |
| Ongoing error | 0.438 | 0.038 | 0.467 | 0.000 | 0.057 | 1.000 |
| Expiring correct | 0.244 | 0.661 | 0.095 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.922 | 0.038 | 0.040 | 0.000 | 0.000 | 1.000 |

State: Florida

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 5,743,294 | 41,887 | 5,190 | 0 | 0 | 5,790,371 |
| Ongoing correct | 0* | 235,145 | 20,990 | 53,197 | 6,549 | 315,882 |
| Ongoing error | 21,898 | 1,754 | 15,197 | 253 | 6,301 | 45,402 |
| Expiring correct | 18,180 | 32,620 | 3,096 | 0 | 0 | 53,896 |
| Expiring error | 12,060 | 578 | 434 | 0 | 0 | 13,072 |
| Total | 5,795,432 | 311,983 | 44,907 | 53,450 | 12,850 | 6,218,623 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 4,516,931 | 12,972 | 2,586 | 0 | 0 | 4,532,490 |
| Ongoing correct | 0* | 37,797 | 11,539 | 15,829 | 4,363 | 69,527 |
| Ongoing error | 12,042 | 796 | 4,134 | 164 | 3,542 | 20,677 |
| Expiring correct | 2,916 | 11,120 | 2,075 | 0 | 0 | 16,112 |
| Expiring error | 7,446 | 371 | 172 | 0 | 0 | 7,988 |
| Total | 4,539,336 | 63,056 | 20,506 | 15,992 | 7,904 | 4,646,795 |
|  | Cell counts for households without earnings |  |  |  |  |  |
| Nonparticipating | 1,222,533 | 28,915 | 2,604 | 0 | 0 | 1,254,051 |
| Ongoing correct | 3,808 | 197,348 | 9,451 | 37,369 | 2,186 | 250,162 |
| Ongoing error | 9,863 | 958 | 11,063 | 89 | 2,759 | 24,733 |
| Expiring correct | 15,285 | 21,491 | 1,021 | 0 | 0 | 37,797 |
| Expiring error | 4,607 | 215 | 262 | 0 | 0 | 5,085 |
| Total | 1,256,096 | 248,927 | 24,401 | 37,458 | 4,946 | 1,571,828 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{array}{r} \text { Non- } \\ \text { participating } \end{array}$ | Ongoing correct | Ongoing error | Expiring correct | $\begin{gathered} \text { Expiring } \\ \text { error } \end{gathered}$ | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.992 | 0.007 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.744 | 0.066 | 0.168 | 0.021 | 1.000 |
| Ongoing error | 0.482 | 0.039 | 0.335 | 0.006 | 0.139 | 1.000 |
| Expiring correct | 0.337 | 0.605 | 0.057 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.923 | 0.044 | 0.033 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.997 | 0.003 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.544 | 0.166 | 0.228 | 0.063 | 1.000 |
| Ongoing error | 0.582 | 0.038 | 0.200 | 0.008 | 0.171 | 1.000 |
| Expiring correct | 0.181 | 0.690 | 0.129 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.932 | 0.046 | 0.021 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.975 | 0.023 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.015 | 0.789 | 0.038 | 0.149 | 0.009 | 1.000 |
| Ongoing error | 0.399 | 0.039 | 0.447 | 0.004 | 0.112 | 1.000 |
| Expiring correct | 0.404 | 0.569 | 0.027 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.906 | 0.042 | 0.052 | 0.000 | 0.000 | 1.000 |

State: Georgia

|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Ongoing | Ongoing | Expiring | Expiring |  |  |
| Current-month status | Non- | Ongicipating | correct | error | correct | error |

Cell counts for total households

| Nonparticipating | 2,616,092 | 24,609 | 2,663 | 0 | 0 | 2,643,363 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 32,452 | 124,212 | 11,736 | 18,821 | 2,058 | 189,277 |
| Ongoing error | 3,083 | 12,803 | 11,746 | 60 | 2,805 | 30,497 |
| Expiring correct | 0* | 25,125 | 3,177 | 0 | 0 | 28,301 |
| Expiring error | 2,273 | 2,158 | 685 | 0 | 0 | 5,117 |
| Total | 2,653,900 | 188,907 | 30,006 | 18,881 | 4,863 | 2,896,557 |


|  |  | 9,187 | 1,419 | 0 | $2,342,075$ |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Nonparticipating | $2,331,469$ | 21,704 | 5,904 | 46,109 |  |  |
| Ongoing correct | 9,628 | 4,765 | 3,044 | 1,282 | 13,424 |  |
| Ongoing error | 4,364 | $0^{*}$ | 9,726 | 2,617 | 0 | 1,251 |
| Expiring correct | 1,436 | 955 | 299 | 0 | 0 |  |
| Expiring error | $2,346,898$ | 46,337 | 13,284 | 0 | 0 | 2,644 |
| Total |  | 7,591 | 2,533 | $2,416,643$ |  |  |


| Nonparticipating | 283,132 | 15,422 | 1,244 | 0 | 0 | 299,798 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 22,978 | 102,507 | 5,831 | 11,230 | 776 |  |
| Ongoing error | $0^{*}$ | 8,038 | 8,701 | 60 | 1,554 | 18,322 |
| Expiring correct | $0^{*}$ | 15,451 | 559 | 0 | 16,010 |  |
| Expiring error | 892 | 1,151 | 386 | 0 | 0 |  |
| Total | 307,002 | 142,570 | 16,722 | 11,290 | 2,430 |  |
|  |  |  |  | 2,930 | 479,914 |  |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.990 | 0.009 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.171 | 0.656 | 0.062 | 0.099 | 0.011 | 1.000 |
| Ongoing error | 0.101 | 0.420 | 0.385 | 0.002 | 0.092 | 1.000 |
| Expiring correct | 0.000* | 0.888 | 0.112 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.444 | 0.422 | 0.134 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.995 | 0.004 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.209 | 0.471 | 0.128 | 0.165 | 0.028 | 1.000 |
| Ongoing error | 0.325 | 0.355 | 0.227 | 0.000 | 0.093 | 1.000 |
| Expiring correct | 0.000* | 0.788 | 0.212 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.534 | 0.355 | 0.111 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.944 | 0.051 | 0.004 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.160 | 0.715 | 0.041 | 0.078 | 0.005 | 1.000 |
| Ongoing error | 0.000* | 0.438 | 0.474 | 0.003 | 0.085 | 1.000 |
| Expiring correct | 0.000* | 0.965 | 0.035 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.367 | 0.474 | 0.159 | 0.000 | 0.000 | 1.000 |

State: Hawaii

|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Ongoing | Ongoing | Expiring | Expiring |  |  |
| Current-month status | participating | correct | error | correct | error | Total |

Nonparticipating
Ongoing correct
Ongoing error
Expiring correct
Expiring error

Cell counts for total households

| Nonparticipating | 317,567 | 2,109 | 352 | 0 | 0 | 320,028 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 1,620 | 36,364 | 3,385 | 2,639 | 215 | 44,223 |
| Ongoing error | 476 | 3,317 | 3,048 | 0 | 217 | 7,059 |
| Expiring correct | 711 | 1,830 | 136 | 0 | 0 | 2,677 |
| Expiring error | 102 | 200 | 123 | 0 | 0 | 425 |
| Total | 320,477 | 43,821 | 7,043 | 2,639 | 432 | 374,413 |


| Nonparticipating | 297,897 | 449 | 138 | 0 | 0 | 298,484 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 776 | 7,911 | 1,696 | 444 | 92 | 10,918 |
| Ongoing error | 0* | 1,985 | 1,131 | 0 | 76 | 3,192 |
| Expiring correct | 0 * | 447 | 16 | 0 | 0 | 463 |
| Expiring error | 0* | 107 | 77 | 0 | 0 | 184 |
| Total | 298,673 | 10,898 | 3,058 | 444 | 168 | 313,241 |

## Total

Nonparticipating
Ongoing correct
Ongoing error
Expiring correct
Expiring error
Total

| 19,336 | 1,660 | 214 | 0 | 0 | 21,210 |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 973 | 28,453 | 1,688 | 2,196 | 123 | 33,434 |
| 642 | 1,333 | 1,917 | 0 | 141 | 4,033 |
| 723 | 1,390 | 120 | 0 | 0 | 2,233 |
| 130 | 87 | 46 | 0 | 0 | 262 |
| 21,804 | 32,922 | 3,985 | 2,196 | 264 | 61,172 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.992 | 0.007 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.037 | 0.822 | 0.077 | 0.060 | 0.005 | 1.000 |
| Ongoing error | 0.067 | 0.470 | 0.432 | 0.000 | 0.031 | 1.000 |
| Expiring correct | 0.266 | 0.684 | 0.051 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.240 | 0.470 | 0.290 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.998 | 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.071 | 0.725 | 0.155 | 0.041 | 0.008 | 1.000 |
| Ongoing error | 0.000* | 0.622 | 0.354 | 0.000 | 0.024 | 1.000 |
| Expiring correct | 0.000* | 0.965 | 0.035 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.000* | 0.580 | 0.420 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.912 | 0.078 | 0.010 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.029 | 0.851 | 0.051 | 0.066 | 0.004 | 1.000 |
| Ongoing error | 0.159 | 0.330 | 0.475 | 0.000 | 0.035 | 1.000 |
| Expiring correct | 0.324 | 0.623 | 0.054 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.495 | 0.330 | 0.175 | 0.000 | 0.000 | 1.000 |

State: Idaho

|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Ongoing | Ongoing | Expiring | Expiring |  |  |
| Current-month status | participating | correct | error | correct | error | Total |

Cell counts for total households

| Nonparticipating | 433,932 | 2,244 | 257 | 0 | 0 | 436,433 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 2,593 | 14,081 | 1,072 | 1,343 | 56 | 19,145 |
| Ongoing error | 26 | 1,337 | 1,105 | 0 | 169 | 2,637 |
| Expiring correct | 0* | 1,318 | 143 | 0 | 0 | 1,461 |
| Expiring error | 72 | 120 | 45 | 0 | 0 | 236 |
| Total | 436,622 | 19,100 | 2,623 | 1,343 | 224 | 459,913 |

Cell counts for households with earnings

| Nonparticipating | 365,305 | 1,123 | 187 | 0 | 0 | 366,615 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 1,374 | 4,609 | 679 | 558 | 56 | 7,277 |
| Ongoing error | 20 | 874 | 629 | 0 | 88 | 1,611 |
| Expiring correct | 0* | 597 | 92 | 0 | 0 | 689 |
| Expiring error | 56 | 80 | 12 | 0 | 0 | 148 |
| Total | 366,755 | 7,283 | 1,600 | 558 | 143 | 376,340 |

## Total

Cell counts for households without earnings

| Nonparticipating | 68,604 | 1,121 | 70 | 0 | 69,795 |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 1,195 | 9,472 | 393 | 0 | 11,845 |  |
| Ongoing error | 4 | 463 | 785 | 1,024 |  |  |
| Expiring correct | 48 | 721 | 51 | 0 | 81 |  |
| Expiring error | 16 | 40 | 33 | 0 | 0 | 80 |
| Total | 69,867 | 11,817 | 1,023 | 0 | 88 |  |
|  |  |  |  | 785 | 81 | 83,573 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.994 | 0.005 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.135 | 0.735 | 0.056 | 0.070 | 0.003 | 1.000 |
| Ongoing error | 0.010 | 0.507 | 0.419 | 0.000 | 0.064 | 1.000 |
| Expiring correct | 0.000* | 0.902 | 0.098 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.303 | 0.507 | 0.190 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.996 | 0.003 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.189 | 0.633 | 0.093 | 0.077 | 0.008 | 1.000 |
| Ongoing error | 0.013 | 0.543 | 0.391 | 0.000 | 0.054 | 1.000 |
| Expiring correct | 0.000* | 0.866 | 0.134 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.376 | 0.543 | 0.082 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.983 | 0.016 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.101 | 0.800 | 0.033 | 0.066 | 0.000 | 1.000 |
| Ongoing error | 0.004 | 0.452 | 0.465 | 0.000 | 0.079 | 1.000 |
| Expiring correct | 0.059 | 0.879 | 0.062 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.175 | 0.452 | 0.372 | 0.000 | 0.000 | 1.000 |

State: Illinois

|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Non- | Ongoing | Ongoing | Expiring | Expiring |  |
| Current-month status | participating | correct | error | correct | error | Total |

Cell counts for total households

| Nonparticipating | 4,084,878 | 23,389 | 1,992 | 0 | 0 | 4,110,260 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 64,613 | 195,853 | 21,171 | 9,806 | 1,004 | 292,446 |
| Ongoing error | 0* | 42,435 | 26,075 | 64 | 1,413 | 69,987 |
| Expiring correct | 0* | 28,391 | 6,571 | 0 | 0 | 34,962 |
| Expiring error | 0* | 1,883 | 897 | 0 | 0 | 2,780 |
| Total | 4,149,491 | 291,951 | 56,706 | 9,870 | 2,417 | 4,510,435 |


|  |  |  | 853 | 0 | $3,595,685$ |
| :--- | :---: | ---: | ---: | ---: | ---: |
| Nonparticipating | $3,589,418$ | 5,415 | 10,949 | 69,680 |  |
| Ongoing correct | 22,138 | 34,842 | 1,239 | 32,748 |  |
| Ongoing error | $0^{*}$ | 22,496 | 10,008 | 18 | 180 |
| Expiring correct | $0^{*}$ | 7,063 | 3,475 | 0 | 0 |
| Expiring error | $0^{*}$ | 640 | 345 | 0 | 0 |
| Total | $3,611,555$ | 70,457 | 25,630 | 1,304 | 985 |
|  |  |  |  |  | $0,709,637$ |


| Nonparticipating | 494,488 | 17,975 | 1,139 | 0 | 0 | 513,602 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 43,303 | 161,010 | 10,222 | 8,566 | 493 | 223,593 |
| Ongoing error | 0* | 19,939 | 16,067 | 0 | 1,233 | 37,239 |
| Expiring correct | 0* | 21,430 | 3,096 | 0 | 0 | 24,526 |
| Expiring error | 145 | 1,141 | 553 | 0 | 0 | 1,838 |
| Total | 537,936 | 221,494 | 31,076 | 8,566 | 1,726 | 800,799 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.994 | 0.006 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.221 | 0.670 | 0.072 | 0.034 | 0.003 | 1.000 |
| Ongoing error | 0.000* | 0.606 | 0.373 | 0.001 | 0.020 | 1.000 |
| Expiring correct | 0.000* | 0.812 | 0.188 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.000* | 0.677 | 0.323 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.998 | 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.318 | 0.500 | 0.157 | 0.018 | 0.007 | 1.000 |
| Ongoing error | 0.000* | 0.687 | 0.306 | 0.002 | 0.005 | 1.000 |
| Expiring correct | 0.000* | 0.670 | 0.330 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.000* | 0.650 | 0.350 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.963 | 0.035 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.194 | 0.720 | 0.046 | 0.038 | 0.002 | 1.000 |
| Ongoing error | 0.000* | 0.535 | 0.431 | 0.000 | 0.033 | 1.000 |
| Expiring correct | 0.000* | 0.874 | 0.126 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.079 | 0.621 | 0.301 | 0.000 | 0.000 | 1.000 |

State: Indiana

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 2,140,198 | 11,123 | 1,120 | 0 | 0 | 2,152,441 |
| Ongoing correct | 16,225 | 75,537 | 3,902 | 7,906 | 499 | 104,067 |
| Ongoing error | 0* | 8,121 | 5,762 | 94 | 498 | 14,474 |
| Expiring correct | 0* | 9,409 | 847 | 0 | 0 | 10,256 |
| Expiring error | 172 | 744 | 156 | 0 | 0 | 1,071 |
| Total | 2,156,595 | 104,933 | 11,787 | 7,999 | 997 | 2,282,311 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 1,836,353 | 4,253 | 745 | 0 | 0 | 1,841,351 |
| Ongoing correct | 4,915 | 15,700 | 1,972 | 3,059 | 310 | 25,957 |
| Ongoing error | 453 | 2,550 | 1,872 | 0 | 250 | 5,125 |
| Expiring correct | 0* | 3,490 | 467 | 0 | 0 | 3,958 |
| Expiring error | 245 | 303 | 61 | 0 | 0 | 608 |
| Total | 1,841,965 | 26,296 | 5,118 | 3,059 | 561 | 1,876,999 |
|  | Cell counts for households without earnings |  |  |  |  |  |
| Nonparticipating | 303,167 | 6,870 | 375 | 0 | 0 | 310,412 |
| Ongoing correct | 11,462 | 59,836 | 1,930 | 4,847 | 188 | 78,264 |
| Ongoing error | 0* | 5,571 | 3,890 | 94 | 248 | 9,802 |
| Expiring correct | 0* | 5,968 | 380 | 0 | 0 | 6,348 |
| Expiring error | 0* | 391 | 95 | 0 | 0 | 486 |
| Total | 314,630 | 78,637 | 6,669 | 4,941 | 436 | 405,312 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.994 | 0.005 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.156 | 0.726 | 0.037 | 0.076 | 0.005 | 1.000 |
| Ongoing error | 0.000* | 0.561 | 0.398 | 0.006 | 0.034 | 1.000 |
| Expiring correct | 0.000* | 0.917 | 0.083 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.160 | 0.694 | 0.145 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.997 | 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.189 | 0.605 | 0.076 | 0.118 | 0.012 | 1.000 |
| Ongoing error | 0.088 | 0.498 | 0.365 | 0.000 | 0.049 | 1.000 |
| Expiring correct | 0.000* | 0.882 | 0.118 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.402 | 0.498 | 0.101 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.977 | 0.022 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.146 | 0.765 | 0.025 | 0.062 | 0.002 | 1.000 |
| Ongoing error | 0.000* | 0.568 | 0.397 | 0.010 | 0.025 | 1.000 |
| Expiring correct | 0.000* | 0.940 | 0.060 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.000* | 0.805 | 0.195 | 0.000 | 0.000 | 1.000 |

State: Iowa

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 1,075,320 | 5,354 | 651 | 0 | 0 | 1,081,325 |
| Ongoing correct | 9,993 | 32,447 | 3,260 | 1,196 | 52 | 46,948 |
| Ongoing error | 0 * | 7,412 | 2,356 | 0 | 170 | 9,938 |
| Expiring correct | 0 * | 1,375 | 236 | 0 | 0 | 1,611 |
| Expiring error | 0* | 262 | 77 | 0 | 0 | 338 |
| Total | 1,085,312 | 46,850 | 6,580 | 1,196 | 222 | 1,140,160 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 902,784 | 1,796 | 432 | 0 | 0 | 905,012 |
| Ongoing correct | 4,109 | 7,688 | 1,897 | 152 | 40 | 13,885 |
| Ongoing error | 0 * | 3,932 | 973 | 0 | 78 | 4,983 |
| Expiring correct | 0 * | 284 | 171 | 0 | 0 | 456 |
| Expiring error | 0* | 137 | 52 | 0 | 0 | 189 |
| Total | 906,893 | 13,837 | 3,525 | 152 | 118 | 924,526 |
|  | Cell counts for households without earnings |  |  |  |  |  |
| Nonparticipating | 172,378 | 3,558 | 219 | 0 | 0 | 176,155 |
| Ongoing correct | 6,042 | 24,759 | 1,363 | 1,043 | 13 | 33,220 |
| Ongoing error | 0 * | 3,480 | 1,383 | 0 | 91 | 4,954 |
| Expiring correct | 0 * | 1,091 | 64 | 0 | 0 | 1,155 |
| Expiring error | 0* | 125 | 25 | 0 | 0 | 150 |
| Total | 178,420 | 33,013 | 3,055 | 1,043 | 104 | 215,634 |

State: Iowa

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.994 | 0.005 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.213 | 0.691 | 0.069 | 0.025 | 0.001 | 1.000 |
| Ongoing error | 0.000* | 0.746 | 0.237 | 0.000 | 0.017 | 1.000 |
| Expiring correct | 0.000* | 0.854 | 0.146 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.000* | 0.773 | 0.227 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.998 | 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.296 | 0.554 | 0.137 | 0.011 | 0.003 | 1.000 |
| Ongoing error | 0.000* | 0.789 | 0.195 | 0.000 | 0.016 | 1.000 |
| Expiring correct | 0.000* | 0.624 | 0.376 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.000* | 0.727 | 0.273 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.979 | 0.020 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.182 | 0.745 | 0.041 | 0.031 | 0.000 | 1.000 |
| Ongoing error | 0.000* | 0.702 | 0.279 | 0.000 | 0.018 | 1.000 |
| Expiring correct | 0.000* | 0.944 | 0.056 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.000* | 0.832 | 0.168 | 0.000 | 0.000 | 1.000 |

State: Kansas

|  | Next-month status |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Current-month status | participating | Ongoing | Ongoing | Expiring | Expiring |
| correct | error | correct | error | Total |  |

Cell counts for total households

|  | 974,784 | 4,016 | 562 | 0 | 0 |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Nonparticipating | 3,927 | 35,112 | 3,691 | 80 | 44,493 |
| Ongoing correct | 517 | 3,848 | 3,025 | 7,473 |  |
| Ongoing error | 272 | 1,313 | 155 | 0 | 02 |
| Expiring correct | 30 | 90 | 55 | 0 | 0 |
| Expiring error | 979,532 | 44,379 | 7,488 | 0 | 0 |
| Total |  |  | 1,740 |  |  |

Cell counts for households with earnings

| Nonparticipating | 836,230 | 1,385 | 327 | 0 | 0 | 837,942 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 1,779 | 7,228 | 2,022 | 342 | 26 | 11,397 |
| Ongoing error | 132 | 2,276 | 911 | 0 | 27 | 3,346 |
| Expiring correct | 0* | 423 | 52 | 0 | 0 | 476 |
| Expiring error | 0* | 37 | 25 | 0 | 0 | 63 |
| Total | 838,142 | 11,350 | 3,338 | 342 | 53 | 853,224 |
|  | Cell counts for households without earnings |  |  |  |  |  |
| Nonparticipating | 138,329 | 2,631 | 236 | 0 | 0 | 141,196 |
| Ongoing correct | 2,237 | 27,884 | 1,668 | 1,341 | 54 | 33,184 |
| Ongoing error | 395 | 1,572 | 2,115 | 0 | 55 | 4,136 |
| Expiring correct | 384 | 897 | 103 | 0 | 0 | 1,383 |
| Expiring error | 46 | 46 | 29 | 0 | 0 | 120 |
| Total | 141,390 | 33,029 | 4,150 | 1,341 | 109 | 180,020 |


|  | Next-month status |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | ---: | ---: |
|  | Non- <br> participating | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error | Total |


| Nonparticipating | 0.995 | 0.004 | 0.001 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Ongoing correct | 0.088 | 0.789 | 0.083 | 0.038 | 0.002 | 1.000 |
| Ongoing error | 0.069 | 0.515 | 0.405 | 0.000 | 0.011 | 1.000 |
| Expiring correct | 0.156 | 0.754 | 0.089 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.173 | 0.515 | 0.312 | 0.000 | 0.000 | 1.000 |

Transition probabilities for households with earnings

| Nonparticipating | 0.998 |
| :--- | :--- |
| Ongoing correct | 0.156 |
| Ongoing error | 0.039 |
| Expiring correct | $0.000^{*}$ |
| Expiring error | $0.000^{*}$ |


| 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- |
| 0.634 | 0.177 | 0.030 | 0.002 | 1.000 |
| 0.680 | 0.272 | 0.000 | 0.008 | 1.000 |
| 0.890 | 0.110 | 0.000 | 0.000 | 1.000 |
| 0.595 | 0.405 | 0.000 | 0.000 | 1.000 |

Transition probabilities for households without earnings

| Nonparticipating | 0.980 | 0.019 | 0.002 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Ongoing correct | 0.067 | 0.840 | 0.050 | 0.040 | 0.002 | 1.000 |
| Ongoing error | 0.095 | 0.380 | 0.511 | 0.000 | 0.013 |  |
| Expiring correct | 0.278 | 0.648 | 0.074 | 0.000 | 0.000 |  |
| Expiring error | 0.378 | 0.380 | 0.242 | 0.000 | 0.000 |  |

State: Kentucky

|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Non- | Ongoing | Ongoing | Expiring | Expiring |  |
| Current-month status | participating | correct | error | correct | error | Total |


| Nonparticipating | $1,393,669$ | 13,871 | 1,225 | 0 | 0 | $1,408,764$ |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 11,364 | 102,933 | 5,997 | 13,346 | 1,394 | 135,034 |
| Ongoing error | $00^{*}$ | 8,806 | 6,809 | 20 | 16,146 |  |
| Expiring correct | 4,246 | 8,562 | 916 | 0 | 13,724 |  |
| Expiring error | 704 | 1,157 | 102 | 0 | 0 | 1,963 |
| Total | $1,409,982$ | 135,328 | 15,048 | 13,376 | 1,897 | $1,575,631$ |

Cell counts for households with earnings

| Nonparticipating | $1,163,425$ | 5,167 | 660 | 0 | $1,169,252$ |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 3,322 | 17,569 | 2,618 | 5,480 | 1,074 |
| Ongoing error | 430 | 3,045 | 1,915 | 29 | 5,064 |
| Expiring correct | 1,330 | 3,728 | 522 | 0 | 0 |
| Expiring error | 604 | 762 | 56 | 0 | 0 |
| Total | $1,169,111$ | 30,271 | 5,771 | 5,509 | 1,422 |
|  |  |  |  | 1,392 | $1,212,054$ |


| Nonparticipating | 229,586 | 8,703 | 565 | 0 | 0 | 238,854 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 8,249 | 85,363 | 3,379 | 7,866 | 320 | 105,177 |
| Ongoing error | $0^{*}$ | 5,761 | 4,893 | 0 | 10,839 |  |
| Expiring correct | 2,878 | 4,892 | 395 | 0 | 0,165 |  |
| Expiring error | 159 | 337 | 46 | 0 | 0 | 542 |
|  | 240,871 | 105,057 | 9,278 | 7,866 | 505 | 363,577 |


|  | Next-month status |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error | Total |

State: Louisiana

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 1,431,494 | 18,280 | 2,332 | 0 | 0 | 1,452,106 |
| Ongoing correct | 8,924 | 111,245 | 6,695 | 22,280 | 1,940 | 151,084 |
| Ongoing error | 2,420 | 5,108 | 8,591 | 0 | 2,942 | 19,060 |
| Expiring correct | 6,451 | 15,135 | 1,235 | 0 | 0 | 22,821 |
| Expiring error | 3,374 | 1,338 | 282 | 0 | 0 | 4,994 |
| Total | 1,452,663 | 151,106 | 19,135 | 22,280 | 4,882 | 1,650,065 |


| Nonparticipating | $1,240,375$ | 6,893 | 1,281 | 0 | 0 | $1,248,549$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 1,043 | 26,289 | 3,741 | 9,774 | 1,618 | 42,464 |
| Ongoing error | 1,714 | 2,100 | 2,941 | 0 | 1,979 | 0 |
| Expiring correct | 2,565 | 6,738 | 705 | 0 | 10,008 |  |
| Expiring error | 2,633 | 879 | 144 | 0 | 0 | 3,657 |
| Total | $1,248,330$ | 42,900 | 8,812 | 9,774 | 3,597 | $1,313,412$ |


| Nonparticipating | 190,631 | 11,387 | 1,051 | 0 | 203,069 |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 8,282 | 84,957 | 2,954 | 12,506 | 322 | 109,021 |
| Ongoing error | 738 | 3,007 | 5,650 | 0 | 10,358 |  |
| Expiring correct | 3,867 | 8,466 | 531 | 0 | 12,864 |  |
| Expiring error | 814 | 389 | 137 | 0 | 0 | 0 |
| Total | 204,332 | 108,206 | 10,323 | 12,506 | 1,285 | 336,653 |

State: Louisiana

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.986 | 0.013 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.059 | 0.736 | 0.044 | 0.147 | 0.013 | 1.000 |
| Ongoing error | 0.127 | 0.268 | 0.451 | 0.000 | 0.154 | 1.000 |
| Expiring correct | 0.283 | 0.663 | 0.054 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.676 | 0.268 | 0.056 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.993 | 0.006 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.025 | 0.619 | 0.088 | 0.230 | 0.038 | 1.000 |
| Ongoing error | 0.196 | 0.240 | 0.337 | 0.000 | 0.227 | 1.000 |
| Expiring correct | 0.256 | 0.673 | 0.070 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.720 | 0.240 | 0.039 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.939 | 0.056 | 0.005 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.076 | 0.779 | 0.027 | 0.115 | 0.003 | 1.000 |
| Ongoing error | 0.071 | 0.290 | 0.545 | 0.000 | 0.093 | 1.000 |
| Expiring correct | 0.301 | 0.658 | 0.041 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.607 | 0.290 | 0.103 | 0.000 | 0.000 | 1.000 |

State: Maine

|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Non- | Ongoing | Ongoing | Expiring | Expiring |  |
| Current-month status | participating | correct | error | correct | error | Total |

Cell counts for total households

|  | 463,181 | 2,553 | 502 | 0 | 0 | 466,237 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Nonparticipating | 1,939 | 29,003 | 2,063 | 6,380 | 425 | 6,810 |
| Ongoing correct | 1,393 | 1,774 | 2,288 | 16 | 608 |  |
| Ongoing error | $0 *$ | 6,075 | 960 | 0 | 0 | 7,035 |
| Expiring correct | 462 | 304 | 266 | 0 | 0 | 1,032 |
| Expiring error | 466,976 | 39,709 | 6,079 | 6,396 | 1,033 | 520,193 |
|  |  |  |  |  |  |  |

Cell counts for households with earnings

|  | 391,687 | 811 | 228 | 0 | 0 | 392,727 |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Nonparticipating | 436 | 2,879 | 648 | 2,182 | 0,404 |  |
| Ongoing correct | 801 | 515 | 281 | 0 | 163 |  |
| Ongoing error | $0 *$ | 2,042 | 532 | 0 | 0 | 2,574 |
| Expiring correct | 236 | 126 | 68 | 0 | 0 | 429 |
| Expiring error | 393,160 | 6,373 | 1,756 | 2,182 | 423 | 403,895 |

Total

| Nonparticipating | 71,582 | 1,742 | 274 | 0 | 0 | 73,598 |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 1,426 | 26,124 | 1,415 | 4,198 | 166 | 43,328 |
| Ongoing error | 582 | 1,259 | 2,007 | 16 | 4,309 |  |
| Expiring correct | $0 *$ | 4,033 | 429 | 0 | 0 | 4,462 |
| Expiring error | 226 | 178 | 198 | 0 | 002 |  |
| Total | 73,815 | 33,336 | 4,323 | 4,214 | 610 | 116,299 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.993 | 0.005 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.049 | 0.729 | 0.052 | 0.160 | 0.011 | 1.000 |
| Ongoing error | 0.229 | 0.292 | 0.376 | 0.003 | 0.100 | 1.000 |
| Expiring correct | 0.000* | 0.863 | 0.137 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.448 | 0.295 | 0.258 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.997 | 0.002 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.068 | 0.450 | 0.101 | 0.341 | 0.040 | 1.000 |
| Ongoing error | 0.455 | 0.293 | 0.160 | 0.000 | 0.093 | 1.000 |
| Expiring correct | 0.000* | 0.793 | 0.207 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.550 | 0.293 | 0.157 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.973 | 0.024 | 0.004 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.043 | 0.784 | 0.042 | 0.126 | 0.005 | 1.000 |
| Ongoing error | 0.135 | 0.292 | 0.466 | 0.004 | 0.103 | 1.000 |
| Expiring correct | 0.000* | 0.904 | 0.096 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.375 | 0.296 | 0.330 | 0.000 | 0.000 | 1.000 |

State: Maryland

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 1,836,252 | 7,772 | 1,101 | 0 | 0 | 1,845,125 |
| Ongoing correct | 1,468 | 64,515 | 4,500 | 10,795 | 1,202 | 82,481 |
| Ongoing error | 3,536 | 1,721 | 9,293 | 0 | 2,052 | 16,602 |
| Expiring correct | 2,957 | 7,248 | 689 | 0 | 0 | 10,894 |
| Expiring error | 2,272 | 341 | 679 | 0 | 0 | 3,293 |
| Total | 1,846,486 | 81,597 | 16,263 | 10,795 | 3,255 | 1,958,395 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 1,622,944 | 1,762 | 227 | 0 | 0 | 1,624,933 |
| Ongoing correct | 0* | 9,304 | 1,943 | 2,494 | 693 | 14,434 |
| Ongoing error | 1,672 | 593 | 3,084 | 0 | 782 | 6,132 |
| Expiring correct | 0* | 2,416 | 447 | 0 | 0 | 2,864 |
| Expiring error | 1,051 | 144 | 296 | 0 | 0 | 1,491 |
| Total | 1,625,667 | 14,221 | 5,997 | 2,494 | 1,475 | 1,649,854 |
|  | Cell counts for households without earnings |  |  |  |  |  |
| Nonparticipating | 212,389 | 6,009 | 875 | 0 | 0 | 219,273 |
| Ongoing correct | 1,909 | 55,211 | 2,558 | 8,301 | 509 | 68,488 |
| Ongoing error | 1,909 | 1,127 | 6,208 | 0 | 1,270 | 10,515 |
| Expiring correct | 3,381 | 4,835 | 241 | 0 | 0 | 8,457 |
| Expiring error | 1,230 | 194 | 384 | 0 | 0 | 1,808 |
| Total | 220,819 | 67,376 | 10,266 | 8,301 | 1,779 | 308,541 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.995 | 0.004 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.018 | 0.782 | 0.055 | 0.131 | 0.015 | 1.000 |
| Ongoing error | 0.213 | 0.104 | 0.560 | 0.000 | 0.124 | 1.000 |
| Expiring correct | 0.271 | 0.665 | 0.063 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.690 | 0.104 | 0.206 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.999 | 0.001 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.645 | 0.135 | 0.173 | 0.048 | 1.000 |
| Ongoing error | 0.273 | 0.097 | 0.503 | 0.000 | 0.128 | 1.000 |
| Expiring correct | 0.000* | 0.844 | 0.156 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.705 | 0.097 | 0.198 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.969 | 0.027 | 0.004 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.028 | 0.806 | 0.037 | 0.121 | 0.007 | 1.000 |
| Ongoing error | 0.182 | 0.107 | 0.590 | 0.000 | 0.121 | 1.000 |
| Expiring correct | 0.400 | 0.572 | 0.029 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.681 | 0.107 | 0.212 | 0.000 | 0.000 | 1.000 |

State: Massachusetts

|  | Next-month status |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Current-month status | Non- | Ongoing | Ongoing | Expiring | Expiring |
| participating | correct | error | correct | error | Total |

Cell counts for total households

|  | 2,49 | 10,146 | 1,311 | 0 | 0 | $2,270,578$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Nonparticipating | $2,259,121$ | 68,845 | 4,538 | 12,826 | 1,301 | $1,09,806$ |
| Ongoing correct | 2,297 | 931 | 4,694 | 55 | 12,193 |  |
| Ongoing error | 5,422 | 8,963 | 1,258 | 0 | 13,173 |  |
| Expiring correct | 2,953 | 1,922 | 200 | 349 | 0 | 0 |
| Expiring error | $2,271,713$ | 89,085 | 12,150 | 12,881 | 2,470 |  |
|  |  |  |  | 2,392 | $2,388,220$ |  |

Cell counts for households with earnings

| Nonparticipating | 1,880,260 | 2,269 | 584 | 0 | 0 | 1,883,114 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 646 | 5,793 | 1,572 | 2,569 | 751 | 11,330 |
| Ongoing error | 2,090 | 304 | 574 | 28 | 278 | 3,274 |
| Expiring correct | 0* | 2,733 | 419 | 0 | 0 | 3,152 |
| Expiring error | 812 | 106 | 130 | 0 | 0 | 1,048 |
| Total | 1,883,808 | 11,206 | 3,279 | 2,597 | 1,029 | 1,901,919 |
|  | Cell counts for households without earnings |  |  |  |  |  |
| Nonparticipating | 378,570 | 7,876 | 727 | 0 | 0 | 387,174 |
| Ongoing correct | 1,472 | 63,052 | 2,966 | 10,258 | 550 | 78,297 |
| Ongoing error | 3,313 | 627 | 4,120 | 27 | 813 | 8,900 |
| Expiring correct | 3,455 | 6,220 | 838 | 0 | 0 | 10,513 |
| Expiring error | 1,095 | 104 | 219 | 0 | 0 | 1,418 |
| Total | 387,905 | 77,879 | 8,870 | 10,284 | 1,363 | 486,302 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.995 | 0.004 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.026 | 0.767 | 0.051 | 0.143 | 0.014 | 1.000 |
| Ongoing error | 0.445 | 0.076 | 0.385 | 0.005 | 0.089 | 1.000 |
| Expiring correct | 0.224 | 0.680 | 0.095 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.778 | 0.081 | 0.141 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.998 | 0.001 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.057 | 0.511 | 0.139 | 0.227 | 0.066 | 1.000 |
| Ongoing error | 0.638 | 0.093 | 0.175 | 0.009 | 0.085 | 1.000 |
| Expiring correct | 0.000* | 0.867 | 0.133 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.775 | 0.101 | 0.124 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.978 | 0.020 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.019 | 0.805 | 0.038 | 0.131 | 0.007 | 1.000 |
| Ongoing error | 0.372 | 0.070 | 0.463 | 0.003 | 0.091 | 1.000 |
| Expiring correct | 0.329 | 0.592 | 0.080 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.772 | 0.073 | 0.154 | 0.000 | 0.000 | 1.000 |

State: Michigan

|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Non- | Ongoing | Ongoing | Expiring | Expiring |  |
| Current-month status | participating | correct | error | correct | error | Total |

Cell counts for total households

| Nonparticipating | $3,477,093$ | 18,980 | 3,143 | 0 | $0,499,216$ |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 5,715 | 173,890 | 16,339 | 215,069 |  |
| Ongoing error | 10,601 | 8,174 | 36,670 | 59,341 |  |
| Expiring correct | 4,001 | 11,778 | 1,484 | 0 | 1,841 |
| Expiring error | 3,674 | 793 | 1,288 | 0 | 0,895 |
| Total | $3,501,085$ | 213,614 | 58,925 | 0 | 0 |
|  |  |  | 17,283 | 5,263 |  |

Cell counts for households with earnings

| Nonparticipating | $2,901,436$ | 6,322 | 1,714 | 0 | 0 | $2,909,471$ |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | $0^{*}$ | 37,519 | 9,029 | 8,112 | 1,444 | 26,103 |
| Ongoing error | 6,456 | 3,778 | 14,021 | 0 | 2,266 | 0 |
| Expiring correct | 1,296 | 5,712 | 1,019 | 0 | 0,026 |  |
| Expiring error | 2,546 | 528 | 630 | 0 | 0 | 0,704 |
| Total | $2,911,734$ | 53,858 | 26,413 | 8,112 | 3,711 | $3,003,827$ |


| Nonparticipating | 574,654 | 12,658 | 1,429 | 0 | 588,742 |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 6,831 | 136,371 | 7,311 | 9,172 | 397 |  |
| Ongoing error | 4,061 | 4,395 | 22,649 | 0 | 1,629 | 32,734 |
| Expiring correct | 2,692 | 6,056 | 465 | 0 | 9,213 |  |
| Expiring error | 1,114 | 275 | 658 | 0 | 0 | 2,047 |
| Total | 589,351 | 159,755 | 32,513 | 9,172 | 2,026 | 792,817 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.994 | 0.005 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.027 | 0.809 | 0.076 | 0.080 | 0.009 | 1.000 |
| Ongoing error | 0.179 | 0.138 | 0.618 | 0.000 | 0.066 | 1.000 |
| Expiring correct | 0.232 | 0.682 | 0.086 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.638 | 0.138 | 0.224 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.997 | 0.002 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.669 | 0.161 | 0.145 | 0.026 | 1.000 |
| Ongoing error | 0.243 | 0.142 | 0.529 | 0.000 | 0.085 | 1.000 |
| Expiring correct | 0.161 | 0.712 | 0.127 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.687 | 0.142 | 0.170 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.976 | 0.022 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.043 | 0.852 | 0.046 | 0.057 | 0.002 | 1.000 |
| Ongoing error | 0.124 | 0.134 | 0.692 | 0.000 | 0.050 | 1.000 |
| Expiring correct | 0.292 | 0.657 | 0.050 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.544 | 0.134 | 0.321 | 0.000 | 0.000 | 1.000 |

State: Minnesota

|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Ongoing | Ongoing | Expiring | Expiring |  |  |
| Current-month status | participating | Ongoing | correct | error | correct | error |

Cell counts for total households

| Nonparticipating | 1,796,230 | 5,551 | 218 | 0 | 0 | 1,802,000 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 26,305 | 56,874 | 3,113 | 544 | 0 | 86,836 |
| Ongoing error | 0* | 20,191 | 3,314 | 0 | 51 | 23,556 |
| Expiring correct | 0 * | 3,718 | 96 | 0 | 0 | 3,814 |
| Expiring error | 0* | 178 | 94 | 0 | 0 | 271 |
| Total | 1,822,536 | 86,511 | 6,836 | 544 | 51 | 1,916,478 |

Cell counts for households with earnings

|  | 1,58 | 120 | 0 | 0 | $1,583,353$ |  |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Nonparticipating | $1,581,606$ | 1,626 | 1,532 | 0 | 17,774 |  |
| Ongoing correct | 8,033 | 8,089 | 120 | 0,267 |  |  |
| Ongoing error | $0^{*}$ | 7,259 | 74 | 0 | 0 | 0 |
| Expiring correct | $0^{*}$ | 782 | 0 | 0 | 0 | 0 |
| Expiring error | $0^{*}$ | 13 | 0 | 0 | 13 |  |
| Total | $1,589,639$ | 17,768 | 2,734 | 120 | $1,610,262$ |  |

## Total

Nonparticipating
Ongoing correct
Ongoing error
Expiring correct
Expiring error
Total

| 214,593 | 3,925 | 98 | 0 | 0 | 218,616 |
| ---: | ---: | ---: | ---: | ---: | ---: |
| 18,303 | 48,785 | 1,581 | 423 | 0 | 69,093 |
| $0^{*}$ | 12,932 | 2,307 | 0 | 51 | 15,289 |
| $0^{*}$ | 2,927 | 22 | 0 | 0 | 2,949 |
| $0^{*}$ | 175 | 94 | 0 | 0 | 268 |
| 232,896 | 68,743 | 4,102 | 423 | 51 | 306,216 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.997 | 0.003 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.303 | 0.655 | 0.036 | 0.006 | 0.000 | 1.000 |
| Ongoing error | 0.000* | 0.857 | 0.141 | 0.000 | 0.002 | 1.000 |
| Expiring correct | 0.000* | 0.975 | 0.025 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.000* | 0.655 | 0.345 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.999 | 0.001 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.452 | 0.455 | 0.086 | 0.007 | 0.000 | 1.000 |
| Ongoing error | 0.000* | 0.878 | 0.122 | 0.000 | 0.000 | 1.000 |
| Expiring correct | 0.000* | 0.914 | 0.086 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.000* | 1.000 | 0.000 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.982 | 0.018 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.265 | 0.706 | 0.023 | 0.006 | 0.000 | 1.000 |
| Ongoing error | 0.000* | 0.846 | 0.151 | 0.000 | 0.003 | 1.000 |
| Expiring correct | 0.000* | 0.992 | 0.008 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.000* | 0.651 | 0.349 | 0.000 | 0.000 | 1.000 |

State: Mississippi


| Nonparticipating | 769,061 | 2,249 | 182 | 0 | 0 | 771,491 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 12,555 | 12,642 | 3,141 | 214 | 27 | 28,579 |
| Ongoing error | 0* | 12,167 | 1,340 | 0 | 0 | 13,507 |
| Expiring correct | 0* | 1,412 | 348 | 0 | 0 | 1,760 |
| Expiring error | 0* | 80 | 32 | 0 | 0 | 112 |
| Total | 781,616 | 28,549 | 5,043 | 214 | 27 | 815,449 |


| Nonparticipating | 110,675 | 4,561 | 286 | 0 | 0 | 115,521 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 10,438 | 60,577 | 2,197 | 3,351 | 77 | 0,639 |
| Ongoing error | $0^{*}$ | 5,594 | 1,873 | 0 | 0,529 |  |
| Expiring correct | $0^{*}$ | 4,676 | 142 | 0 | 0,817 |  |
| Expiring error | $0^{*}$ | 180 | 52 | 0 | 0 | 233 |
| Total | 121,113 | 75,587 | 4,550 | 3,351 | 138 | 204,739 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.992 | 0.008 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.215 | 0.699 | 0.051 | 0.034 | 0.001 | 1.000 |
| Ongoing error | 0.000* | 0.844 | 0.153 | 0.000 | 0.003 | 1.000 |
| Expiring correct | 0.000* | 0.925 | 0.075 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.000* | 0.797 | 0.203 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.997 | 0.003 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.439 | 0.442 | 0.110 | 0.007 | 0.001 | 1.000 |
| Ongoing error | 0.000* | 0.901 | 0.099 | 0.000 | 0.000 | 1.000 |
| Expiring correct | 0.000* | 0.802 | 0.198 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.000* | 0.712 | 0.288 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.958 | 0.039 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.136 | 0.790 | 0.029 | 0.044 | 0.001 | 1.000 |
| Ongoing error | 0.000* | 0.743 | 0.249 | 0.000 | 0.008 | 1.000 |
| Expiring correct | 0.000* | 0.971 | 0.029 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.000* | 0.775 | 0.225 | 0.000 | 0.000 | 1.000 |

State: Missouri

|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Non- | Ongoing | Ongoing | Expiring | Expiring |  |
| Current-month status | participating | correct | error | correct | error | Total |

Cell counts for total households

| Nonparticipating | $2,031,166$ | 13,701 | 1,156 | 0 | 0 | $2,046,024$ |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 16,322 | 92,185 | 5,815 | 21,257 | 2,214 | 137,793 |
| Ongoing error | $0^{*}$ | 8,857 | 6,856 | 68 | 2,231 | 0 |
| Expiring correct | $0^{*}$ | 20,788 | 2,471 | 0 | 23,258 |  |
| Expiring error | 1,472 | 2,297 | 653 | 0 | 0 | 4,422 |
| Total | $2,048,960$ | 137,827 | 16,951 | 21,324 | 4,445 | $2,229,508$ |

## Cell counts for households with earnings

| Nonparticipating | 1,742,642 | 5,163 | 654 | 0 | 0 | 1,748,459 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 5,139 | 14,937 | 2,208 | 7,510 | 1,385 | 31,179 |
| Ongoing error | 1,666 | 2,269 | 1,736 | 68 | 956 | 6,695 |
| Expiring correct | 0* | 8,217 | 1,659 | 0 | 0 | 9,876 |
| Expiring error | 1,164 | 810 | 348 | 0 | 0 | 2,322 |
| Total | 1,750,612 | 31,397 | 6,605 | 7,578 | 2,341 | 1,798,532 |
|  | Cell counts for households without earnings |  |  |  |  |  |
| Nonparticipating | 286,071 | 8,538 | 502 | 0 | 0 | 295,112 |
| Ongoing correct | 11,505 | 77,248 | 3,607 | 13,746 | 830 | 106,936 |
| Ongoing error | 0* | 6,587 | 5,120 | 0 | 1,275 | 12,982 |
| Expiring correct | 290 | 12,739 | 811 | 0 | 0 | 13,840 |
| Expiring error | 481 | 1,318 | 305 | 0 | 0 | 2,105 |
| Total | 298,348 | 106,430 | 10,346 | 13,746 | 2,104 | 430,976 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.993 | 0.007 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.118 | 0.669 | 0.042 | 0.154 | 0.016 | 1.000 |
| Ongoing error | 0.000* | 0.492 | 0.381 | 0.004 | 0.124 | 1.000 |
| Expiring correct | 0.000* | 0.894 | 0.106 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.333 | 0.519 | 0.148 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.997 | 0.003 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.165 | 0.479 | 0.071 | 0.241 | 0.044 | 1.000 |
| Ongoing error | 0.249 | 0.339 | 0.259 | 0.010 | 0.143 | 1.000 |
| Expiring correct | 0.000* | 0.832 | 0.168 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.501 | 0.349 | 0.150 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.969 | 0.029 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.108 | 0.722 | 0.034 | 0.129 | 0.008 | 1.000 |
| Ongoing error | 0.000* | 0.507 | 0.394 | 0.000 | 0.098 | 1.000 |
| Expiring correct | 0.021 | 0.920 | 0.059 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.229 | 0.626 | 0.145 | 0.000 | 0.000 | 1.000 |

State: Montana

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 331,098 | 1,818 | 172 | 0 | 0 | 333,089 |
| Ongoing correct | 5,626 | 13,436 | 1,596 | 701 | 28 | 21,387 |
| Ongoing error | 0* | 5,136 | 1,478 | 0 | 58 | 6,673 |
| Expiring correct | 0* | 905 | 98 | 0 | 0 | 1,003 |
| Expiring error | 0* | 132 | 13 | 0 | 0 | 146 |
| Total | 336,724 | 21,428 | 3,358 | 701 | 86 | 362,298 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 280,979 | 459 | 159 | 0 | 0 | 281,597 |
| Ongoing correct | 1,596 | 3,561 | 1,075 | 127 | 13 | 6,372 |
| Ongoing error | 0* | 2,131 | 866 | 0 | 15 | 3,011 |
| Expiring correct | 0* | 194 | 70 | 0 | 0 | 264 |
| Expiring error | 1 | 30 | 0 | 0 | 0 | 31 |
| Total | 282,575 | 6,374 | 2,169 | 127 | 28 | 291,274 |
|  | Cell counts for households without earnings |  |  |  |  |  |
| Nonparticipating | 50,038 | 1,359 | 14 | 0 | 0 | 51,411 |
| Ongoing correct | 4,111 | 9,875 | 521 | 574 | 15 | 15,097 |
| Ongoing error | 0* | 3,006 | 613 | 0 | 43 | 3,661 |
| Expiring correct | 0* | 673 | 28 | 0 | 0 | 701 |
| Expiring error | 0* | 141 | 13 | 0 | 0 | 154 |
| Total | 54,149 | 15,054 | 1,189 | 574 | 58 | 71,024 |



Transition probabilities for total households

| Nonparticipating | 0.994 | 0.005 | 0.001 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Ongoing correct | 0.263 | 0.628 | 0.075 | 0.033 | 0.001 | 1.000 |
| Ongoing error | $0.000^{*}$ | 0.770 | 0.222 | 0.000 | 0.009 |  |
| Expiring correct | $0.000^{*}$ | 0.902 | 0.098 | 0.000 | 0.000 | 1.000 |
| Expiring error | $0.000^{*}$ | 0.908 | 0.092 | 0.000 | 0.000 | 1.000 |

Expiring error
0.000*
$0.908 \quad 0.092$
0.000
1.000

Transition probabilities for households with earnings

| 0.002 | 0.001 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- |
| 0.559 | 0.169 | 0.020 | 0.002 | 1.000 |
| 0.708 | 0.287 | 0.000 | 0.005 | 1.000 |
| 0.736 | 0.264 | 0.000 | 0.000 | 1.000 |
| 0.981 | 0.000 | 0.000 | 0.000 | 1.000 |

Transition probabilities for households without earnings

| Nonparticipating | 0.973 | 0.026 | 0.000 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Ongoing correct | 0.272 | 0.654 | 0.035 | 0.038 | 0.001 | 1.000 |
| Ongoing error | $0.000^{*}$ | 0.821 | 0.167 | 0.000 | 0.012 | 1.000 |
| Expiring correct | $0.000^{*}$ | 0.960 | 0.040 | 0.000 | 0.000 |  |
| Expiring error | $0.000^{*}$ | 0.913 | 0.087 | 0.000 | 0.000 | 1.000 |

State: Nebraska

|  | Next-month status |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Current-month status | Non- | Ongoing | Ongoing | Expiring | Expiring |
| participating | correct | error | correct | error | Total |

Cell counts for total households

| Nonparticipating | 620,354 | 1,888 | 344 | 0 | 0 | 622,586 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 4,718 | 18,465 | 2,480 | 1,932 | 485 | 28,081 |
| Ongoing error | 0* | 3,970 | 2,476 | 0 | 225 | 6,672 |
| Expiring correct | 0* | 3,206 | 488 | 0 | 0 | 3,694 |
| Expiring error | 34 | 473 | 211 | 0 | 0 | 719 |
| Total | 625,106 | 28,004 | 5,999 | 1,932 | 710 | 661,752 |


| Nonparticipating | 519,955 | 670 | 135 | 0 | 0 | 520,759 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 1,263 | 4,067 | 1,386 | 594 | 316 | 7,627 |
| Ongoing error | 314 | 1,583 | 1,425 | 0 | 91 | 3,413 |
| Expiring correct | 0* | 1,089 | 320 | 0 | 0 | 1,409 |
| Expiring error | 107 | 192 | 115 | 0 | 0 | 415 |
| Total | 521,639 | 7,602 | 3,381 | 594 | 407 | 533,623 |

Cell counts for households without earnings

| Nonparticipating | 100,096 | 1,219 | 210 | 0 | 0 | 101,524 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 3,372 | 14,398 | 1,094 | 1,338 | 0,370 |  |
| Ongoing error | $0^{*}$ | 2,387 | 1,051 | 0 | 135 | 0,573 |
| Expiring correct | $0^{*}$ | 2,120 | 168 | 0 | 0,288 |  |
| Expiring error | $0^{*}$ | 278 | 96 | 0 | 0 | 374 |
| Total | 103,468 | 20,402 | 2,618 | 1,338 | 303 | 128,129 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.996 | 0.003 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.168 | 0.658 | 0.088 | 0.069 | 0.017 | 1.000 |
| Ongoing error | 0.000* | 0.595 | 0.371 | 0.000 | 0.034 | 1.000 |
| Expiring correct | 0.000* | 0.868 | 0.132 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.047 | 0.659 | 0.294 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.998 | 0.001 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.166 | 0.533 | 0.182 | 0.078 | 0.041 | 1.000 |
| Ongoing error | 0.092 | 0.464 | 0.418 | 0.000 | 0.027 | 1.000 |
| Expiring correct | 0.000* | 0.773 | 0.227 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.258 | 0.464 | 0.278 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.986 | 0.012 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.166 | 0.707 | 0.054 | 0.066 | 0.008 | 1.000 |
| Ongoing error | 0.000* | 0.668 | 0.294 | 0.000 | 0.038 | 1.000 |
| Expiring correct | 0.000* | 0.927 | 0.073 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.000* | 0.743 | 0.257 | 0.000 | 0.000 | 1.000 |

State: Nevada

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 690,533 | 3,657 | 350 | 0 | 0 | 694,540 |
| Ongoing correct | 5,222 | 16,698 | 1,930 | 1,681 | 106 | 25,636 |
| Ongoing error | 1,035 | 1,673 | 358 | 13 | 48 | 3,127 |
| Expiring correct | 0* | 3,323 | 411 | 0 | 0 | 3,734 |
| Expiring error | 63 | 90 | 14 | 0 | 0 | 167 |
| Total | 696,853 | 25,441 | 3,062 | 1,694 | 154 | 727,205 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 592,261 | 992 | 261 | 0 | 0 | 593,515 |
| Ongoing correct | 656 | 2,275 | 909 | 494 | 47 | 4,381 |
| Ongoing error | 1,022 | 354 | 89 | 0 | 33 | 1,499 |
| Expiring correct | 0* | 705 | 199 | 0 | 0 | 904 |
| Expiring error | 50 | 20 | 14 | 0 | 0 | 85 |
| Total | 593,990 | 4,346 | 1,473 | 494 | 80 | 600,384 |
|  | Cell counts for households without earnings |  |  |  |  |  |
| Nonparticipating | 98,222 | 2,665 | 88 | 0 | 0 | 100,975 |
| Ongoing correct | 4,610 | 14,423 | 1,020 | 1,186 | 59 | 21,299 |
| Ongoing error | 17 | 1,319 | 269 | 13 | 15 | 1,632 |
| Expiring correct | 0* | 2,621 | 212 | 0 | 0 | 2,833 |
| Expiring error | 15 | 67 | 0 | 0 | 0 | 82 |
| Total | 102,864 | 21,095 | 1,589 | 1,199 | 73 | 126,821 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.994 | 0.005 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.204 | 0.651 | 0.075 | 0.066 | 0.004 | 1.000 |
| Ongoing error | 0.331 | 0.535 | 0.114 | 0.004 | 0.015 | 1.000 |
| Expiring correct | 0.000* | 0.890 | 0.110 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.375 | 0.539 | 0.086 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.998 | 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.150 | 0.519 | 0.208 | 0.113 | 0.011 | 1.000 |
| Ongoing error | 0.682 | 0.236 | 0.060 | 0.000 | 0.022 | 1.000 |
| Expiring correct | 0.000* | 0.780 | 0.220 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.595 | 0.236 | 0.169 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.973 | 0.026 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.216 | 0.677 | 0.048 | 0.056 | 0.003 | 1.000 |
| Ongoing error | 0.010 | 0.808 | 0.165 | 0.008 | 0.009 | 1.000 |
| Expiring correct | 0.000* | 0.925 | 0.075 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.184 | 0.816 | 0.000 | 0.000 | 0.000 | 1.000 |

State: New Hampshire

|  | Next-month status |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Current-month status | participating | Ongoing | Ongoing | Expiring | Expiring |
| correct | error | correct | error | Total |  |


| Nonparticipating | 443,903 | 1,523 | 241 | 0 | 0 | 445,666 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 0* | 9,774 | 844 | 2,341 | 294 | 13,253 |
| Ongoing error | 667 | 256 | 902 | 13 | 380 | 2,218 |
| Expiring correct | 1,208 | 998 | 153 | 0 | 0 | 2,359 |
| Expiring error | 531 | 82 | 65 | 0 | 0 | 679 |
| Total | 446,310 | 12,633 | 2,205 | 2,354 | 674 | 464,177 |


|  |  | 346 | 102 | 0 | 0 | 397,272 |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Nonparticipating | 396,823 | $0^{*}$ | 1,436 | 433 | 2,388 |  |
| Ongoing correct | 208 | 129 | 204 | 163 | 167 |  |
| Ongoing error | 202 | 196 | 62 | 0 | 0 | 461 |
| Expiring correct | 255 | 66 | 13 | 0 | 0 | 334 |
| Expiring error | 397,489 | 2,173 | 715 | 469 | 331 | 401,176 |


| Nonparticipating | 47,013 | 1,176 | 138 | 0 | 0 | 48,327 |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | $0 *$ | 8,338 | 511 | 10,865 |  |  |
| Ongoing error | 485 | 127 | 698 | 1,885 | 0 | 213 |
| Expiring correct | 1,054 | 790 | 91 | 0 | 0 | 1,923 |
| Expiring error | 269 | 29 | 52 | 0 | 0 | 351 |
| Total | 48,821 | 10,461 | 1,490 | 1,885 | 343 | 63,001 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.996 | 0.003 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.737 | 0.064 | 0.177 | 0.022 | 1.000 |
| Ongoing error | 0.301 | 0.115 | 0.407 | 0.006 | 0.171 | 1.000 |
| Expiring correct | 0.512 | 0.423 | 0.065 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.782 | 0.121 | 0.096 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.999 | 0.001 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.601 | 0.139 | 0.191 | 0.068 | 1.000 |
| Ongoing error | 0.289 | 0.178 | 0.282 | 0.019 | 0.232 | 1.000 |
| Expiring correct | 0.439 | 0.425 | 0.136 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.763 | 0.197 | 0.040 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.973 | 0.024 | 0.003 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.767 | 0.047 | 0.174 | 0.012 | 1.000 |
| Ongoing error | 0.318 | 0.084 | 0.459 | 0.000 | 0.140 | 1.000 |
| Expiring correct | 0.545 | 0.408 | 0.047 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.768 | 0.084 | 0.148 | 0.000 | 0.000 | 1.000 |

State: New Jersey

|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Non- | Ongoing | Ongoing | Expiring | Expiring |  |
| Current-month status | participating | correct | error | correct | error | Total |


| Nonparticipating | $2,833,427$ | 5,010 | 452 | 0 | $2,838,889$ |  |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | $0^{*}$ | 104,975 | 5,829 | 16,084 | 1,190 | 128,078 |
| Ongoing error | 4,677 | 633 | 13,919 | 0 | 2,862 | 0 |
| Expiring correct | 4,635 | 10,302 | 1,061 | 0 | 15,997 |  |
| Expiring error | 3,408 | 116 | 526 | 0 | 0 | 4,051 |
| Total | $2,846,147$ | 121,036 | 21,788 | 4,052 | $3,009,107$ |  |

Cell counts for households with earnings

| Nonparticipating | $2,397,878$ | 677 | 67 | 0 | 0 | $2,398,622$ |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | $0 *$ | 12,368 | 1,945 | 2,598 | 530 | 0,441 |
| Ongoing error | 1,535 | 140 | 3,970 | 0 | 0,601 |  |
| Expiring correct | 291 | 1,863 | 406 | 0 | 0 |  |
| Expiring error | 1,338 | 31 | 117 | 0 | 0 | 1,486 |
| Total | $2,401,042$ | 15,079 | 6,505 | 2,598 | 1,485 | $2,426,710$ |


| Nonparticipating | 435,662 | 4,333 | 386 | 0 | 0 | 440,381 |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | $0^{*}$ | 92,607 | 3,884 | 13,486 | 660 | 110,637 |
| Ongoing error | 3,082 | 494 | 9,948 | 0 | 1,907 | 0,431 |
| Expiring correct | 4,295 | 8,441 | 654 | 0 | 13,391 |  |
| Expiring error | 2,065 | 82 | 410 | 0 | 0 |  |
| Total | 445,105 | 105,956 | 15,282 | 13,486 | 2,567 | 582,397 |


|  | Next-month status |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Current-month status | Non- | Ongoing | Ongoing | Expiring | Expiring |
| carticipating | correct | error | correct | error | Total |

Transition probabilities for total households

| Nonparticipating | 0.998 | 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Ongoing correct | $0.000^{*}$ | 0.820 | 0.046 | 0.126 | 0.009 | 1.000 |
| Ongoing error | 0.212 | 0.029 | 0.630 | 0.000 | 0.130 |  |
| Expiring correct | 0.290 | 0.644 | 0.066 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.841 | 0.029 | 0.130 | 0.000 | 0.000 | 1.000 |

Transition probabilities for households with earnings
Nonparticipating
Ongoing correct
Ongoing error
Expiring correct
Expiring error

| 1.000 | 0.000 |
| :--- | :--- |
| $0.000^{*}$ | 0.709 |
| 0.233 | 0.021 |
| 0.114 | 0.728 |

0.000
0.112
0.602
0.159
0.079
0.000
0.149
0.000
0.000
0.000

| 0.000 | 1.000 |
| :--- | :--- |
| 0.030 | 1.000 |
| 0.145 | 1.000 |
| 0.000 | 1.000 |
| 0.000 | 1.000 |

Transition probabilities for households without earnings

| Nonparticipating | 0.989 | 0.010 | 0.001 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Ongoing correct | $0.000^{*}$ | 0.837 | 0.035 | 0.122 | 0.006 | 1.000 |
| Ongoing error | 0.200 | 0.032 | 0.645 | 0.000 | 0.124 | 1.000 |
| Expiring correct | 0.321 | 0.630 | 0.049 | 0.000 | 0.000 |  |
| Expiring error | 0.808 | 0.032 | 0.160 | 0.000 | 0.000 | 1.000 |

State: New Mexico

|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Current-month status | Non- <br> participating | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error | Total |


| Nonparticipating | 582,383 | 6,316 | 1,073 | 0 | 0 | 589,772 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 0* | 29,851 | 2,949 | 12,834 | 1,637 | 47,271 |
| Ongoing error | 3,024 | 532 | 2,438 | 0 | 1,339 | 7,333 |
| Expiring correct | 7,733 | 4,488 | 670 | 0 | 0 | 12,891 |
| Expiring error | 2,597 | 217 | 172 | 0 | 0 | 2,986 |
| Total | 595,736 | 41,404 | 7,303 | 12,834 | 2,976 | 660,254 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 506,213 | 2,199 | 529 | 0 | 0 | 508,941 |
| Ongoing correct | 0* | 6,641 | 1,262 | 5,497 | 794 | 14,194 |
| Ongoing error | 1,447 | 178 | 482 | 0 | 608 | 2,715 |
| Expiring correct | 3,122 | 2,044 | 344 | 0 | 0 | 5,511 |
| Expiring error | 1,248 | 93 | 78 | 0 | 0 | 1,420 |
| Total | 512,031 | 11,156 | 2,694 | 5,497 | 1,403 | 532,781 |


| Nonparticipating | 76,133 | 4,117 | 545 | 0 | 80,795 |  |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | $0^{*}$ | 23,210 | 1,688 | 7,336 | 843 | 33,077 |
| Ongoing error | 1,590 | 354 | 1,956 | 0 | 731 | 4,631 |
| Expiring correct | 4,627 | 2,447 | 326 | 0 | 0 | 1,401 |
| Expiring error | 1,355 | 120 | 94 | 0 | 0 |  |
| Total | 83,706 | 30,248 | 4,609 | 7,336 | 1,574 | 127,473 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.987 | 0.011 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.631 | 0.062 | 0.271 | 0.035 | 1.000 |
| Ongoing error | 0.412 | 0.073 | 0.332 | 0.000 | 0.183 | 1.000 |
| Expiring correct | 0.600 | 0.348 | 0.052 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.870 | 0.073 | 0.058 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.995 | 0.004 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.468 | 0.089 | 0.387 | 0.056 | 1.000 |
| Ongoing error | 0.533 | 0.066 | 0.177 | 0.000 | 0.224 | 1.000 |
| Expiring correct | 0.567 | 0.371 | 0.062 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.879 | 0.066 | 0.055 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.942 | 0.051 | 0.007 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.702 | 0.051 | 0.222 | 0.025 | 1.000 |
| Ongoing error | 0.343 | 0.076 | 0.422 | 0.000 | 0.158 | 1.000 |
| Expiring correct | 0.625 | 0.331 | 0.044 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.864 | 0.076 | 0.060 | 0.000 | 0.000 | 1.000 |

State: New York

| Current-month status | $\begin{array}{r} \text { Non- } \\ \text { participating } \end{array}$ | Next-month status |  |  |  | Total |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  |  | Ongoing correct | Ongoing error | Expiring correct | Expiring error |  |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 6,201,559 | 26,020 | 4,218 | 0 | 0 | 6,231,797 |
| Ongoing correct | 0* | 479,327 | 30,924 | 56,603 | 2,715 | 569,569 |
| Ongoing error | 27,795 | 3,002 | 66,066 | 355 | 12,051 | 109,270 |
| Expiring correct | 15,193 | 37,997 | 4,773 | 0 | 0 | 57,962 |
| Expiring error | 12,105 | 466 | 2,606 | 0 | 0 | 15,178 |
| Total | 6,256,652 | 546,813 | 108,587 | 56,959 | 14,766 | 6,983,777 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 5,335,762 | 5,089 | 1,886 | 0 | 0 | 5,342,737 |
| Ongoing correct | 0* | 64,131 | 10,464 | 10,574 | 1,158 | 86,326 |
| Ongoing error | 9,092 | 741 | 10,778 | 166 | 4,777 | 25,553 |
| Expiring correct | 371 | 8,953 | 1,521 | 0 | 0 | 10,846 |
| Expiring error | 4,828 | 214 | 974 | 0 | 0 | 6,015 |
| Total | 5,350,052 | 79,127 | 25,622 | 10,740 | 5,935 | 5,471,478 |
|  | Cell counts for households without earnings |  |  |  |  |  |
| Nonparticipating | 865,812 | 20,931 | 2,331 | 0 | 0 | 889,074 |
| Ongoing correct | 0* | 415,196 | 20,461 | 46,030 | 1,557 | 483,243 |
| Ongoing error | 18,695 | 2,262 | 55,289 | 189 | 7,274 | 83,708 |
| Expiring correct | 14,832 | 29,028 | 3,251 | 0 | 0 | 47,111 |
| Expiring error | 7,261 | 268 | 1,632 | 0 | 0 | 9,162 |
| Total | 906,600 | 467,685 | 82,964 | 46,218 | 8,831 | 1,512,299 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.995 | 0.004 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.842 | 0.054 | 0.099 | 0.005 | 1.000 |
| Ongoing error | 0.254 | 0.027 | 0.605 | 0.003 | 0.110 | 1.000 |
| Expiring correct | 0.262 | 0.656 | 0.082 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.798 | 0.031 | 0.172 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.999 | 0.001 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.743 | 0.121 | 0.122 | 0.013 | 1.000 |
| Ongoing error | 0.356 | 0.029 | 0.422 | 0.007 | 0.187 | 1.000 |
| Expiring correct | 0.034 | 0.826 | 0.140 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.803 | 0.035 | 0.162 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.974 | 0.024 | 0.003 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.859 | 0.042 | 0.095 | 0.003 | 1.000 |
| Ongoing error | 0.223 | 0.027 | 0.660 | 0.002 | 0.087 | 1.000 |
| Expiring correct | 0.315 | 0.616 | 0.069 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.793 | 0.029 | 0.178 | 0.000 | 0.000 | 1.000 |

State: North Carolina

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 2,789,873 | 18,170 | 1,755 | 0 | 0 | 2,809,798 |
| Ongoing correct | 17,892 | 114,436 | 8,415 | 23,554 | 2,879 | 167,176 |
| Ongoing error | 4,546 | 7,750 | 6,865 | 188 | 1,616 | 20,965 |
| Expiring correct | 0* | 24,368 | 3,152 | 0 | 0 | 27,520 |
| Expiring error | 2,561 | 1,753 | 317 | 0 | 0 | 4,631 |
| Total | 2,814,873 | 166,476 | 20,503 | 23,742 | 4,495 | 3,030,090 |


| Nonparticipating | $2,407,811$ | 6,943 | 1,019 | 0 | 0 | $2,415,773$ |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 3,755 | 17,375 | 4,272 | 9,528 | 1,730 | 9,469 |
| Ongoing error | 4,146 | 2,311 | 1,974 | 91 | 11,355 |  |
| Expiring correct | $0^{*}$ | 9,448 | 1,907 | 0 | 0 |  |
| Expiring error | 1,814 | 688 | 209 | 0 | 0 | 2,711 |
| Total | $2,417,527$ | 36,765 | 9,381 | 9,619 | 2,677 | $2,475,968$ |


| Nonparticipating | 381,261 | 11,227 | 736 | 0 | 0 | 393,224 |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 14,725 | 97,061 | 4,143 | 14,027 | 1,149 | 131,104 |
| Ongoing error | 462 | 5,438 | 4,891 | 97 | 11,558 |  |
| Expiring correct | $0 *$ | 15,060 | 1,245 | 0 | 16,305 |  |
| Expiring error | 899 | 925 | 108 | 0 | 0 | 1,932 |
| Total | 397,347 | 129,712 | 11,123 | 14,123 | 1,818 | 554,122 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.993 | 0.006 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.107 | 0.685 | 0.050 | 0.141 | 0.017 | 1.000 |
| Ongoing error | 0.217 | 0.370 | 0.327 | 0.009 | 0.077 | 1.000 |
| Expiring correct | 0.000* | 0.885 | 0.115 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.553 | 0.379 | 0.068 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.997 | 0.003 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.102 | 0.474 | 0.117 | 0.260 | 0.047 | 1.000 |
| Ongoing error | 0.438 | 0.244 | 0.208 | 0.010 | 0.100 | 1.000 |
| Expiring correct | 0.000* | 0.832 | 0.168 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.669 | 0.254 | 0.077 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.970 | 0.029 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.112 | 0.740 | 0.032 | 0.107 | 0.009 | 1.000 |
| Ongoing error | 0.040 | 0.471 | 0.423 | 0.008 | 0.058 | 1.000 |
| Expiring correct | 0.000* | 0.924 | 0.076 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.465 | 0.479 | 0.056 | 0.000 | 0.000 | 1.000 |

State: North Dakota

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 233,272 | 868 | 106 | 0 | 0 | 234,246 |
| Ongoing correct | 2,207 | 8,459 | 740 | 707 | 44 | 12,156 |
| Ongoing error | 0* | 1,952 | 350 | 28 | 24 | 2,354 |
| Expiring correct | 0* | 843 | 142 | 0 | 0 | 985 |
| Expiring error | 0* | 105 | 10 | 0 | 0 | 115 |
| Total | 235,479 | 12,226 | 1,348 | 735 | 68 | 249,856 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 196,658 | 370 | 74 | 0 | 0 | 197,102 |
| Ongoing correct | 940 | 2,644 | 586 | 283 | 44 | 4,497 |
| Ongoing error | 0* | 1,071 | 201 | 21 | 14 | 1,307 |
| Expiring correct | 0* | 390 | 108 | 0 | 0 | 498 |
| Expiring error | 0* | 66 | 0 | 0 | 0 | 66 |
| Total | 197,597 | 4,541 | 969 | 304 | 58 | 203,470 |
|  | Cell counts for households without earnings |  |  |  |  |  |
| Nonparticipating | 36,598 | 498 | 32 | 0 | 0 | 37,128 |
| Ongoing correct | 1,284 | 5,815 | 154 | 424 | 0 | 7,676 |
| Ongoing error | 0* | 880 | 149 | 8 | 10 | 1,047 |
| Expiring correct | 0 * | 460 | 34 | 0 | 0 | 494 |
| Expiring error | 0* | 32 | 10 | 0 | 0 | 42 |
| Total | 37,882 | 7,685 | 378 | 431 | 10 | 46,387 |


|  | Next-month status |  |  |  |  |  |
| :--- | ---: | :--- | :--- | :--- | :--- | :--- | :--- |
| Current-month status | Norticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |

## Transition probabilities for total households

| Nonparticipating | 0.996 | 0.004 | 0.000 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Ongoing correct | 0.182 | 0.696 | 0.061 | 0.058 | 0.004 |  |
| Ongoing error | $0.000^{*}$ | 0.829 | 0.149 | 0.012 | 0.010 |  |
| Expiring correct | $0.000^{*}$ | 0.856 | 0.144 | 0.000 | 1.000 |  |
| Expiring error | $0.000^{*}$ | 0.912 | 0.088 | 0.000 | 1.000 |  |
|  |  |  | 0.000 | 0.000 | 1.000 |  |

Expiring error
0.000*
0.912
0.088
1.000
Nonparticipating
Ongoing correct
Ongoing error
Expiring correct
Expiring error
Nonparticipating
Ongoing correct
Ongoing error
Expiring correct
Expiring error
0.998
0.209
$0.000^{*}$
$0.000^{*}$
$0.000^{*}$

Transition probabilities for households with earnings

| 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- |
| 0.588 | 0.130 | 0.063 | 0.010 | 1.000 |
| 0.820 | 0.154 | 0.016 | 0.011 | 1.000 |
| 0.783 | 0.217 | 0.000 | 0.000 | 1.000 |
| 1.000 | 0.000 | 0.000 | 0.000 | 1.000 |

Transition probabilities for households without earnings

| Nonparticipating | 0.986 | 0.013 | 0.001 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Ongoing correct | 0.167 | 0.758 | 0.020 | 0.055 | 0.000 | 1.000 |
| Ongoing error | $0.000^{*}$ | 0.841 | 0.142 | 0.007 | 0.010 | 1.000 |
| Expiring correct | $0.000^{*}$ | 0.932 | 0.068 | 0.000 | 0.000 |  |
| Expiring error | $0.000^{*}$ | 0.758 | 0.242 | 0.000 | 0.000 | 1.000 |

State: Ohio

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{array}{r} \text { Non- } \\ \text { participating } \end{array}$ | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 4,034,999 | 19,196 | 2,105 | 0 | 0 | 4,056,300 |
| Ongoing correct | 26,948 | 179,452 | 9,717 | 23,547 | 2,178 | 241,842 |
| Ongoing error | 0* | 14,005 | 14,790 | 0 | 1,876 | 30,672 |
| Expiring correct | 0* | 26,398 | 2,271 | 0 | 0 | 28,668 |
| Expiring error | 1,557 | 2,014 | 757 | 0 | 0 | 4,328 |
| Total | 4,063,505 | 241,066 | 29,639 | 23,547 | 4,054 | 4,361,811 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 3,355,048 | 6,481 | 999 | 0 | 0 | 3,362,529 |
| Ongoing correct | 3,772 | 24,094 | 4,592 | 10,275 | 1,604 | 44,338 |
| Ongoing error | 3,121 | 3,342 | 2,803 | 0 | 1,102 | 10,368 |
| Expiring correct | 0* | 9,776 | 1,355 | 0 | 0 | 11,131 |
| Expiring error | 1,325 | 914 | 596 | 0 | 0 | 2,835 |
| Total | 3,363,267 | 44,608 | 10,345 | 10,275 | 2,707 | 3,431,202 |

Cell counts for households without earnings

| Nonparticipating | 677,025 | 12,715 | 1,106 | 0 | 0 | 090,846 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 22,692 | 155,358 | 5,125 | 13,272 | 573 | 197,020 |
| Ongoing error | $0^{*}$ | 10,663 | 11,987 | 0 | 774 | 23,425 |
| Expiring correct | $0^{*}$ | 16,915 | 915 | 0 | 17,830 |  |
| Expiring error | 521 | 807 | 161 | 0 | 0 | 1,488 |
| Total | 700,238 | 196,458 | 19,294 | 13,272 | 1,347 | 930,609 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.995 | 0.005 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.111 | 0.742 | 0.040 | 0.097 | 0.009 | 1.000 |
| Ongoing error | 0.000* | 0.457 | 0.482 | 0.000 | 0.061 | 1.000 |
| Expiring correct | 0.000* | 0.921 | 0.079 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.360 | 0.465 | 0.175 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.998 | 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.085 | 0.543 | 0.104 | 0.232 | 0.036 | 1.000 |
| Ongoing error | 0.301 | 0.322 | 0.270 | 0.000 | 0.106 | 1.000 |
| Expiring correct | 0.000* | 0.878 | 0.122 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.467 | 0.322 | 0.210 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.980 | 0.018 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.115 | 0.789 | 0.026 | 0.067 | 0.003 | 1.000 |
| Ongoing error | 0.000* | 0.455 | 0.512 | 0.000 | 0.033 | 1.000 |
| Expiring correct | 0.000* | 0.949 | 0.051 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.350 | 0.542 | 0.108 | 0.000 | 0.000 | 1.000 |

State: Oklahoma

|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Current-month status | Non- <br> participating | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error | Total |

Nonparticipating
Ongoing correct
Ongoing error
Expiring correct
Expiring error
Total

| $1,181,225$ | 12,602 | 2,200 |
| ---: | ---: | ---: |
| 15,843 | 56,402 | 5,771 |
| 3,233 | 6,724 | 3,966 |
| $0^{*}$ | 11,232 | 2,085 |
| 1,104 | 1,221 | 279 |
| $1,201,405$ | 88,182 | 14,301 |


| 0 |
| ---: |
| 8,176 |
| 22 |
| 0 |
| 0 |
| 8,199 |


| 0 | $1,196,028$ |
| ---: | ---: |
| 2,121 | 88,313 |
| 443 | 14,388 |
| 0 | 13,317 |
| 0 | 2,604 |
| 2,564 | $1,314,651$ |

Cell counts for households with earnings

| Nonparticipating | $1,000,102$ | 5,236 | 1,343 | 0 | 0 | $1,006,682$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 4,272 | 8,558 | 2,690 | 3,856 | 1,649 | 21,025 |
| Ongoing error | 3,508 | 1,892 | 701 | 22 | 209 | 6,333 |
| Expiring correct | $0^{*}$ | 4,837 | 1,458 | 0 | 0 | 6,295 |
| Expiring error | 1,172 | 568 | 139 | 0 | 0 | 1,879 |
| Total | $1,009,054$ | 21,090 | 6,331 | 3,878 | 1,858 | $1,042,212$ |


| Nonparticipating | 180,208 | 7,366 | 857 | 0 | 0 | 188,431 |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 11,989 | 4,844 | 3,081 | 4,321 | 472 | 67,707 |
| Ongoing error | $0^{*}$ | 4,833 | 3,264 | 0 | 234 | 8,331 |
| Expiring correct | $0^{*}$ | 6,614 | 627 | 0 | 0 | 7,241 |
| Expiring error | 154 | 435 | 140 | 0 | 0 | 729 |
| Total | 192,351 | 67,092 | 7,969 | 4,321 | 706 | 272,439 |


|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Non- | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error | Total |

## Transition probabilities for total households

| Nonparticipating | 0.988 | 0.011 | 0.002 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Ongoing correct | 0.179 | 0.639 | 0.065 | 0.093 | 0.024 | 1.000 |
| Ongoing error | 0.225 | 0.467 | 0.276 | 0.002 | 0.031 |  |
| Expiring correct | $0.000^{*}$ | 0.843 | 0.157 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.424 | 0.469 | 0.107 | 0.000 | 0.000 | 1.000 |

Expiring error
0.424
0.469
$0.107 \quad 0.000$
1.000

Transition probabilities for households with earnings

| 0.993 | 0.005 | 0.001 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- |
| 0.203 | 0.407 | 0.128 | 0.183 | 0.078 | 1.000 |
| 0.554 | 0.299 | 0.111 | 0.004 | 0.033 | 1.000 |
| $0.000^{*}$ | 0.768 | 0.232 | 0.000 | 0.000 | 1.000 |
| 0.624 | 0.302 | 0.074 | 0.000 | 0.000 | 1.000 |

Transition probabilities for households without earnings

| Nonparticipating | 0.956 | 0.039 | 0.005 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Ongoing correct | 0.177 | 0.707 | 0.046 | 0.064 | 0.007 | 1.000 |
| Ongoing error | $0.000^{*}$ | 0.580 | 0.392 | 0.000 | 0.028 |  |
| Expiring correct | $0.000^{*}$ | 0.913 | 0.087 | 0.000 | 0.000 |  |
| Expiring error | 0.211 | 0.596 | 0.192 | 0.000 | 0.000 | 1.000 |

State: Oregon

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 1,217,018 | 6,277 | 1,250 | 0 | 0 | 1,224,544 |
| Ongoing correct | 29,750 | 55,049 | 8,225 | 3,372 | 292 | 96,689 |
| Ongoing error | 0* | 25,905 | 5,467 | 112 | 304 | 31,788 |
| Expiring correct | 0* | 8,822 | 1,344 | 0 | 0 | 10,166 |
| Expiring error | 0* | 927 | 56 | 0 | 0 | 983 |
| Total | 1,246,768 | 96,979 | 16,342 | 3,484 | 597 | 1,364,170 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 1,036,246 | 1,830 | 565 | 0 | 0 | 1,038,641 |
| Ongoing correct | 10,318 | 10,873 | 4,348 | 1,088 | 235 | 26,862 |
| Ongoing error | 0* | 10,799 | 1,885 | 58 | 191 | 12,933 |
| Expiring correct | 0* | 2,879 | 990 | 0 | 0 | 3,869 |
| Expiring error | 0* | 569 | 29 | 0 | 0 | 599 |
| Total | 1,046,564 | 26,950 | 7,818 | 1,146 | 426 | 1,082,904 |
|  | Cell counts for households without earnings |  |  |  |  |  |
| Nonparticipating | 180,784 | 4,447 | 684 | 0 | 0 | 185,915 |
| Ongoing correct | 19,420 | 44,176 | 3,878 | 2,284 | 58 | 69,815 |
| Ongoing error | 0* | 15,106 | 3,582 | 54 | 114 | 18,855 |
| Expiring correct | 0* | 5,992 | 354 | 0 | 0 | 6,347 |
| Expiring error | 0* | 308 | 26 | 0 | 0 | 334 |
| Total | 200,204 | 70,029 | 8,525 | 2,338 | 171 | 281,267 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.994 | 0.005 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.308 | 0.569 | 0.085 | 0.035 | 0.003 | 1.000 |
| Ongoing error | 0.000* | 0.815 | 0.172 | 0.004 | 0.010 | 1.000 |
| Expiring correct | 0.000* | 0.868 | 0.132 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.000* | 0.943 | 0.057 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.998 | 0.002 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.384 | 0.405 | 0.162 | 0.041 | 0.009 | 1.000 |
| Ongoing error | 0.000* | 0.835 | 0.146 | 0.004 | 0.015 | 1.000 |
| Expiring correct | 0.000* | 0.744 | 0.256 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.000* | 0.951 | 0.049 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.972 | 0.024 | 0.004 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.278 | 0.633 | 0.056 | 0.033 | 0.001 | 1.000 |
| Ongoing error | 0.000* | 0.801 | 0.190 | 0.003 | 0.006 | 1.000 |
| Expiring correct | 0.000* | 0.944 | 0.056 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.000* | 0.921 | 0.079 | 0.000 | 0.000 | 1.000 |

State: Pennsylvania

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{array}{r} \text { Non- } \\ \text { participating } \end{array}$ | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 4,309,278 | 20,190 | 1,731 | 0 | 0 | 4,331,198 |
| Ongoing correct | 15,210 | 241,941 | 19,306 | 19,034 | 876 | 296,368 |
| Ongoing error | 1,985 | 20,147 | 26,436 | 0 | 1,805 | 50,374 |
| Expiring correct | 6,382 | 11,664 | 1,350 | 0 | 0 | 19,396 |
| Expiring error | 634 | 1,100 | 1,016 | 0 | 0 | 2,750 |
| Total | 4,333,490 | 295,042 | 49,840 | 19,034 | 2,682 | 4,700,087 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 3,490,770 | 4,662 | 434 | 0 | 0 | 3,495,866 |
| Ongoing correct | 5,332 | 47,535 | 10,867 | 3,547 | 786 | 68,067 |
| Ongoing error | 0* | 13,326 | 9,952 | 0 | 735 | 24,013 |
| Expiring correct | 1,085 | 1,691 | 812 | 0 | 0 | 3,588 |
| Expiring error | 112 | 901 | 521 | 0 | 0 | 1,534 |
| Total | 3,497,299 | 68,114 | 22,587 | 3,547 | 1,521 | 3,593,069 |

Cell counts for households without earnings

| Nonparticipating | 817,280 | 15,528 | 1,296 | 0 | 034,105 |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 9,807 | 194,407 | 8,439 | 15,487 | 228,229 |
| Ongoing error | 3,291 | 6,821 | 16,484 | 0 | 27,667 |
| Expiring correct | 5,391 | 9,872 | 538 | 0 | 1,071 |
| Expiring error | 421 | 300 | 495 | 0 | 0 |
| Total | 836,191 | 226,928 | 27,253 | 15,487 | 1,216 |
|  |  |  |  | 1,161 | $1,107,019$ |

State: Pennsylvania

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{array}{r} \text { Non- } \\ \text { participating } \end{array}$ | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.995 | 0.005 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.051 | 0.816 | 0.065 | 0.064 | 0.003 | 1.000 |
| Ongoing error | 0.039 | 0.400 | 0.525 | 0.000 | 0.036 | 1.000 |
| Expiring correct | 0.329 | 0.601 | 0.070 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.231 | 0.400 | 0.369 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.999 | 0.001 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.078 | 0.698 | 0.160 | 0.052 | 0.012 | 1.000 |
| Ongoing error | 0.000* | 0.555 | 0.414 | 0.000 | 0.031 | 1.000 |
| Expiring correct | 0.302 | 0.471 | 0.226 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.073 | 0.587 | 0.340 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.980 | 0.019 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.043 | 0.852 | 0.037 | 0.068 | 0.000 | 1.000 |
| Ongoing error | 0.119 | 0.247 | 0.596 | 0.000 | 0.039 | 1.000 |
| Expiring correct | 0.341 | 0.625 | 0.034 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.347 | 0.247 | 0.407 | 0.000 | 0.000 | 1.000 |

State: Rhode Island

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{array}{r} \text { Non- } \\ \text { participating } \\ \hline \end{array}$ | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 371,761 | 1,283 | 94 | 0 | 0 | 373,139 |
| Ongoing correct | 370 | 21,104 | 1,364 | 3,342 | 112 | 26,291 |
| Ongoing error | 0* | 2,047 | 1,652 | 0 | 152 | 3,850 |
| Expiring correct | 1,589 | 1,737 | 128 | 0 | 0 | 3,454 |
| Expiring error | 28 | 174 | 80 | 0 | 0 | 283 |
| Total | 373,748 | 26,345 | 3,318 | 3,342 | 264 | 407,018 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 310,274 | 272 | 40 | 0 | 0 | 310,587 |
| Ongoing correct | 314 | 3,172 | 787 | 243 | 69 | 4,585 |
| Ongoing error | 92 | 805 | 637 | 0 | 29 | 1,563 |
| Expiring correct | 0* | 308 | 64 | 0 | 0 | 372 |
| Expiring error | 12 | 55 | 40 | 0 | 0 | 107 |
| Total | 310,691 | 4,612 | 1,569 | 243 | 98 | 317,214 |
|  | Cell counts for households without earnings |  |  |  |  |  |
| Nonparticipating | 61,177 | 1,011 | 54 | 0 | 0 | 62,242 |
| Ongoing correct | 141 | 17,932 | 577 | 3,100 | 43 | 21,792 |
| Ongoing error | 0* | 1,242 | 1,014 | 0 | 123 | 2,379 |
| Expiring correct | 1,726 | 1,425 | 63 | 0 | 0 | 3,214 |
| Expiring error | 13 | 124 | 40 | 0 | 0 | 177 |
| Total | 63,057 | 21,733 | 1,749 | 3,100 | 166 | 89,805 |


|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
|  | Non- <br> participating | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error | Total |

Transition probabilities for total households

| Nonparticipating | 0.996 | 0.003 | 0.000 | 0.000 | 0.000 | 1.000 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 0.014 | 0.803 | 0.052 | 0.127 | 0.004 | 1.000 |
| Ongoing error | 0.000* | 0.532 | 0.429 | 0.000 | 0.039 | 1.000 |
| Expiring correct | 0.460 | 0.503 | 0.037 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.101 | 0.615 | 0.284 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.999 | 0.001 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.068 | 0.692 | 0.172 | 0.053 | 0.015 | 1.000 |
| Ongoing error | 0.059 | 0.515 | 0.408 | 0.000 | 0.018 | 1.000 |
| Expiring correct | 0.000* | 0.827 | 0.173 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.111 | 0.515 | 0.374 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.983 | 0.016 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.006 | 0.823 | 0.026 | 0.142 | 0.002 | 1.000 |
| Ongoing error | 0.000* | 0.522 | 0.426 | 0.000 | 0.052 | 1.000 |
| Expiring correct | 0.537 | 0.443 | 0.020 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.071 | 0.701 | 0.228 | 0.000 | 0.000 | 1.000 |

State: South Carolina

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{array}{r} \text { Non- } \\ \text { participating } \end{array}$ | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 1,352,323 | 8,045 | 609 | 0 | 0 | 1,360,976 |
| Ongoing correct | 13,461 | 75,703 | 4,567 | 4,720 | 242 | 98,693 |
| Ongoing error | 4,287 | 2,216 | 4,004 | 0 | 289 | 10,797 |
| Expiring correct | 0* | 12,158 | 1,465 | 0 | 0 | 13,623 |
| Expiring error | 319 | 113 | 120 | 0 | 0 | 552 |
| Total | 1,370,390 | 98,236 | 10,764 | 4,720 | 531 | 1,484,642 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 1,131,591 | 2,498 | 232 | 0 | 0 | 1,134,321 |
| Ongoing correct | 5,688 | 18,551 | 2,520 | 776 | 125 | 27,661 |
| Ongoing error | 2,822 | 908 | 944 | 0 | 59 | 4,733 |
| Expiring correct | 0* | 5,569 | 983 | 0 | 0 | 6,552 |
| Expiring error | 130 | 38 | 30 | 0 | 0 | 197 |
| Total | 1,140,232 | 27,564 | 4,708 | 776 | 184 | 1,173,464 |

Cell counts for households without earnings

| Nonparticipating | 220,810 | 5,547 | 377 | 0 | 0 | 226,734 |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 7,700 | 57,152 | 2,047 | 3,944 | 117 | 2,961 |
| Ongoing error | 1,460 | 1,308 | 3,060 | 0 | 2,059 |  |
| Expiring correct | $0 *$ | 6,588 | 482 | 0 | 0,070 |  |
| Expiring error | 188 | 76 | 90 | 0 | 0 | 354 |
| Total | 230,158 | 70,671 | 6,056 | 3,944 | 348 | 311,178 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{array}{r} \text { Non- } \\ \text { participating } \\ \hline \end{array}$ | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.994 | 0.006 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.136 | 0.767 | 0.046 | 0.048 | 0.002 | 1.000 |
| Ongoing error | 0.397 | 0.205 | 0.371 | 0.000 | 0.027 | 1.000 |
| Expiring correct | 0.000* | 0.892 | 0.108 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.578 | 0.205 | 0.217 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.998 | 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.206 | 0.671 | 0.091 | 0.028 | 0.005 | 1.000 |
| Ongoing error | 0.596 | 0.192 | 0.199 | 0.000 | 0.012 | 1.000 |
| Expiring correct | 0.000* | 0.850 | 0.150 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.658 | 0.192 | 0.150 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.974 | 0.024 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.109 | 0.805 | 0.029 | 0.056 | 0.002 | 1.000 |
| Ongoing error | 0.241 | 0.216 | 0.505 | 0.000 | 0.038 | 1.000 |
| Expiring correct | 0.000* | 0.932 | 0.068 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.530 | 0.216 | 0.254 | 0.000 | 0.000 | 1.000 |

State: South Dakota

|  | Next-month status |  |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Current-month status | Non- <br> participating | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error | Total |

Cell counts for total households

| Nonparticipating | 265,764 | 1,548 | 48 | 0 | 0 | 267,361 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 347 | 13,025 | 358 | 1,085 | 23 | 14,839 |
| Ongoing error | 406 | 0 | 245 | 0 | 58 | 709 |
| Expiring correct | 827 | 276 | 11 | 0 | 0 | 1,114 |
| Expiring error | 48 | 0 | 35 | 0 | 0 | 83 |
| Total | 267,392 | 14,850 | 698 | 1,085 | 81 | 284,106 |
|  | Cell counts for households with earnings |  |  |  |  |  |


| Nonparticipating | 225,801 | 614 | 23 | 0 | 0 | 226,439 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 87 | 5,022 | 262 | 318 | 23 | 5,712 |
| Ongoing error | 285 | 0 | 94 | 0 | 12 | 391 |
| Expiring correct | 242 | 85 | 0 | 0 | 0 | 327 |
| Expiring error | 37 | 0 | 0 | 0 | 0 | 37 |
|  | 226,452 | 5,721 | 379 | 318 | 35 | 232,906 |


| Nonparticipating | 39,961 | 934 | 25 | 0 | 40,920 |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 262 | 8,003 | 95 | 0 | 9,128 |  |
| Ongoing error | 121 | 0 | 151 | 367 | 46 |  |
| Expiring correct | 586 | 191 | 11 | 0 | 0 | 788 |
| Expiring error | 10 | 0 | 35 | 0 | 0 |  |
| Total | 40,940 | 9,128 | 319 | 0 | 46 |  |


|  | Next-month status |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
|  | Non- <br> Current-month status | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error |
|  | Transition probabilities for total households |  |  |  |  |

State: Tennessee

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{array}{r} \text { Non- } \\ \text { participating } \end{array}$ | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 1,941,130 | 12,314 | 1,350 | 0 | 0 | 1,954,794 |
| Ongoing correct | 0* | 125,921 | 6,666 | 39,963 | 3,400 | 175,950 |
| Ongoing error | 6,979 | 1,095 | 6,431 | 186 | 2,613 | 17,304 |
| Expiring correct | 18,103 | 20,412 | 1,804 | 0 | 0 | 40,319 |
| Expiring error | 5,035 | 457 | 686 | 0 | 0 | 6,178 |
| Total | 1,971,246 | 160,199 | 16,936 | 40,150 | 6,013 | 2,194,545 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 1,691,768 | 4,726 | 725 | 0 | 0 | 1,697,218 |
| Ongoing correct | $0^{*}$ | 19,300 | 3,033 | 14,565 | 2,307 | 39,205 |
| Ongoing error | 3,768 | 183 | 1,155 | 124 | 1,217 | 6,447 |
| Expiring correct | 5,542 | 7,929 | 1,307 | 0 | 0 | 14,777 |
| Expiring error | 3,243 | 172 | 188 | 0 | 0 | 3,603 |
| Total | 1,704,321 | 32,309 | 6,408 | 14,689 | 3,524 | 1,761,250 |
|  | Cell counts for households without earnings |  |  |  |  |  |
| Nonparticipating | 249,318 | 7,588 | 626 | 0 | 0 | 257,532 |
| Ongoing correct | 0* | 106,621 | 3,632 | 25,399 | 1,094 | 136,746 |
| Ongoing error | 3,226 | 912 | 5,276 | 62 | 1,396 | 10,872 |
| Expiring correct | 12,532 | 12,539 | 497 | 0 | 0 | 25,567 |
| Expiring error | 1,849 | 231 | 498 | 0 | 0 | 2,578 |
| Total | 266,926 | 127,891 | 10,529 | 25,461 | 2,490 | 433,296 |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.993 | 0.006 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.716 | 0.038 | 0.227 | 0.019 | 1.000 |
| Ongoing error | 0.403 | 0.063 | 0.372 | 0.011 | 0.151 | 1.000 |
| Expiring correct | 0.449 | 0.506 | 0.045 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.815 | 0.074 | 0.111 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.997 | 0.003 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.492 | 0.077 | 0.372 | 0.059 | 1.000 |
| Ongoing error | 0.584 | 0.028 | 0.179 | 0.019 | 0.189 | 1.000 |
| Expiring correct | 0.375 | 0.537 | 0.088 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.900 | 0.048 | 0.052 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.968 | 0.029 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.780 | 0.027 | 0.186 | 0.008 | 1.000 |
| Ongoing error | 0.297 | 0.084 | 0.485 | 0.006 | 0.128 | 1.000 |
| Expiring correct | 0.490 | 0.490 | 0.019 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.717 | 0.090 | 0.193 | 0.000 | 0.000 | 1.000 |

State: Texas

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{array}{r} \text { Non- } \\ \text { participating } \end{array}$ | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 6,665,351 | 52,086 | 2,816 | 0 | 0 | 6,720,253 |
| Ongoing correct | 29,532 | 284,309 | 14,949 | 76,552 | 4,558 | 409,900 |
| Ongoing error | 0* | 15,386 | 12,809 | 85 | 6,207 | 34,486 |
| Expiring correct | 26,761 | 49,379 | 1,692 | 0 | 0 | 77,832 |
| Expiring error | 4,905 | 5,006 | 1,239 | 0 | 0 | 11,150 |
| Total | 6,726,550 | 406,166 | 33,505 | 76,637 | 10,765 | 7,253,623 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 5,924,138 | 21,195 | 1,681 | 0 | 0 | 5,947,013 |
| Ongoing correct | 0* | 89,034 | 9,332 | 37,672 | 3,318 | 139,356 |
| Ongoing error | 2,342 | 6,477 | 6,335 | 85 | 3,743 | 18,981 |
| Expiring correct | 18,163 | 19,006 | 787 | 0 | 0 | 37,956 |
| Expiring error | 4,283 | 2,512 | 471 | 0 | 0 | 7,266 |
| Total | 5,948,926 | 138,224 | 18,605 | 37,756 | 7,061 | 6,150,573 |

Cell counts for households without earnings

| Nonparticipating | 736,932 | 30,891 | 1,135 | 0 | 0 | 768,959 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 31,192 | 195,275 | 5,618 | 38,881 | 1,240 | 272,205 |
| Ongoing error | $0^{*}$ | 8,909 | 6,474 | 0 | 2,464 | 4,847 |
| Expiring correct | 8,593 | 30,632 | 906 | 0 | 0 | 0,130 |
| Expiring error | 907 | 2,235 | 767 | 0 | 3,909 |  |
|  | 267,942 | 14,900 | 38,881 | 3,703 | $1,103,050$ |  |


| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.992 | 0.008 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.072 | 0.694 | 0.036 | 0.187 | 0.011 | 1.000 |
| Ongoing error | 0.000* | 0.446 | 0.371 | 0.002 | 0.180 | 1.000 |
| Expiring correct | 0.344 | 0.634 | 0.022 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.440 | 0.449 | 0.111 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.996 | 0.004 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.639 | 0.067 | 0.270 | 0.024 | 1.000 |
| Ongoing error | 0.123 | 0.341 | 0.334 | 0.004 | 0.197 | 1.000 |
| Expiring correct | 0.479 | 0.501 | 0.021 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.589 | 0.346 | 0.065 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.958 | 0.040 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.115 | 0.717 | 0.021 | 0.143 | 0.005 | 1.000 |
| Ongoing error | 0.000* | 0.499 | 0.363 | 0.000 | 0.138 | 1.000 |
| Expiring correct | 0.214 | 0.763 | 0.023 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.232 | 0.572 | 0.196 | 0.000 | 0.000 | 1.000 |

State: Utah

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 662,401 | 2,841 | 341 | 0 | 0 | 665,583 |
| Ongoing correct | 2,347 | 16,890 | 1,941 | 3,500 | 609 | 25,288 |
| Ongoing error | 804 | 1,621 | 1,116 | 0 | 535 | 4,075 |
| Expiring correct | 0* | 3,349 | 512 | 0 | 0 | 3,861 |
| Expiring error | 534 | 453 | 152 | 0 | 0 | 1,139 |
| Total | 666,086 | 25,153 | 4,062 | 3,500 | 1,144 | 699,945 |


| Nonparticipating | 582,345 | 929 | 283 | 0 | 0 | 583,557 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 68 | 4,177 | 1,115 | 1,606 | 432 | 2,399 |
| Ongoing error | 795 | 732 | 395 | 0 | 2,198 |  |
| Expiring correct | 61 | 1,230 | 335 | 0 | 0 | 1,626 |
| Expiring error | 417 | 235 | 54 | 0 | 707 |  |
|  | 583,686 | 7,304 | 2,183 | 1,606 | 708 | 595,487 |

Cell counts for households without earnings

| Nonparticipating | 80,019 | 1,912 | 58 | 0 | 0 | 1,990 |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 2,245 | 12,713 | 826 | 1,894 | 177 | 17,855 |
| Ongoing error | 6 | 888 | 720 | 0 | 1,873 |  |
| Expiring correct | $0^{*}$ | 2,131 | 177 | 0 | 0 | 2,308 |
| Expiring error | 130 | 205 | 97 | 0 | 0 | 432 |
|  | 17,850 | 1,879 | 1,894 | 436 | 104,459 |  |

State: Utah

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.995 | 0.004 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.093 | 0.668 | 0.077 | 0.138 | 0.024 | 1.000 |
| Ongoing error | 0.197 | 0.398 | 0.274 | 0.000 | 0.131 | 1.000 |
| Expiring correct | 0.000* | 0.867 | 0.133 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.469 | 0.398 | 0.133 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.998 | 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.009 | 0.565 | 0.151 | 0.217 | 0.058 | 1.000 |
| Ongoing error | 0.362 | 0.333 | 0.180 | 0.000 | 0.125 | 1.000 |
| Expiring correct | 0.038 | 0.757 | 0.206 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.590 | 0.333 | 0.077 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.976 | 0.023 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.126 | 0.712 | 0.046 | 0.106 | 0.010 | 1.000 |
| Ongoing error | 0.003 | 0.474 | 0.385 | 0.000 | 0.138 | 1.000 |
| Expiring correct | 0.000* | 0.923 | 0.077 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.301 | 0.474 | 0.225 | 0.000 | 0.000 | 1.000 |

State: Vermont

|  | Next-month status |  |  |  |  |  |
| :--- | :---: | :---: | :---: | ---: | ---: | ---: |
| Current-month <br> status | Non- <br> participating | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error | Total |


| Nonparticipating | 186,330 | 257 | 107 | 0 | 0 | 186,693 |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
| Ongoing correct | 0* | 2,425 | 452 | 338 | 118 | 3,333 |
| Ongoing error | 350 | 188 | 555 | 0 | 68 | 1,162 |
| Expiring correct | 99 | 207 | 39 | 0 | 0 | 344 |
| Expiring error | 158 | 31 | 0 | 0 | 0 | 188 |
| Total | 186,937 | 3,107 | 1,152 | 338 | 186 | 191,721 |


| Nonparticipating | 28,965 | 881 | 65 | 0 | 0 | 29,910 |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | $0^{*}$ | 10,670 | 472 | 1,381 | 0 | 12,524 |
| Ongoing error | 251 | 252 | 922 | 0 | 74 | 0 |
| Expiring correct | 721 | 629 | 53 | 0 | 1,400 |  |
| Expiring error | 66 | 13 | 0 | 0 | 0 | 79 |
|  | 12,446 | 1,512 | 1,381 | 74 | 45,416 |  |

State: Vermont

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.994 | 0.005 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.826 | 0.058 | 0.108 | 0.007 | 1.000 |
| Ongoing error | 0.226 | 0.165 | 0.555 | 0.000 | 0.053 | 1.000 |
| Expiring correct | 0.470 | 0.478 | 0.052 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.835 | 0.165 | 0.000 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.998 | 0.001 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.728 | 0.136 | 0.101 | 0.035 | 1.000 |
| Ongoing error | 0.302 | 0.162 | 0.478 | 0.000 | 0.058 | 1.000 |
| Expiring correct | 0.286 | 0.601 | 0.112 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.838 | 0.162 | 0.000 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.968 | 0.029 | 0.002 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.852 | 0.038 | 0.110 | 0.000 | 1.000 |
| Ongoing error | 0.167 | 0.168 | 0.615 | 0.000 | 0.050 | 1.000 |
| Expiring correct | 0.514 | 0.449 | 0.037 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.832 | 0.168 | 0.000 | 0.000 | 0.000 | 1.000 |

State: Virginia

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 2,479,114 | 9,863 | 1,222 | 0 | 0 | 2,490,199 |
| Ongoing correct | 0* | 90,200 | 8,195 | 17,430 | 3,432 | 119,258 |
| Ongoing error | 7,268 | 3,079 | 5,732 | 40 | 1,387 | 17,507 |
| Expiring correct | 2,282 | 13,405 | 2,002 | 0 | 0 | 17,688 |
| Expiring error | 4,033 | 874 | 0 | 0 | 0 | 4,908 |
| Total | 2,492,698 | 117,421 | 17,151 | 17,470 | 4,819 | 2,649,559 |


| Nonparticipating | $2,151,777$ | 2,897 | 727 | 0 | 0 | $2,155,401$ |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | $0^{*}$ | 16,846 | 5,053 | 5,392 | 2,902 | 30,193 |
| Ongoing error | 5,834 | 1,193 | 1,841 | 0 | 597 | 0,465 |
| Expiring correct | $0^{*}$ | 5,555 | 1,613 | 0 | 7,168 |  |
| Expiring error | 3,099 | 447 | 0 | 0 | 0 | 3,546 |
| Total | $2,160,710$ | 26,939 | 9,234 | 5,392 | 3,499 | $2,205,774$ |


| Nonparticipating | 323,392 | 6,966 | 495 | 0 | 0 | 330,852 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 2,203 | 73,354 | 3,142 | 12,038 | 530 | 9,267 |
| Ongoing error | 1,438 | 1,886 | 3,891 | 40 | 790 | 0 |
| Expiring correct | 3,920 | 7,951 | 389 | 0 | 12,259 |  |
| Expiring error | 1,036 | 326 | 0 | 0 | 0 | 1,362 |
|  | 331,988 | 90,482 | 7,917 | 12,078 | 1,320 | 443,786 |

State: Virginia

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | $\begin{array}{r} \text { Non- } \\ \text { participating } \end{array}$ | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.996 | 0.004 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.756 | 0.069 | 0.146 | 0.029 | 1.000 |
| Ongoing error | 0.415 | 0.176 | 0.327 | 0.002 | 0.079 | 1.000 |
| Expiring correct | 0.129 | 0.758 | 0.113 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.822 | 0.178 | 0.000 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.998 | 0.001 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.558 | 0.167 | 0.179 | 0.096 | 1.000 |
| Ongoing error | 0.616 | 0.126 | 0.195 | 0.000 | 0.063 | 1.000 |
| Expiring correct | 0.000* | 0.775 | 0.225 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.874 | 0.126 | 0.000 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.977 | 0.021 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.024 | 0.804 | 0.034 | 0.132 | 0.006 | 1.000 |
| Ongoing error | 0.179 | 0.234 | 0.484 | 0.005 | 0.098 | 1.000 |
| Expiring correct | 0.320 | 0.649 | 0.032 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.761 | 0.239 | 0.000 | 0.000 | 0.000 | 1.000 |

State: Washington

|  | Next-month status |  |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: | :---: |
| Current-month <br> status | Non- <br> participating | Ongoing <br> correct | Ongoing <br> error | Expiring <br> correct | Expiring <br> error | Total |

Cell counts for total households

| Nonparticipating | $2,096,392$ | 12,714 | 1,842 | 0 | 0 | $2,110,947$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 4,762 | 78,720 | 7,090 | 16,938 | 1,453 | 108,964 |
| Ongoing error | 4,097 | 6,165 | 4,313 | 88 | 1,027 | 15,691 |
| Expiring correct | 6,043 | 9,347 | 1,499 | 0 | 0 | 16,889 |
| Expiring error | 1,089 | 991 | 408 | 0 | 2,488 |  |
| Total | $2,112,383$ | 107,938 | 15,152 | 17,026 | 2,480 | $2,254,978$ |
|  |  |  |  |  |  |  |


| Nonparticipating | $1,792,455$ | 3,202 | 978 | 0 | 0 |
| :--- | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 795 | 13,184 | 3,131 | 3,484 | 595 |
| Ongoing error | 2,505 | 2,166 | 1,228 | 58 | 258 |
| Expiring correct | 720 | 2,247 | 517 | 0 | 0,214 |
| Expiring error | 295 | 303 | 250 | 0 | 0,484 |
| Total | $1,796,769$ | 21,101 | 6,104 | 3,542 | 0 |
|  |  |  |  | 853 | $1,828,370$ |


| Nonparticipating | 304,109 | 9,512 | 864 | 0 | 0 | 314,484 |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | 3,833 | 65,537 | 3,959 | 13,454 | 858 | 9,640 |
| Ongoing error | 1,574 | 4,000 | 3,086 | 30 | 769 | 0 |
| Expiring correct | 5,315 | 7,091 | 982 | 0 | 13,388 |  |
| Expiring error | 782 | 698 | 158 | 0 | 0 | 1,638 |
|  | 315,613 | 86,836 | 9,048 | 13,484 | 1,626 | 426,609 |

State: Washington

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.993 | 0.006 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.044 | 0.722 | 0.065 | 0.155 | 0.013 | 1.000 |
| Ongoing error | 0.261 | 0.393 | 0.275 | 0.006 | 0.065 | 1.000 |
| Expiring correct | 0.358 | 0.553 | 0.089 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.438 | 0.399 | 0.164 | 0.000 | 0.000 | 1.000 |
|  | Transition probabilities for households with earnings |  |  |  |  |  |
| Nonparticipating | 0.998 | 0.002 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.038 | 0.622 | 0.148 | 0.164 | 0.028 | 1.000 |
| Ongoing error | 0.403 | 0.349 | 0.198 | 0.009 | 0.042 | 1.000 |
| Expiring correct | 0.207 | 0.645 | 0.148 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.348 | 0.358 | 0.295 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.967 | 0.030 | 0.003 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.044 | 0.748 | 0.045 | 0.154 | 0.010 | 1.000 |
| Ongoing error | 0.166 | 0.423 | 0.326 | 0.003 | 0.081 | 1.000 |
| Expiring correct | 0.397 | 0.530 | 0.073 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.477 | 0.426 | 0.096 | 0.000 | 0.000 | 1.000 |

State: West Virginia

|  | Next-month status |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Current-month <br> status | Non- | Ongoing | Ongoing | Expiring | Expiring |

Cell counts for total households

| Nonparticipating | 604,049 | 5,120 | 545 | 0 | 0 | 609,714 |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | $0^{*}$ | 62,252 | 4,104 | 15,307 | 1,355 | 83,018 |
| Ongoing error | 4,132 | 113 | 4,600 | 0 | 1,307 | 10,152 |
| Expiring correct | 9,304 | 5,379 | 556 | 0 | 15,239 |  |
| Expiring error | 2,529 | 30 | 137 | 0 | 0 | 2,695 |
| Total | 620,014 | 72,894 | 9,941 | 2,662 | 720,819 |  |
|  |  |  |  |  |  |  |


|  |  |  |  |  |  |  |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Nonparticipating | 484,600 | 1,750 | 343 | 0 | 0 | 486,693 |
| Ongoing correct | $0^{*}$ | 9,619 | 1,902 | 6,310 | 838 | 18,669 |
| Ongoing error | 1,944 | 80 | 799 | 0 | 667 | 0 |
| Expiring correct | 4,248 | 1,623 | 333 | 0 | 6,203 |  |
| Expiring error | 1,453 | 35 | 24 | 0 | 0 | 1,512 |
| Total | 492,244 | 13,107 | 3,400 | 6,310 | 1,506 | 516,568 |

Cell counts for households without earnings

| Nonparticipating | 119,390 | 3,370 | 201 | 0 | 0 | 122,961 |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | $0^{*}$ | 52,634 | 2,202 | 8,997 | 516 | 64,350 |
| Ongoing error | 2,214 | 33 | 3,801 | 0 | 640 | 0 |
| Expiring correct | 5,098 | 3,745 | 223 | 0 | 9,066 |  |
| Expiring error | 1,069 | 6 | 113 | 0 | 0 | 1,187 |
|  | 127,770 | 59,787 | 6,541 | 8,997 | 1,156 | 204,251 |

State: West Virginia

|  |  |  | Next-month status |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Current-month <br> status | Non- | Ongoing | Ongoing | Expiring | Expiring |

Transition probabilities for total households

| Nonparticipating | 0.991 | 0.008 | 0.001 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Ongoing correct | $0.000^{*}$ | 0.750 | 0.049 | 0.184 | 0.016 | 1.000 |
| Ongoing error | 0.407 | 0.011 | 0.453 | 0.000 | 0.129 | 1.000 |
| Expiring correct | 0.611 | 0.353 | 0.036 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.938 | 0.011 | 0.051 | 0.000 | 0.000 | 1.000 |

Expiring error
0.938

Transition probabilities for households with earnings
Nonparticipating
Ongoing correct
Ongoing error
Expiring correct
Expiring error
0.996
$0.000^{*}$
0.557
0.685
0.961

| 0.004 | 0.001 |
| :--- | :--- |
| 0.515 | 0.102 |
| 0.023 | 0.229 |
| 0.262 | 0.054 |
| 0.023 | 0.016 |

0.000
0.338
0.000
0.000
0.000
0.00
1.000

Nonparticipating
correc

Expiring correct
0.961
$\begin{array}{cc}\text { Transition probabilities for hou } \\ & \\ 0.027 & 0.002 \\ 0.818 & 0.034 \\ 0.005 & 0.568 \\ 0.413 & 0.025 \\ 0.005 & 0.095\end{array}$

| Nonparticipating | 0.971 | 0.027 | 0.002 | 0.000 | 0.000 | 1.000 |
| :--- | :--- | :--- | :--- | :--- | :--- | :--- |
| Ongoing correct | $0.000^{*}$ | 0.818 | 0.034 | 0.140 | 0.008 | 1.000 |
| Ongoing error | 0.331 | 0.005 | 0.568 | 0.000 | 0.096 |  |
| Expiring correct | 0.562 | 0.413 | 0.025 | 0.000 | 0.000 |  |
| Expiring error | 0.900 | 0.005 | 0.095 | 0.000 | 0.000 | 1.000 |

State: Wisconsin

|  | Next-month status |  |  |  |  |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Current-month <br> status | Non- | Ongoing | Ongoing | Expiring | Expiring |

Cell counts for total households

| Nonparticipating | $1,990,174$ | 7,171 | 1,260 | 0 | 0 | $1,998,605$ |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | $0^{*}$ | 39,737 | 3,824 | 10,668 | 1,269 | 55,498 |
| Ongoing error | 2,743 | 2,264 | 2,832 | 32 | 1,425 | 0,297 |
| Expiring correct | 4,878 | 4,857 | 1,086 | 0 | 10,821 |  |
| Expiring error | 1,780 | 674 | 273 | 0 | 0 | 2,727 |
| Total | $1,999,575$ | 54,703 | 9,276 | 2,694 | $2,076,949$ |  |
|  |  |  |  |  |  |  |


| Nonparticipating | $1,642,911$ | 2,588 | 790 | 0 | 0 | $1,646,290$ |
| :--- | ---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | $0 *$ | 8,763 | 1,948 | 4,323 | 640 | 15,674 |
| Ongoing error | 1,642 | 1,022 | 890 | 17 | 4,384 |  |
| Expiring correct | 1,641 | 2,080 | 613 | 0 | 4,334 |  |
| Expiring error | 986 | 347 | 133 | 0 | 0 | 0 |
| Total | $1,647,180$ | 14,801 | 4,374 | 4,339 | 1,455 | $1,672,150$ |


| Nonparticipating | 347,296 | 4,582 | 470 | 0 | 0 | 352,348 |
| :--- | :---: | ---: | ---: | ---: | ---: | ---: |
| Ongoing correct | $0^{*}$ | 30,974 | 1,876 | 6,345 | 628 | 39,824 |
| Ongoing error | 1,091 | 1,243 | 1,942 | 16 | 610 | 0 |
| Expiring correct | 3,214 | 2,781 | 473 | 0 | 6,468 |  |
| Expiring error | 794 | 323 | 140 | 0 | 0 | 1,257 |
|  | 352,395 | 39,902 | 4,902 | 6,361 | 1,239 | 404,799 |

State: Wisconsin

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.996 | 0.004 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.716 | 0.069 | 0.192 | 0.023 | 1.000 |
| Ongoing error | 0.295 | 0.244 | 0.305 | 0.003 | 0.153 | 1.000 |
| Expiring correct | 0.451 | 0.449 | 0.100 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.653 | 0.247 | 0.100 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households with earnings |  |  |  |  |  |  |
| Nonparticipating | 0.998 | 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.559 | 0.124 | 0.276 | 0.041 | 1.000 |
| Ongoing error | 0.374 | 0.233 | 0.203 | 0.004 | 0.186 | 1.000 |
| Expiring correct | 0.379 | 0.480 | 0.141 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.672 | 0.237 | 0.091 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.986 | 0.013 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.778 | 0.047 | 0.159 | 0.016 | 1.000 |
| Ongoing error | 0.223 | 0.254 | 0.396 | 0.003 | 0.125 | 1.000 |
| Expiring correct | 0.497 | 0.430 | 0.073 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.632 | 0.257 | 0.111 | 0.000 | 0.000 | 1.000 |

State: Wyoming

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Non- participating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Cell counts for total households |  |  |  |  |  |
| Nonparticipating | 180,401 | 1,018 | 73 | 0 | 0 | 181,492 |
| Ongoing correct | 188 | 5,022 | 239 | 1,636 | 36 | 7,122 |
| Ongoing error | 159 | 190 | 147 | 0 | 37 | 533 |
| Expiring correct | 750 | 827 | 47 | 0 | 0 | 1,624 |
| Expiring error | 45 | 29 | 7 | 0 | 0 | 80 |
| Total | 181,543 | 7,086 | 513 | 1,636 | 73 | 190,852 |
|  | Cell counts for households with earnings |  |  |  |  |  |
| Nonparticipating | 152,287 | 371 | 50 | 0 | 0 | 152,707 |
| Ongoing correct | 0* | 1,666 | 157 | 908 | 36 | 2,768 |
| Ongoing error | 127 | 90 | 78 | 0 | 30 | 324 |
| Expiring correct | 338 | 531 | 23 | 0 | 0 | 892 |
| Expiring error | 43 | 19 | 7 | 0 | 0 | 69 |
| Total | 152,794 | 2,676 | 315 | 908 | 66 | 156,760 |
|  | Cell counts for households without earnings |  |  |  |  |  |
| Nonparticipating | 28,059 | 647 | 23 | 0 | 0 | 28,729 |
| Ongoing correct | 246 | 3,356 | 82 | 728 | 0 | 4,413 |
| Ongoing error | 31 | 101 | 69 | 0 | 8 | 208 |
| Expiring correct | 406 | 300 | 24 | 0 | 0 | 730 |
| Expiring error | 6 | 5 | 0 | 0 | 0 | 11 |
| Total | 28,748 | 4,410 | 198 | 728 | 8 | 34,092 |

State: Wyoming

| Current-month status | Next-month status |  |  |  |  |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Nonparticipating | Ongoing correct | Ongoing error | Expiring correct | Expiring error | Total |
|  | Transition probabilities for total households |  |  |  |  |  |
| Nonparticipating | 0.994 | 0.006 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.026 | 0.705 | 0.034 | 0.230 | 0.005 | 1.000 |
| Ongoing error | 0.298 | 0.357 | 0.275 | 0.000 | 0.070 | 1.000 |
| Expiring correct | 0.462 | 0.509 | 0.029 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.561 | 0.357 | 0.081 | 0.000 | 0.000 | 1.000 |
|  | Transition probabilities for households with earnings |  |  |  |  |  |
| Nonparticipating | 0.997 | 0.002 | 0.000 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.000* | 0.602 | 0.057 | 0.328 | 0.013 | 1.000 |
| Ongoing error | 0.391 | 0.277 | 0.240 | 0.000 | 0.092 | 1.000 |
| Expiring correct | 0.379 | 0.595 | 0.026 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.629 | 0.277 | 0.095 | 0.000 | 0.000 | 1.000 |
| Transition probabilities for households without earnings |  |  |  |  |  |  |
| Nonparticipating | 0.977 | 0.023 | 0.001 | 0.000 | 0.000 | 1.000 |
| Ongoing correct | 0.056 | 0.761 | 0.019 | 0.165 | 0.000 | 1.000 |
| Ongoing error | 0.150 | 0.484 | 0.330 | 0.000 | 0.036 | 1.000 |
| Expiring correct | 0.556 | 0.411 | 0.033 | 0.000 | 0.000 | 1.000 |
| Expiring error | 0.516 | 0.484 | 0.000 | 0.000 | 0.000 | 1.000 |

## Appendix E Discrete Time Markov Chains

This material is drawn from the following website for the University of Texas at Dallas, which provides Fall 2003 course notes for "Computer Sciences 6352: Performance of Computer Systems and Networks" (instructor: Assistant Professor Jason Jue):
http://www.utdallas.edu/~jjue/cs6352/markov.

## Discrete-Time Markov Chains

A Markov chain is a discrete state space process in which the next state depends only on the present state.
For a discrete time system, if $X_{n}$ is the state of the system at time $n$, then $\left\{X_{n}: n \geq 0\right\}$ is a Markov chain if:
$\operatorname{Pr}\left[X_{n}=j \mid X_{n-1}=i_{n-1}, X_{n-2}=i_{n-2}, \ldots, X_{0}=i_{0}\right]=\operatorname{Pr}\left[X_{n}=j \mid X_{n-1}=i_{n-1}\right]$,
i.e., the state $(j)$ of the system at time $n$ depends only on the state $(i)$ of the system at time $n-1$, and does not depend on any other state before time $n-1$.

## State Probabilities

The state probability, denoted as $\pi_{j}(n)$, is the probability that the process is in state $j$ at time $n$.

$$
\pi_{j}(n)=\operatorname{Pr}\left\{X_{n}=j\right\}
$$

The state probability vector is denoted as $\Pi(n)$, and consists of all of the state probabilities for a given time $n$.

$$
\Pi(n)=\left[\begin{array}{llll}
\pi_{0}(n) & \pi_{1}(n) & \pi_{2}(n) & \cdots
\end{array}\right]
$$

Note that the sum over the elements in $\Pi(n)$ is equal to 1 .

$$
\sum_{j} \pi_{j}(n)=1
$$

## Transition Probabilities

The one-step transition probability is the probability of transitioning from one state to another in a single step. The Markov chain is said to be time homogeneous if the transition probabilities from one state to another are independent of time index $n$.

$$
p_{i j}=\operatorname{Pr}\left\{X_{n}=j \mid X_{n-1}=i\right\}
$$

The transition probability matrix, $P$, is the matrix consisting of the one-step transition probabilities, $p_{i j}$.
The $m$-step transition probability is the probability of transitioning from state $i$ to state $j$ in $m$ steps.

$$
p_{i j}^{(m)}=\operatorname{Pr}\left\{X_{n+m}=j \mid X_{n}=i\right\}
$$

The $m$-step transition matrix whose elements are the $m$-step transition probabilities $p_{i j}^{(m)}$ is denoted as $P^{(m)}$.

The $m$-step transition probabilities can be found from the single-step transition probabilities as follows.

To transition from $i$ to $j$ in $m$ steps, the process can first transition from $i$ to $r$ in $m-k$ steps, and then transition from $r$ to $j$ in $k$ steps, where $0<k<m$.

$$
p_{i j}^{(m)}=\sum_{r} p_{i r}^{m-k} p_{r j}^{k}
$$

In matrix form, this becomes:

$$
P^{(m)}=P^{(m-k)} P^{(k)}
$$

Setting $k=m-1$ yields:

$$
P^{(m)}=P \cdot P^{(m-1)}
$$

From this equation we can see that:

$$
P^{(m-1)}=P \cdot P^{(m-2)}
$$

Substituting this back into the previous equation yields:

$$
P^{(m)}=P \cdot P \cdot P^{(m-2)}
$$

Continuing these substitutions, eventually we have:

$$
P^{(m)}=P \cdot P \cdot P \cdot \cdot P=P^{m}
$$

Therefore, the $m$-step transition probability matrix can be found by multiplying the single-step probability matrix by itself $m$ times.

The state vector at time $m$ can also be found in terms of the transition probability matrix and the intial state vector $\Pi(0)$. We first observe that:

$$
\pi_{j}(m)=\sum_{i} \pi_{i}(m-1) p_{i j}
$$

In vector and matrix form, this becomes:

$$
\Pi(m)=\Pi(m-1) P
$$

We also find that, through substitution:

$$
\Pi(m-1)=\Pi(m-2) P
$$

or,

$$
\Pi(m)=\Pi(m-2) P \cdot P
$$

Continuing the substitution yields:

$$
\Pi(m)=\Pi(0) P^{m_{i}}
$$

where $\Pi(0)$ is the vector containing the initial probabilities of being in each state at time 0 .

## Long-Run Behavior of Markov Chains

As the time index $m$ approaches infinity, a Markov chain may settle down and exhibit steady-state behavior. If the following limit exists:

$$
\lim _{m \rightarrow \infty} p_{i j}^{(m)}=\pi_{j}
$$

for all values of $i$, then the $\left\{\pi_{j}\right\}$ are the limiting or steady-state probabilities.
Looking at the state probability as $m$ approaches infinity, we see that:

$$
\begin{align*}
\lim _{m \rightarrow \infty} \pi_{j}(m) & =\lim _{m \rightarrow \infty} \sum_{i} \pi_{i}(0) p_{i j}(m)  \tag{1}\\
& =\sum_{i} \pi_{i}(0) \lim _{m \rightarrow \infty} p_{i j}(m) \\
& =\sum_{i} \pi_{i}(0) \pi_{j} \\
& =\pi_{j} \sum_{i} \pi_{i}(0) \\
& =\pi_{j}
\end{align*}
$$

When the limiting probabilities exist, they can be found using the following equations:

$$
\Pi=\Pi P
$$

and

$$
\sum_{i} \pi_{i}=1
$$

where

$$
\Pi=\left[\begin{array}{llll}
\pi_{0} & \pi_{1} & \pi_{2} & \cdots
\end{array}\right]
$$


[^0]:    1 See Kabbani and Wilde (2003) and Kornfeld (2002).

[^1]:    2 For a detailed analysis of the factors influencing food stamp caseloads between 1987 and 1999, see Kornfeld (2002).

    3 Alternative estimates show national food stamp participation rates among program-eligibles that are higher in each year than those noted here, but with a similar trend over time. In Shirm and Castner (2002), the national participation rate was estimated at 60 percent for 1998, 58 percent for 1999, and 59 percent for 2000.

[^2]:    $4 \quad$ See Affholter and Kramer (1987), Chapter 3.
    5 The following discussion draws heavily from Rosenbaum and Super (2001).

[^3]:    ${ }^{6}$ Food and Nutrition Service (May 24, 2002).

[^4]:    9 In this study, we use the term "aggregate participation rate" to refer here to the percentage of all households that participate in the program. This is to distinguish it from the alternative measure (referred to here as the "conditional participation rate") that indicates the percentage of program-eligible households that participate.

[^5]:    10 There are some combinations of $a$ and $b$ for which no equilibrium is reached (for instance, if $a=1$ and $b=1$ ), but such scenarios are implausible ones that do not reflect the real-world monthly dynamics.

[^6]:    ${ }^{11}$ As indicated in Chapter Three, one can derive the equilibrium value of the case error rate $\left(\mathrm{r}^{*}\right)$, as $\mathrm{e}^{*} / \mathrm{p}^{*}$, where $\mathrm{p}^{*}$ is the equilibrium value for the share of all households that are food stamp participants (as defined above) and $\mathrm{e}^{*}$ is the equilibrium value for the share of all households that are error cases.

[^7]:    12 The assumption also reflects a recognized limitation of the QC data. As will be explained later, the QC data do not accurately identify all cases that are overdue for recertification.

[^8]:    13 Similarly, separate models could be constructed for agency-related error and client-related error. The distinction between the two error types can be somewhat arbitrary, however, and one needs to establish rules for classifying multiple-error cases that contain both agency and client errors.

    14 The correlation coefficient of 0.923 is based on state-reported values of the total case error rate and the official values of the total dollar error rate in fiscal year 2001, for the 50 states and the District of Columbia.

[^9]:    ${ }^{15}$ A subsample of each state's QC sample is then subject to validation by federal QC staff, and the official error statistics incorporate the findings of both the state reviews and federal re-reviews. The federal rereview findings are not used in this study, as the federal subsample is about one-third the size of the full QC sample.

[^10]:    16 The annual QC data files, including details on error cases not included in public-use data sets, were made available to us by Mathematica Policy Research (MPR), at the request of the Food and Nutrition Service (FNS). We wish to acknowledge the assistance provided by Jenny Genser of FNS and Karen Cunnyngham of MPR in making the detailed data available for this analysis.

[^11]:    Note: Row entries may not sum to the indicated row total due to rounding. Asterisk (*) indicates an imputed zero value.

[^12]:    ${ }^{17}$ In this particular example, the computed marginal row total for nonparticipating households, 99.015 million (for "March 2001"), equals the corresponding first-column total, also 99.015 million (for "April 2001"). These two values are not equal by assumption.

[^13]:    18 An alternative specification of the model was suggested by the Economic Research Service, employing a different approach to the calculation of row totals in the transition matrix that yielded lower case closure rates. This alternative model was considered but was not adopted, in part because the long-term outcomes implied by the transition matrix appeared less credible than those shown in Exhibit 12. In particular, under the alternative formulation, the long-term participation rate for total households was 7.91 percent, implying a substantial increase in the active caseload. The corresponding long-term case error rate was 11.03 percent.

[^14]:    19 See Bartlett et al. (2004). The study derived its nationally representative estimates for June 2000 from data collected at 109 local food stamp offices.

[^15]:    20 For the data element "sample month in certification," FNS Handbook 310 instructs QC reviewers that "this entry should indicate how far into the certification period the sample month occurs." The instructions go on to say, however, that "for households that are participating in months for which they have not been certified, enter the number of months beyond the end of the household's certification period." This implies that a case one month overdue for recertification could be misclassified as a newly certified case. Consistent with this logic, the 11.2 percent caseload share for expiring cases in 2000 (shown in Exhibit 14) is lower than the corresponding 12.5 percent from Bartlett et al. (2004).

[^16]:    ${ }^{21}$ Several aspects of this study should be re-emphasized, to avoid confusion. First, for expositional ease we refer to case transitions as occurring from the "current" month to the "next" month. The expected pattern of these month-to-month transitions is derived from QC data, indicating case status in the review month and enabling us to infer the case status in the prior month. One should thus regard the QC review month as the "next" month; the month preceding the QC review is the "current" month. Second, the term "expiring cases" refers to cases that are not necessarily closing; instead, these are cases that are at the end of their certification period and are thus subject to a recertification.

[^17]:    22 These model-derived case error rates differ from those shown in Exhibit 2 for several reasons. First and most importantly, the model applies a consistent $\$ 25$ error threshold for all years. (In Exhibit 2, the error rates prior to 2000 reflect a $\$ 5$ error threshold. As shown in Exhibit 8, in those years approximately onethird of recorded errors amounted to less than $\$ 25$ for eligible cases.) Second, Guam and the Virgin Islands are excluded from the model-derived estimates.

[^18]:    23 See, for instance, Rosenbaum (2000).

[^19]:    1 This variable is drawn from the "detailed error findings" recorded on the Quality Control Review Schedule, and is available only in the backup files made available for this analysis. The date of occurrence is recorded for up to nine errors that may have been discovered during the QC review.

[^20]:    ${ }^{2}$ We also tested the model under alternative assumptions regarding this time window. It appears that reviewers may use the occurrence date to indicate the date of the underlying change in household circumstances that precipitated the error, rather than the date of onset of the QC error itself. Once a change in circumstances occurs, there is an allowable time interval for the client to report the change and a further time interval for the agency to act on the reported change.
    ${ }^{3}$ This assumption may overstate the number of cases previously in error, to the extent that: (a) an interim change may serve to prevent an error from arising, through timely benefit adjustment; or (b) an interim change may be recorded even though the caseworker simply processed a monthly or quarterly report without a change in benefit. However, the assumption may also understate the number of cases previously in error, by not taking account of "transient" errors that "self-correct"-i.e., errors associated with unreported short-term changes in household circumstances, such as temporary increases in income. Such cases "self-correct" (without an interim change) when the household's circumstances return to the status quo ante.

[^21]:    4 We tested alternative specifications of this relationship. The one shown above was found to provide the most plausible results, based on information known about the length of assigned certification periods. If the mean length of certification period is $n$, the number of cases subject to recertification in any given month should be approximately $1 / \mathrm{n}$. One alternative tested and rejected was to consider a case to be expiring if LASTCERT \$CERTMTH. Under this definition the share of total cases classified as "expiring" each month was less than 0.100 . This proportion was implausibly low, given that the mean certification length is between 9.5 and 10.0. The need to calibrate the model under alternative specifications arose from the fact that the QC review schedule records only the month and year (but not the day) of key events in a case's history.
    5 The model will tend to understate somewhat the number of cases that are due or overdue for recertification, for the following reason. In entering the variable LASTCERT, QC reviewers are instructed that "For households that are participating in months for which they have not been certified, enter the number of months beyond the end of the household's certification period." For instance, if a case did not undergo its scheduled recertification at the $12^{\text {th }}$ month and is now participating at the $15^{\text {th }}$ month, LASTCERT should be coded as 3 , not 15 . It is apparent, however, that not all reviewers follow this instruction, as there are cases for which LASTCERT exceeds CERTMTH.

[^22]:    ${ }^{6}$ Here, once again, we tested and rejected alternative specifications, including YRMONTH-OCCDATE\#1. The latter definition was found to assign an implausibly large number of cases to $\mathrm{Q}_{35}$ and $\mathrm{Q}_{55}$, resulting in large negative entries for the $\mathrm{Q}_{31}$ and $\mathrm{Q}_{51}$. The specification adopted here appears consistent with the QC rules that call upon reviewers to disregard errors (or "exclude variances") associated with changes in client circumstances that occur in the month immediately prior to the review month.

[^23]:    7 The calculation of row and column residuals must proceed on a specified order. The value $\mathrm{Q}_{42}$ is the last to be derived, preceded by $\mathrm{Q}_{52}$ and the first-column cells.

[^24]:    8 See Mills (1990).

[^25]:    9 It is necessary to first derive $\mathrm{Q}_{51}$ before computing $\mathrm{Q}_{41}$.

